3.2.2 Random Effect Model (ランダム効果モデル)

Model:

$$y_{it} = X_{it}\beta + v_i + u_{it}, \qquad i = 1, 2, \cdots, n, \quad t = 1, 2, \cdots, T$$

where *i* indicates individual and *t* denotes time.

The assumptions on the error terms v_i and u_{it} are:

$$E(v_i|X) = E(u_{it}|X) = 0 \text{ for all } i,$$

$$V(v_i|X) = \sigma_v^2 \text{ for all } i, \qquad V(u_{it}|X) = \sigma_u^2 \text{ for all } i \text{ and } t,$$

$$Cov(v_i, v_j|X) = 0 \text{ for } i \neq j, \qquad Cov(u_{it}, u_{js}|X) = 0 \text{ for } i \neq j \text{ and } t \neq s,$$

$$Cov(v_i, u_{jt}|X) = 0 \text{ for all } i, j \text{ and } t.$$

Note that X includes X_{it} for $i = 1, 2, \dots, n$ and $t = 1, 2, \dots, T$.

In a matrix form with respect to $t = 1, 2, \dots, T$, we have the following:

$$y_{i} = X_{i}\beta + v_{i}1_{T} + u_{i}, \qquad i = 1, 2, \cdots, n,$$
where $y_{i} = \begin{pmatrix} y_{i1} \\ y_{i2} \\ \vdots \\ y_{iT} \end{pmatrix}, X_{i} = \begin{pmatrix} X_{i1} \\ X_{i2} \\ \vdots \\ X_{iT} \end{pmatrix} \text{ and } u_{i} = \begin{pmatrix} u_{i1} \\ u_{i2} \\ \vdots \\ u_{iT} \end{pmatrix} \text{ are } T \times 1, T \times k \text{ and } T \times 1, \text{ respectively.}$

$$u_{i} \sim N(0, \sigma_{u}^{2}I_{T}) \text{ and } v_{i}1_{T} \sim N(0, \sigma_{v}^{2}) \implies v_{i}1_{T} + u_{i} \sim N(0, \sigma_{v}^{2}I_{T}I_{T}' + \sigma_{u}^{2}I_{T}).$$

Again, in a matrix form with respect to *i*, we have the following:

where
$$y = \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix}$$
, $X = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{pmatrix}$, $v = \begin{pmatrix} v_1 1_T \\ v_2 1_T \\ \vdots \\ v_n 1_T \end{pmatrix}$ and $u = \begin{pmatrix} u_1 \\ u_2 \\ \vdots \\ u_n \end{pmatrix}$ are $nT \times 1$, $nT \times k$, $nT \times 1$ and

 $nT \times 1$, respectively.

The distribution of u + v is given by:

$$v + u \sim N(0, I_n \otimes (\sigma_v^2 \mathbf{1}_T \mathbf{1}_T' + \sigma_u^2 I_T))$$

The likelihood function is given by:

$$L(\beta, \sigma_{\nu}^{2}, \sigma_{u}^{2}) = (2\pi)^{-nT/2} \left| I_{n} \otimes (\sigma_{\nu}^{2} \mathbf{1}_{T} \mathbf{1}_{T}' + \sigma_{u}^{2} I_{T}) \right|^{-1/2} \\ \times \exp\left(-\frac{1}{2}(y - X\beta)' \left(I_{n} \otimes (\sigma_{\nu}^{2} \mathbf{1}_{T} \mathbf{1}_{T}' + \sigma_{u}^{2} I_{T}) \right)^{-1} (y - X\beta) \right).$$

Remember that $f(x) = (2\pi)^{-k/2} |\Sigma|^{-1/2} \exp\left(-\frac{1}{2}(x-\mu)'\Sigma^{-1}(x-\mu)\right)$ when $X \sim N(\mu, \Sigma)$, where X denotes a k-variate random variable.

The estimators of β , σ_v^2 and σ_u^2 are given by maximizing the following log-likelihood function:

$$\log L(\beta, \sigma_v^2, \sigma_u^2) = -\frac{nT}{2} \log(2\pi) - \frac{1}{2} \log \left| I_n \otimes (\sigma_v^2 \mathbf{1}_T \mathbf{1}_T' + \sigma_u^2 I_T) \right| -\frac{1}{2} (y - X\beta)' \left(I_n \otimes (\sigma_v^2 \mathbf{1}_T \mathbf{1}_T' + \sigma_u^2 I_T) \right)^{-1} (y - X\beta).$$

MLE of β , denoted by $\tilde{\beta}$, is given by:

$$\begin{split} \tilde{\beta} &= \left(X' \Big(I_n \otimes (\sigma_v^2 \mathbf{1}_T \mathbf{1}'_T + \sigma_u^2 I_T) \Big)^{-1} X \Big)^{-1} X' \Big(I_n \otimes (\sigma_v^2 \mathbf{1}_T \mathbf{1}'_T + \sigma_u^2 I_T) \Big)^{-1} y \\ &= \Big(\sum_{i=1}^n X'_i (\sigma_v^2 \mathbf{1}_T \mathbf{1}'_T + \sigma_u^2 I_T)^{-1} X_i \Big)^{-1} \Big(\sum_{i=1}^n X'_i (\sigma_v^2 \mathbf{1}_T \mathbf{1}'_T + \sigma_u^2 I_T)^{-1} y_i \Big), \end{split}$$

which is equivalent to GLS.

