

## 8. Lag Operator (ラグ作要素) :

$$L^\tau y_t = y_{t-\tau}, \quad \tau = 1, 2, \dots$$

## 9. Likelihood Function (尤度関数) — Innovation Form :

The joint distribution of  $y_1, y_2, \dots, y_T$  is written as:

$$\begin{aligned} f(y_1, y_2, \dots, y_T) &= f(y_T | y_{T-1}, \dots, y_1) f(y_{T-1}, \dots, y_1) \\ &= f(y_T | y_{T-1}, \dots, y_1) f(y_{T-1} | y_{T-2}, \dots, y_1) f(y_{T-2}, \dots, y_1) \\ &\quad \vdots \\ &= f(y_T | y_{T-1}, \dots, y_1) f(y_{T-1} | y_{T-2}, \dots, y_1) \cdots f(y_2 | y_1) f(y_1) \\ &= f(y_1) \prod_{t=2}^T f(y_t | y_{t-1}, \dots, y_1). \end{aligned}$$

Therefore, the log-likelihood function is given by:

$$\log f(y_1, y_2, \dots, y_T) = \log f(y_1) + \sum_{t=2}^T \log f(y_t | y_{t-1}, \dots, y_1).$$

Under the normality assumption,  $f(y_t | y_{t-1}, \dots, y_1)$  is given by the normal distribution with conditional mean  $E(y_t | y_{t-1}, \dots, y_1)$  and conditional variance  $\text{Var}(y_t | y_{t-1}, \dots, y_1)$ .

## 6.2 Autoregressive Model (自己回帰モデル or AR モデル)

### 1. AR( $p$ ) Model :

$$y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \dots + \phi_p y_{t-p} + \epsilon_t,$$

which is rewritten as:

$$\phi(L)y_t = \epsilon_t,$$

where

$$\phi(L) = 1 - \phi_1 L - \phi_2 L^2 - \dots - \phi_p L^p.$$

## 2. Stationarity (定常性) :

Suppose that all the  $p$  solutions of  $x$  from  $\phi(x) = 0$  are real numbers

When the  $p$  solutions are greater than one,  $y_t$  is stationary.

Suppose that the  $p$  solutions include imaginary numbers.

When the  $p$  solutions are outside unit circle,  $y_t$  is stationary.

## 3. Partial Autocorrelation Coefficient (偏自己相關係數), $\phi_{k,k}$ :

The partial autocorrelation coefficient between  $y_t$  and  $y_{t-k}$ , denoted by  $\phi_{k,k}$ , is a measure of strength of the relationship between  $y_t$  and  $y_{t-k}$ , after removing

influence of  $y_{t-1}, \dots, y_{t-k+1}$ .

$$\phi_{1,1} = \rho(1)$$

$$\begin{pmatrix} 1 & \rho(1) \\ \rho(1) & 1 \end{pmatrix} \begin{pmatrix} \phi_{2,1} \\ \phi_{2,2} \end{pmatrix} = \begin{pmatrix} \rho(1) \\ \rho(2) \end{pmatrix}$$

$$\begin{pmatrix} 1 & \rho(1) & \rho(2) \\ \rho(1) & 1 & \rho(1) \\ \rho(2) & \rho(1) & 1 \end{pmatrix} \begin{pmatrix} \phi_{3,1} \\ \phi_{3,2} \\ \phi_{3,3} \end{pmatrix} = \begin{pmatrix} \rho(1) \\ \rho(2) \\ \rho(3) \end{pmatrix}$$

⋮

$$\begin{pmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(k-1) \\ \rho(1) & 1 & & \rho(k-3) & \rho(k-2) \\ \vdots & \vdots & & \vdots & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & 1 \end{pmatrix} \begin{pmatrix} \phi_{k,1} \\ \phi_{k,2} \\ \vdots \\ \phi_{k,k-1} \\ \phi_{k,k} \end{pmatrix} = \begin{pmatrix} \rho(1) \\ \rho(2) \\ \vdots \\ \rho(k) \end{pmatrix}$$

