

# 1 Time Series Analysis (時系列分析)

## 1.1 Introduction

代表的テキスト：

- J.D. Hamilton (1994) *Econometric Analysis*  
沖本・井上訳 (2006) 『時系列解析(上・下)』
- A.C. Harvey (1981) *Time Series Models*  
国友・山本訳 (1985) 『時系列モデル入門』
- 沖本竜義 (2010) 『経済・ファイナンスデータの計量時系列分析』

## 1. **Sstationarity (定常性) :**

Let  $y_1, y_2, \dots, y_T$  be time series data.

### (a) **Weak Stationarity (弱定常性) :**

$$E(y_t) = \mu,$$

$$E((y_t - \mu)(y_{t-\tau} - \mu)) = \gamma(\tau), \quad \tau = 0, 1, 2, \dots$$

The first and second moments depend on time difference, not time itself.

(b) **Strong Stationarity (強定常性) :**

Let  $f(y_{t_1}, y_{t_2}, \dots, y_{t_r})$  be the joint distribution of  $y_{t_1}, y_{t_2}, \dots, y_{t_r}$ .

$$f(y_{t_1}, y_{t_2}, \dots, y_{t_r}) = f(y_{t_1+\tau}, y_{t_2+\tau}, \dots, y_{t_r+\tau})$$

All the moments are same for all  $\tau$ .

## 2. Ergodicity (エルゴード性) :

As time difference between two data is large, the two data become independent.

$y_1, y_2, \dots, y_T$  is said to be ergodic in mean when  $\bar{y}$  converges in probability to  $E(y_t)$ .

## 3. Auto-covariance Function (自己共分散関数) :

$$E((y_t - \mu)(y_{t-\tau} - \mu)) = \gamma(\tau), \quad \tau = 0, 1, 2, \dots$$

$$\gamma(\tau) = \gamma(-\tau)$$

4. **Auto-correlation Function (自己相関関数) :**

$$\rho(\tau) = \frac{E((y_t - \mu)(y_{t-\tau} - \mu))}{\sqrt{\text{Var}(y_t)} \sqrt{\text{Var}(y_{t-\tau})}} = \frac{\gamma(\tau)}{\gamma(0)}$$

Note that  $\text{Var}(y_t) = \text{Var}(y_{t-\tau}) = \gamma(0)$ .

5. **Sample Mean (標本平均) :**

$$\hat{\mu} = \frac{1}{T} \sum_{t=1}^T y_t$$

6. **Sample Auto-covariance (標本自己共分散) :**

$$\hat{\gamma}(\tau) = \frac{1}{T} \sum_{t=\tau+1}^T (y_t - \hat{\mu})(y_{t-\tau} - \hat{\mu})$$

7. **Corelogram** (コレログラム, or 標本自己相関関数) :

$$\hat{\rho}(\tau) = \frac{\hat{\gamma}(\tau)}{\hat{\gamma}(0)}$$

8. **Lag Operator** (ラグ作要素) :

$$L^\tau y_t = y_{t-\tau}, \quad \tau = 1, 2, \dots$$

9. **Likelihood Function** (尤度関数) — Innovation Form :

The joint distribution of  $y_1, y_2, \dots, y_T$  is written as:

$$\begin{aligned} f(y_1, y_2, \dots, y_T) &= f(y_T | y_{T-1}, \dots, y_1) f(y_{T-1}, \dots, y_1) \\ &= f(y_T | y_{T-1}, \dots, y_1) f(y_{T-1} | y_{T-2}, \dots, y_1) f(y_{T-2}, \dots, y_1) \\ &\quad \vdots \\ &= f(y_T | y_{T-1}, \dots, y_1) f(y_{T-1} | y_{T-2}, \dots, y_1) \cdots f(y_2 | y_1) f(y_1) \\ &= f(y_1) \prod_{t=2}^T f(y_t | y_{t-1}, \dots, y_1). \end{aligned}$$

Therefore, the log-likelihood function is given by:

$$\log f(y_1, y_2, \dots, y_T) = \log f(y_1) + \sum_{t=2}^T \log f(y_t | y_{t-1}, \dots, y_1).$$

Under the normality assumption,  $f(y_t | y_{t-1}, \dots, y_1)$  is given by the normal distribution with conditional mean  $E(y_t | y_{t-1}, \dots, y_1)$  and conditional variance  $\text{Var}(y_t | y_{t-1}, \dots, y_1)$ .