Note that $\tilde{\beta}$ is not operational, because $\hat{\beta}$ depends on σ_v^2 and σ_u^2 .

3.3 Hausman's Specification Error (特定化誤差) Test

Regression model:

 $y = X\beta + u$, $y : n \times 1$, $X : n \times k$, $\beta : k \times 1$, $u : n \times 1$.

Suppose that *X* is stochastic.

If E(u|X) = 0, OLSE $\hat{\beta}$ is unbiased because of $\hat{\beta} = (X'X)^{-1}X'y = \beta + (X'X)^{-1}X'u$ and $E((X'X)^{-1}X'u) = 0.$

However, If $E(u|X) \neq 0$, OLSE $\hat{\beta}$ is biased and inconsistent.

Therefore, we need to check if X is correlated with u or not.

\implies Hausman's Specification Error Test

The null and alternative hypotheses are:

- H_0 : X and u are independent, i.e., Cov(X, u) = 0,
- H_1 : X and u are not independent.

Suppose that we have two estimators $\hat{\beta}_0$ and $\hat{\beta}_1$, which have the following properties:

- $\hat{\beta}_0$ is consistent and efficient under H_0 , but is not consistent under H_1 ,
- $\hat{\beta}_1$ is consistent under both H_0 and H_1 , but is not efficient under H_0 .

Under the conditions above, we have the following test statistic:

$$(\hat{\beta}_1 - \hat{\beta}_0)' \Big(\mathbf{V}(\hat{\beta}_1) - \mathbf{V}(\hat{\beta}_0) \Big)^{-1} (\hat{\beta}_1 - \hat{\beta}_0) \longrightarrow \chi^2(k).$$

Example: $\hat{\beta}_0$ is OLS, while $\hat{\beta}_1$ is IV such as 2SLS.

Hausman, J.A. (1978) "Specification Tests in Econometrics," *Econometrica*, Vol.46, No.6, pp.1251–1271.

3.4 Choice of Fixed Effect Model or Random Effect Model

3.4.1 The Case where *X* is Correlated with *u* — Review

The standard regression model is given by:

$$y = X\beta + u, \qquad u \sim N(0, \sigma^2 I_n)$$

OLS is:

$$\hat{\beta} = (X'X)^{-1}X'y = \beta + (X'X)^{-1}X'u.$$

If *X* is not correlated with *u*, i.e., E(X'u) = 0, we have the result: $E(\hat{\beta}) = \beta$.

However, if X is correlated with u, i.e., $E(X'u) \neq 0$, we have the result: $E(\hat{\beta}) \neq \beta$. $\implies \hat{\beta}$ is biased. Assume that in the limit we have the followings:

$$(\frac{1}{n}X'X)^{-1} \longrightarrow M_{xx}^{-1},$$

$$\frac{1}{n}X'u \longrightarrow M_{xu} \neq 0 \text{ when } X \text{ is correlated with } u.$$

Therefore, even in the limit,

$$\operatorname{plim}\hat{\beta} = \beta + M_{xx}^{-1}M_{xu} \neq \beta,$$

which implies that $\hat{\beta}$ is not a consistent estimator of β .

Thus, in the case where X is correlated with u, OLSE $\hat{\beta}$ is neither unbiased nor consistent.

3.4.2 Fixed Effect Model or Random Effect Model

Usually, in the random effect model, we can consider that v_i is correlated with X_{it} .

[Reason:]

 v_i includes the unobserved variables in the *i*th individual, i.e., ability, intelligence, and so on.

 X_{it} represents the observed variables in the *i*th individual, i.e., income, assets, and so on.

The unobserved variables v_i are related to the observed variables X_{it} .

Therefore, we consider that v_i is correlated with X_{it} .

Thus, in the case of the random effect model, usually we cannot use OLS or GLS. In order to use the random effect model, we need to test whether v_i is uncorrelated with X_{it} . Apply Hausman's test.

- H_0 : X_{it} and e_{it} are independent (\longrightarrow Use the random effect model),
- H_1 : X_{it} and e_{it} are not independent (\longrightarrow Use the fixed effect model),

where $e_{it} = v_i + u_{it}$.

Note that:

- We can use the random effect model under H_0 , but not under H_1 .
- We can use the fixed effect model under both H_0 and H_1 .
- The random effect model is more efficient than the fixed effect model under H_0 . Therefore, under H_0 we should use the random effect model, rather than the fixed effect model.