Use Cramer's rule (クラメールの公式) to obtain  $\phi_{k,k}$ .

$$\phi_{k,k} = \frac{\begin{vmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(1) \\ \rho(1) & 1 & & & \rho(k-3) \rho(2) \\ \vdots & \vdots & & & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & \rho(k) \end{vmatrix}}{\begin{vmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(k-1) \\ \rho(1) & 1 & & & \rho(k-3) \rho(k-2) \\ \vdots & \vdots & & & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & 1 \end{vmatrix}}$$

**Example: AR(1) Model:**  $y_t = \phi_1 y_{t-1} + \epsilon_t$

1. The stationarity condition is: the solution of  $\phi(x) = 1 - \phi_1 x = 0$ , i.e.,  $x = 1/\phi_1$ , is greater than one in absolute value, or equivalently,  $|\phi_1| < 1$ .

2. Rewriting the AR(1) model,

$$\begin{aligned}y_t &= \phi_1 y_{t-1} + \epsilon_t \\&= \phi_1^2 y_{t-2} + \epsilon_t + \phi_1 \epsilon_{t-1} \\&= \phi_1^3 y_{t-3} + \epsilon_t + \phi_1 \epsilon_{t-1} + \phi_1^2 \epsilon_{t-2} \\&\quad \vdots \\&= \phi_1^s y_{t-s} + \epsilon_t + \phi_1 \epsilon_{t-1} + \cdots + \phi_1^{s-1} \epsilon_{t-s+1}.\end{aligned}$$

As  $s$  is large,  $\phi_1^s$  approaches zero.  $\implies$  Stationarity condition

3. For stationarity,  $y_t = \phi_1 y_{t-1} + \epsilon_t$  is rewritten as:

$$y_t = \epsilon_t + \phi_1 \epsilon_{t-1} + \phi_1^2 \epsilon_{t-2} + \cdots$$

MA representation of AR model.

(MA will be discussed later.)

#### 4. Mean of AR(1) process, $\mu$

$$\begin{aligned}\mu &= E(y_t) = E(\epsilon_t + \phi_1 \epsilon_{t-1} + \phi_1^2 \epsilon_{t-2} + \dots) \\ &= E(\epsilon_t) + \phi_1 E(\epsilon_{t-1}) + \phi_1^2 E(\epsilon_{t-2}) + \dots = 0\end{aligned}$$

#### 5. Autocovariance and autocorrelation functions of the AR(1) process:

Rewriting the AR(1) process, we have:

$$y_t = \phi_1^\tau y_{t-\tau} + \epsilon_t + \phi_1 \epsilon_{t-1} + \dots + \phi_1^{\tau-1} \epsilon_{t-\tau+1}.$$

Therefore, the autocovariance function of AR(1) process is:

$$\begin{aligned}\gamma(\tau) &= E((y_t - \mu)(y_{t-\tau} - \mu)) = E(y_t y_{t-\tau}) \\ &= E\left((\phi_1^\tau y_{t-\tau} + \epsilon_t + \phi_1 \epsilon_{t-1} + \dots + \phi_1^{\tau-1} \epsilon_{t-\tau+1}) y_{t-\tau}\right) \\ &= \phi_1^\tau E(y_{t-\tau} y_{t-\tau}) + E(\epsilon_t y_{t-\tau}) + \phi_1 E(\epsilon_{t-1} y_{t-\tau}) + \dots + \phi_1^{\tau-1} E(\epsilon_{t-\tau+1} y_{t-\tau}) \\ &= \phi_1^\tau \gamma(0).\end{aligned}$$