## 1.2 Autoregressive Model (自己回帰モデル or AR モデル)

### 1. AR( $p$ ) Model :

$$y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \epsilon_t,$$

which is rewritten as:

$$\phi(L)y_t = \epsilon_t,$$

where

$$\phi(L) = 1 - \phi_1 L - \phi_2 L^2 - \cdots - \phi_p L^p.$$

## 2. **Stationarity** (定常性) :

Suppose that all the  $p$  solutions of  $x$  from  $\phi(x) = 0$  are real numbers

When the  $p$  solutions are greater than one,  $y_t$  is stationary.

Suppose that the  $p$  solutions include imaginary numbers.

When the  $p$  solutions are outside unit circle,  $y_t$  is stationary.

### 3. **Partial Autocorrelation Coefficient (偏自己相関係数), $\phi_{k,k}$ :**

The partial autocorrelation coefficient between  $y_t$  and  $y_{t-k}$ , denoted by  $\phi_{k,k}$ , is a measure of strength of the relationship between  $y_t$  and  $y_{t-k}$ , after removing influence of  $y_{t-1}, \dots, y_{t-k+1}$ .

$$\phi_{1,1} = \rho(1)$$

$$\begin{pmatrix} 1 & \rho(1) \\ \rho(1) & 1 \end{pmatrix} \begin{pmatrix} \phi_{2,1} \\ \phi_{2,2} \end{pmatrix} = \begin{pmatrix} \rho(1) \\ \rho(2) \end{pmatrix}$$

$$\begin{pmatrix} 1 & \rho(1) & \rho(2) \\ \rho(1) & 1 & \rho(1) \\ \rho(2) & \rho(1) & 1 \end{pmatrix} \begin{pmatrix} \phi_{3,1} \\ \phi_{3,2} \\ \phi_{3,3} \end{pmatrix} = \begin{pmatrix} \rho(1) \\ \rho(2) \\ \rho(3) \end{pmatrix}$$

⋮

$$\begin{pmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(k-1) \\ \rho(1) & 1 & & \rho(k-3) & \rho(k-2) \\ \vdots & \vdots & & \vdots & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & 1 \end{pmatrix} \begin{pmatrix} \phi_{k,1} \\ \phi_{k,2} \\ \vdots \\ \phi_{k,k-1} \\ \phi_{k,k} \end{pmatrix} = \begin{pmatrix} \rho(1) \\ \rho(2) \\ \vdots \\ \rho(k) \end{pmatrix}$$

Use Cramer's rule (クラメールの公式) to obtain  $\phi_{k,k}$ .

$$\phi_{k,k} = \frac{\begin{vmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(1) \\ \rho(1) & 1 & & \rho(k-3) & \rho(2) \\ \vdots & \vdots & & \vdots & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & \rho(k) \end{vmatrix}}{\begin{vmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(k-1) \\ \rho(1) & 1 & & \rho(k-3) & \rho(k-2) \\ \vdots & \vdots & & \vdots & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & 1 \end{vmatrix}}$$

**Example: AR(1) Model:**  $y_t = \phi_1 y_{t-1} + \epsilon_t$

1. The stationarity condition is: the solution of  $\phi(x) = 1 - \phi_1 x = 0$ , i.e.,  $x = 1/\phi_1$ , is greater than one, or equivalently,  $\phi_1 < 1$ .
2. Rewriting the AR(1) model,

$$\begin{aligned}y_t &= \phi_1 y_{t-1} + \epsilon_t \\&= \phi_1^2 y_{t-2} + \epsilon_t + \phi_1 \epsilon_{t-1} \\&= \phi_1^3 y_{t-3} + \epsilon_t + \phi_1 \epsilon_{t-1} + \phi_1^2 \epsilon_{t-2} \\&\vdots\end{aligned}$$

$$= \phi_1^s y_{t-s} + \epsilon_t + \phi_1 \epsilon_{t-1} + \dots + \phi_1^{s-1} \epsilon_{t-s+1}.$$