The autocorrelation function of AR(1) process is:

$$\rho(\tau) = \frac{\gamma(\tau)}{\gamma(0)} = \phi_1^\tau.$$

Multiply  $y_{t-\tau}$  on both sides of the AR(1) process and take the expectation:

$$\begin{aligned} E(y_t y_{t-\tau}) &= \phi_1 E(y_{t-1} y_{t-\tau}) + E(\epsilon_t y_{t-\tau}) \\ \gamma(\tau) &= \begin{cases} \phi_1 \gamma(\tau - 1), & \text{for } \tau \neq 0, \\ \phi_1 \gamma(\tau - 1) + \sigma^2, & \text{for } \tau = 0. \end{cases} \end{aligned}$$

Using  $\gamma(\tau) = \gamma(-\tau)$ ,  $\gamma(\tau)$  for  $\tau = 0$  is given by:

$$\gamma(0) = \phi_1 \gamma(1) + \sigma^2 = \phi_1^2 \gamma(0) + \sigma^2.$$

Note that  $\gamma(1) = \phi_1 \gamma(0)$ .

Therefore,  $\gamma(0)$  is given by:

$$\gamma(0) = \frac{\sigma^2}{1 - \phi_1^2}$$

6. Partial autocorrelation function of AR(1) process:

$$\phi_{1,1} = \rho(1) = \phi_1$$

$$\phi_{2,2} = \frac{\begin{vmatrix} 1 & \rho(1) \\ \rho(1) & \rho(2) \end{vmatrix}}{\begin{vmatrix} 1 & \rho(1) \\ \rho(1) & 1 \end{vmatrix}} = \frac{\rho(2) - \rho(1)^2}{1 - \rho(1)^2} = 0$$

7. Estimation of AR(1) model:

(a) Likelihood function

$$\log f(y_T, \dots, y_1) = \log f(y_1) + \sum_{t=1}^T \log f(y_t | y_{t-1}, \dots, y_1)$$

$$\begin{aligned}
&= -\frac{1}{2} \log(2\pi) - \frac{1}{2} \log\left(\frac{\sigma^2}{1-\phi_1^2}\right) - \frac{1}{\sigma^2/(1-\phi_1^2)} y_1^2 \\
&\quad - \frac{T-1}{2} \log(2\pi) - \frac{T-1}{2} \log(\sigma^2) - \frac{1}{\sigma^2} \sum_{t=2}^T (y_t - \phi_1 y_{t-1})^2 \\
\\
&= -\frac{T}{2} \log(2\pi) - \frac{T}{2} \log(\sigma^2) - \frac{1}{2} \log\left(\frac{1}{1-\phi_1^2}\right) \\
&\quad - \frac{1}{2\sigma^2/(1-\phi_1^2)} y_1^2 - \frac{1}{2\sigma^2} \sum_{t=2}^T (y_t - \phi_1 y_{t-1})^2
\end{aligned}$$

Note as follows:

$$\begin{aligned}
f(y_1) &= \frac{1}{\sqrt{2\pi\sigma^2/(1-\phi_1^2)}} \exp\left(-\frac{1}{2\sigma^2/(1-\phi_1^2)} y_1^2\right) \\
f(y_t|y_{t-1}, \dots, y_1) &= \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left(-\frac{1}{2\sigma^2} (y_t - \phi_1 y_{t-1})^2\right)
\end{aligned}$$

$$\frac{\partial \log f(y_T, \dots, y_1)}{\partial \sigma^2} = -\frac{T}{2} \frac{1}{\sigma^2} + \frac{1}{2\sigma^4/(1-\phi_1^2)} y_1^2 + \frac{1}{2\sigma^4} \sum_{t=2}^T (y_t - \phi_1 y_{t-1})^2 = 0$$

$$\frac{\partial \log f(y_T, \dots, y_1)}{\partial \phi_1} = -\frac{\phi_1}{1-\phi_1^2} + \frac{\phi_1}{\sigma^2} y_1^2 + \frac{1}{\sigma^2} \sum_{t=2}^T (y_t - \phi_1 y_{t-1}) y_{t-1} = 0$$

The MLE of  $\phi_1$  and  $\sigma^2$  satisfies the above two equation.

$$\begin{aligned}\tilde{\sigma}^2 &= \frac{1}{T} \left( (1 - \tilde{\phi}_1^2) y_1^2 + \sum_{t=2}^T (y_t - \tilde{\phi}_1 y_{t-1})^2 \right) \\ \tilde{\phi}_1 &= \frac{\sum_{t=2}^T y_t y_{t-1}}{\sum_{t=2}^T y_{t-1}^2} + \left( \tilde{\phi}_1 y_1^2 - \frac{\tilde{\sigma}^2 \tilde{\phi}_1}{1 - \tilde{\phi}_1^2} \right) / \sum_{t=2}^T y_{t-1}^2\end{aligned}$$

(b) Ordinary Least Squares (OLS) Method

$$S(\phi_1) = \sum_{t=2}^T (y_t - \phi_1 y_{t-1})^2$$

is minimized with respect to  $\phi_1$ .

$$\begin{aligned}\hat{\phi}_1 &= \frac{\sum_{t=2}^T y_{t-1} y_t}{\sum_{t=2}^T y_{t-1}^2} = \phi_1 + \frac{\sum_{t=2}^T y_{t-1} \epsilon_t}{\sum_{t=2}^T y_{t-1}^2} = \phi_1 + \frac{(1/T) \sum_{t=2}^T y_{t-1} \epsilon_t}{(1/T) \sum_{t=2}^T y_{t-1}^2} \\ &\longrightarrow \phi_1 + \frac{E(y_{t-1} \epsilon_t)}{E(y_{t-1}^2)} = \phi_1\end{aligned}$$

OLSE of  $\phi_1$  is a consistent estimator.

The following equations are utilized.

$$E(y_{t-1} \epsilon_t) = 0$$

$$E(y_{t-1}^2) = \text{Var}(y_{t-1}) = \gamma(0)$$

8. Asymptotic distribution of OLSE  $\hat{\phi}_1$ :

$$\sqrt{T}(\hat{\phi}_1 - \phi_1) \longrightarrow N(0, 1 - \phi_1^2)$$

**Proof:**

$y_{t-1}\epsilon_t$ ,  $t = 1, 2, \dots, T$ , are distributed with mean zero and variance  $\frac{\sigma_\epsilon^4}{1 - \phi_1^2}$ .

From the central limit theorem,

$$\frac{(1/T) \sum_{t=1}^T y_{t-1}\epsilon_t}{\sqrt{\sigma_\epsilon^4/(1 - \phi_1^2)/\sqrt{T}}} \longrightarrow N(0, 1)$$

Rewriting,

$$\frac{1}{\sqrt{T}} \sum_{t=1}^T y_{t-1}\epsilon_t \longrightarrow N(0, \frac{\sigma_\epsilon^4}{1 - \phi_1^2}).$$

Next,

$$\frac{1}{T} \sum_{t=1}^T y_{t-1}^2 \longrightarrow E(y_{t-1}^2) = \gamma(0) = \frac{\sigma_\epsilon^2}{1 - \phi_1^2}$$

yields:

$$\sqrt{T}(\hat{\phi}_1 - \phi_1) = \frac{(1/\sqrt{T}) \sum_{t=1}^T y_{t-1} \epsilon_t}{(1/T) \sum_{t=1}^T y_{t-1}^2} \rightarrow N(0, 1 - \phi_1^2)$$

9. Some formulas:

(a) Central Limit Theorem

Random variables  $x_1, x_2, \dots, x_T$  are mutually independently distributed with mean  $\mu$  and variance  $\sigma^2$ .

Define  $\bar{x} = (1/T) \sum_{t=1}^T x_t$ .