As  $s$  is large,  $\phi_1^s$  approaches zero.  $\implies$  Stationarity condition

3. For stationarity,  $y_t = \phi_1 y_{t-1} + \epsilon_t$  is rewritten as:

$$y_t = \epsilon_t + \phi_1 \epsilon_{t-1} + \phi_1^2 \epsilon_{t-2} + \dots$$

MA representation of AR model.

(MA will be discussed later.)

4. Mean of AR(1) process,  $\mu$

$$\begin{aligned}\mu &= E(y_t) = E(\epsilon_t + \phi_1\epsilon_{t-1} + \phi_1^2\epsilon_{t-2} + \dots) \\ &= E(\epsilon_t) + \phi_1E(\epsilon_{t-1}) + \phi_1^2E(\epsilon_{t-2}) + \dots = 0\end{aligned}$$

5. Autocovariance and autocorrelation functions of the AR(1) process:

Rewriting the AR(1) process, we have:

$$y_t = \phi_1^\tau y_{t-\tau} + \epsilon_t + \phi_1\epsilon_{t-1} + \dots + \phi_1^{\tau-1}\epsilon_{t-\tau+1}.$$

Therefore, the autocovariance function of AR(1) process is:

$$\begin{aligned}\gamma(\tau) &= E((y_t - \mu)(y_{t-\tau} - \mu)) = E(y_t y_{t-\tau}) \\ &= E\left((\phi_1^\tau y_{t-\tau} + \epsilon_t + \phi_1 \epsilon_{t-1} + \dots + \phi_1^{\tau-1} \epsilon_{t-\tau+1})y_{t-\tau}\right) \\ &= \phi_1^\tau E(y_{t-\tau} y_{t-\tau}) + E(\epsilon_t y_{t-\tau}) + \phi_1 E(\epsilon_{t-1} y_{t-\tau}) + \dots + \phi_1^{\tau-1} E(\epsilon_{t-\tau+1} y_{t-\tau}) \\ &= \phi_1^\tau \gamma(0).\end{aligned}$$

The autocorrelation function of AR(1) process is:

$$\rho(\tau) = \frac{\gamma(\tau)}{\gamma(0)} = \phi_1^\tau.$$

Multiply  $y_{t-\tau}$  on both sides of the AR(1) process and take the expectation:

$$E(y_t y_{t-\tau}) = \phi_1 E(y_{t-1} y_{t-\tau}) + E(\epsilon_t y_{t-\tau})$$

$$\gamma(\tau) = \begin{cases} \phi_1 \gamma(\tau - 1), & \text{for } \tau \neq 0, \\ \phi_1 \gamma(\tau - 1) + \sigma^2, & \text{for } \tau = 0. \end{cases}$$

Using  $\gamma(\tau) = \gamma(-\tau)$ ,  $\gamma(\tau)$  for  $\tau = 0$  is given by:

$$\gamma(0) = \phi_1 \gamma(1) + \sigma^2 = \phi_1^2 \gamma(0) + \sigma^2.$$

Note that  $\gamma(1) = \phi_1 \gamma(0)$ .

Therefore,  $\gamma(0)$  is given by:

$$\gamma(0) = \frac{\sigma^2}{1 - \phi_1^2}$$

6. Partial autocorrelation function of AR(1) process:

$$\phi_{1,1} = \rho(1) = \phi_1$$

$$\phi_{2,2} = \frac{\begin{vmatrix} 1 & \rho(1) \\ \rho(1) & \rho(2) \end{vmatrix}}{\begin{vmatrix} 1 & \rho(1) \\ \rho(1) & 1 \end{vmatrix}} = \frac{\rho(2) - \rho(1)^2}{1 - \rho(1)^2} = 0$$