Then,

$$\frac{\bar{x} - E(\bar{x})}{\sqrt{V(\bar{x})}} = \frac{\bar{x} - \mu}{\sigma / \sqrt{T}} \rightarrow N(0, 1)$$

(b) Central Limit Theorem II

Random variables  $x_1, x_2, \dots, x_T$  are distributed with mean  $\mu$  and variance  $\sigma^2$ .

Define  $\bar{x} = (1/T) \sum_{t=1}^T x_t$ .

Then,

$$\frac{\bar{x} - E(\bar{x})}{\sqrt{V(\bar{x})}} \rightarrow N(0, 1)$$

(c) Let  $x$  and  $y$  be random variables.

$y$  converges in distribution to a distribution, and  $x$  converges in probability to a fixed value.

Then,  $xy$  converges in distribution.

For example, consider:

$$y \rightarrow N(\mu, \sigma^2), \quad x \rightarrow c.$$

Then, we obtain:

$$xy \rightarrow N(c\mu, c^2\sigma^2)$$

10. **AR(1) +drift:**  $y_t = \mu + \phi_1 y_{t-1} + \epsilon_t$

Mean:

Using the lag operator,

$$\phi(L)y_t = \mu + \epsilon_t$$

where  $\phi(L) = 1 - \phi_1 L$ .

Multiply  $\phi(L)^{-1}$  on both sides. Then, when  $|\phi_1| < 1$ , we have:

$$y_t = \phi(L)^{-1}\mu + \phi(L)^{-1}\epsilon_t.$$

Taking the expectation on both sides,

$$\begin{aligned} E(y_t) &= \phi(L)^{-1}\mu + \phi(L)^{-1}E(\epsilon_t) \\ &= \phi(1)^{-1}\mu = \frac{\mu}{1 - \phi_1} \end{aligned}$$

**Example: AR(2) Model:** Consider  $y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \epsilon_t$ .

1. The stationarity condition is: two solutions of  $x$  from  $\phi(x) = 1 - \phi_1 x - \phi_2 x^2 = 0$  are outside the unit circle.
2. Rewriting the AR(2) model,

$$(1 - \phi_1 L - \phi_2 L^2)y_t = \epsilon_t.$$

Let  $1/\alpha_1$  and  $1/\alpha_2$  be the solutions of  $\phi(x) = 0$ .

Then, the AR(2) model is written as:

$$(1 - \alpha_1 L)(1 - \alpha_2 L)y_t = \epsilon_t,$$

which is rewritten as:

$$y_t = \frac{1}{(1 - \alpha_1 L)(1 - \alpha_2 L)}\epsilon_t$$

$$= \left( \frac{\alpha_1/(\alpha_1 - \alpha_2)}{1 - \alpha_1 L} + \frac{-\alpha_2/(\alpha_1 - \alpha_2)}{1 - \alpha_2 L} \right) \epsilon_t$$

### 3. Mean of AR(2) Model:

When  $y_t$  is stationary, i.e.,  $\alpha_1$  and  $\alpha_2$  are within the unit circle,

$$\mu = E(y_t) = E(\phi(L)\epsilon_t) = 0$$

### 4. Autocovariance Function of AR(2) Model:

$$\begin{aligned}\gamma(\tau) &= E((y_t - \mu)(y_{t-\tau} - \mu)) = E(y_t y_{t-\tau}) \\ &= E((\phi_1 y_{t-1} + \phi_2 y_{t-2} + \epsilon_t) y_{t-\tau}) \\ &= \phi_1 E(y_{t-1} y_{t-\tau}) + \phi_2 E(y_{t-2} y_{t-\tau}) + E(\epsilon_t y_{t-\tau}) \\ &= \begin{cases} \phi_1 \gamma(\tau - 1) + \phi_2 \gamma(\tau - 2), & \text{for } \tau \neq 0, \\ \phi_1 \gamma(\tau - 1) + \phi_2 \gamma(\tau - 2) + \sigma_\epsilon^2, & \text{for } \tau = 0. \end{cases}\end{aligned}$$

The initial condition is obtained by solving the following three equations:

$$\gamma(0) = \phi_1\gamma(1) + \phi_2\gamma(2) + \sigma_\epsilon^2,$$

$$\gamma(1) = \phi_1\gamma(0) + \phi_2\gamma(1),$$

$$\gamma(2) = \phi_1\gamma(1) + \phi_2\gamma(0).$$

Therefore, the initial conditions are given by:

$$\gamma(0) = \left(\frac{1 - \phi_2}{1 + \phi_2}\right) \frac{\sigma_\epsilon^2}{(1 - \phi_2)^2 - \phi_1^2},$$

$$\gamma(1) = \frac{\phi_1}{1 - \phi_2} \gamma(0) = \left(\frac{\phi_1}{1 - \phi_2}\right) \left(\frac{1 - \phi_2}{1 + \phi_2}\right) \frac{\sigma_\epsilon^2}{(1 - \phi_2)^2 - \phi_1^2}.$$

Given  $\gamma(0)$  and  $\gamma(1)$ , we obtain  $\gamma(\tau)$  as follows:

$$\gamma(\tau) = \phi_1\gamma(\tau - 1) + \phi_2\gamma(\tau - 2), \quad \text{for } \tau = 2, 3, \dots$$

## 5. Another solution for $\gamma(0)$ :

From  $\gamma(0) = \phi_1\gamma(1) + \phi_2\gamma(2) + \sigma_\epsilon^2$ ,

$$\gamma(0) = \frac{\sigma_\epsilon^2}{1 - \phi_1\rho(1) - \phi_2\rho(2)}$$

where

$$\rho(1) = \frac{\phi_1}{1 - \phi_2}, \quad \rho(2) = \phi_1\rho(1) + \phi_2 = \frac{\phi_1^2 + (1 - \phi_2)\phi_2}{1 - \phi_2}.$$

## 6. Autocorrelation Function of AR(2) Model:

Given  $\rho(1)$  and  $\rho(2)$ ,

$$\rho(\tau) = \phi_1\rho(\tau - 1) + \phi_2\rho(\tau - 2), \quad \text{for } \tau = 3, 4, \dots,$$

7.  $\phi_{k,k}$  = Partial Autocorrelation Coefficient of AR(2) Process:

$$\begin{pmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(k-1) \\ \rho(1) & 1 & & \rho(k-3) & \rho(k-2) \\ \vdots & \vdots & & \vdots & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & 1 \end{pmatrix} \begin{pmatrix} \phi_{k,1} \\ \phi_{k,2} \\ \vdots \\ \phi_{k,k-1} \\ \phi_{k,k} \end{pmatrix} = \begin{pmatrix} \rho(1) \\ \rho(2) \\ \vdots \\ \rho(k) \end{pmatrix},$$

for  $k = 1, 2, \dots$ .

$$\phi_{k,k} = \frac{\begin{vmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(1) \\ \rho(1) & 1 & & & \rho(k-3) \rho(2) \\ \vdots & \vdots & & & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & \rho(k) \end{vmatrix}}{\begin{vmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(k-1) \\ \rho(1) & 1 & & & \rho(k-3) \rho(k-2) \\ \vdots & \vdots & & & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & 1 \end{vmatrix}}$$

Autocovariance Functions:

$$\gamma(1) = \phi_1\gamma(0) + \phi_2\gamma(1),$$

$$\gamma(2) = \phi_1\gamma(1) + \phi_2\gamma(0),$$

$$\gamma(\tau) = \phi_1\gamma(\tau - 1) + \phi_2\gamma(\tau - 2), \quad \text{for } \tau = 3, 4, \dots$$