## 7. Estimation of AR(1) model:

### (a) Likelihood function

$$\begin{aligned}\log f(y_T, \dots, y_1) &= \log f(y_1) + \sum_{t=1}^T \log f(y_t | y_{t-1}, \dots, y_1) \\ &= -\frac{1}{2} \log(2\pi) - \frac{1}{2} \log\left(\frac{\sigma^2}{1 - \phi_1^2}\right) - \frac{1}{\sigma^2 / (1 - \phi_1^2)} y_1^2 \\ &\quad - \frac{T-1}{2} \log(2\pi) - \frac{T-1}{2} \log(\sigma^2) - \frac{1}{\sigma^2} \sum_{t=2}^T (y_t - \phi_1 y_{t-1})^2\end{aligned}$$

$$\begin{aligned}
&= -\frac{T}{2} \log(2\pi) - \frac{T}{2} \log(\sigma^2) - \frac{1}{2} \log\left(\frac{1}{1 - \phi_1^2}\right) \\
&\quad - \frac{1}{2\sigma^2/(1 - \phi_1^2)} y_1^2 - \frac{1}{2\sigma^2} \sum_{t=2}^T (y_t - \phi_1 y_{t-1})^2
\end{aligned}$$

Note as follows:

$$\begin{aligned}
f(y_1) &= \frac{1}{\sqrt{2\pi\sigma^2/(1 - \phi_1^2)}} \exp\left(-\frac{1}{2\sigma^2/(1 - \phi_1^2)} y_1^2\right) \\
f(y_t|y_{t-1}, \dots, y_1) &= \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2} (y_t - \phi_1 y_{t-1})^2\right)
\end{aligned}$$

$$\frac{\partial \log f(y_T, \dots, y_1)}{\partial \sigma^2} = -\frac{T}{2} \frac{1}{\sigma^2} + \frac{1}{2\sigma^4/(1-\phi_1^2)} y_1^2 + \frac{1}{2\sigma^4} \sum_{t=2}^T (y_t - \phi_1 y_{t-1})^2 = 0$$

$$\frac{\partial \log l(y_T, \dots, y_1)}{\partial \phi_1} = -\frac{\phi_1}{1-\phi_1^2} + \frac{\phi_1}{\sigma^2} y_1^2 + \frac{1}{\sigma^2} \sum_{t=2}^T (y_t - \phi_1 y_{t-1}) y_{t-1} = 0$$

The MLE of  $\phi_1$  and  $\sigma^2$  satisfies the above two equation.

$$\tilde{\sigma}^2 = \frac{1}{T} \left( (1 - \tilde{\phi}_1^2) y_1^2 + \sum_{t=2}^T (y_t - \tilde{\phi}_1 y_{t-1})^2 \right)$$

$$\tilde{\phi}_1 = \frac{\sum_{t=2}^T y_t y_{t-1}}{\sum_{t=2}^T y_{t-1}^2} + \left( \tilde{\phi}_1 y_1^2 - \frac{\tilde{\sigma}^2 \tilde{\phi}_1}{1 - \tilde{\phi}_1^2} \right) / \sum_{t=2}^T y_{t-1}^2$$

(b) Ordinary Least Squares (OLS) Method

$$S(\phi_1) = \sum_{t=2}^T (y_t - \phi_1 y_{t-1})^2$$

is minimized with respect to  $\phi_1$ .

$$\begin{aligned}\hat{\phi}_1 &= \frac{\sum_{t=2}^T y_{t-1} y_t}{\sum_{t=2}^T y_{t-1}^2} = \phi_1 + \frac{\sum_{t=2}^T y_{t-1} \epsilon_t}{\sum_{t=2}^T y_{t-1}^2} = \phi_1 + \frac{(1/T) \sum_{t=2}^T y_{t-1} \epsilon_t}{(1/T) \sum_{t=2}^T y_{t-1}^2} \\ &\longrightarrow \phi_1 + \frac{E(y_{t-1} \epsilon_t)}{E(y_{t-1}^2)} = \phi_1\end{aligned}$$

OLSE of  $\phi_1$  is a consistent estimator.