Autocorrelation Functions:

$$\rho(1) = \phi_1 + \phi_2\rho(1) = \frac{\phi_1}{1 - \phi_2},$$

$$\rho(2) = \phi_1\rho(1) + \phi_2 = \frac{\phi_1^2}{1 - \phi_2} + \phi_2,$$

$$\rho(\tau) = \phi_1\rho(\tau - 1) + \phi_2\rho(\tau - 2), \quad \text{for } \tau = 3, 4, \dots$$

$$\phi_{1,1} = \rho(1) = \frac{\phi_1}{1 - \phi_2}$$

$$\phi_{2,2} = \frac{\begin{vmatrix} 1 & \rho(1) \\ \rho(1) & \rho(2) \\ 1 & \rho(1) \end{vmatrix}}{\begin{vmatrix} \rho(1) & 1 \end{vmatrix}} = \frac{\rho(2) - \rho(1)^2}{1 - \rho(1)^2} = \phi_2$$

$$\phi_{3,3} = \frac{\begin{vmatrix} 1 & \rho(1) & \rho(1) \\ \rho(1) & 1 & \rho(2) \\ \rho(2) & \rho(1) & \rho(3) \end{vmatrix}}{\begin{vmatrix} 1 & \rho(1) & \rho(2) \\ \rho(1) & 1 & \rho(1) \\ \rho(2) & \rho(1) & 1 \end{vmatrix}}$$

$$= \frac{(\rho(3) - \rho(1)\rho(2)) - \rho(1)^2(\rho(3) - \rho(1)) + \rho(2)\rho(1)(\rho(2) - 1)}{(1 - \rho(1)^2) - \rho(1)^2(1 - \rho(2)) + \rho(2)(\rho(1)^2 - \rho(2))} = 0.$$

## 8. Log-Likelihood Function — Innovation Form:

$$\log f(y_T, \dots, y_1) = \log f(y_2, y_1) + \sum_{t=3}^T \log f(y_t | y_{t-1}, \dots, y_1)$$

where

$$f(y_2, y_1) = \frac{1}{2\pi} \begin{vmatrix} \gamma(0) & \gamma(1) \\ \gamma(1) & \gamma(0) \end{vmatrix}^{-1/2} \exp\left(-\frac{1}{2}(y_1 \ y_2) \begin{pmatrix} \gamma(0) & \gamma(1) \\ \gamma(1) & \gamma(0) \end{pmatrix}^{-1} \begin{pmatrix} y_1 \\ y_2 \end{pmatrix}\right),$$

$$f(y_t | y_{t-1}, \dots, y_1) = \frac{1}{\sqrt{2\pi\sigma_\epsilon^2}} \exp\left(-\frac{1}{2\sigma_\epsilon^2}(y_t - \phi_1 y_{t-1} - \phi_2 y_{t-2})^2\right).$$

Note as follows:

$$\begin{pmatrix} \gamma(0) & \gamma(1) \\ \gamma(1) & \gamma(0) \end{pmatrix} = \gamma(0) \begin{pmatrix} 1 & \rho(1) \\ \rho(1) & 1 \end{pmatrix} = \gamma(0) \begin{pmatrix} 1 & \phi_1/(1-\phi_2) \\ \phi_1/(1-\phi_2) & 1 \end{pmatrix}.$$

9. **AR(2) +drift:**  $y_t = \mu + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \epsilon_t$

Mean:

Rewriting the AR(2)+drift model,

$$\phi(L)y_t = \mu + \epsilon_t$$

where  $\phi(L) = 1 - \phi_1 L - \phi_2 L^2$ .

Under the stationarity assumption, we can rewrite the AR(2)+drift model as follows:

$$y_t = \phi(L)^{-1}\mu + \phi(L)^{-1}\epsilon_t.$$

Therefore,

$$E(y_t) = \phi(L)^{-1}\mu + \phi(L)^{-1}E(\epsilon_t) = \phi(1)^{-1}\mu = \frac{\mu}{1 - \phi_1 - \phi_2}$$