7.11 ARCH and GARCH Models

Autoregressive Conditional Heteroskedasticity (ARCH)

Generalized Autoregressive Conditional Heteroskedasticity (GARCH)

1. ARCH (p) Model

   $$\epsilon_t|\epsilon_{t-1}, \epsilon_{t-2}, \cdots, \epsilon_1 \sim N(0, h_t),$$

   where,

   $$h_t = \alpha_0 + \alpha_1\epsilon_{t-1}^2 + \cdots + \alpha_p\epsilon_{t-p}^2.$$

   The unconditional variance of ϵ_t is:

   $$\sigma_\epsilon^2 = \frac{\alpha_0}{1 - \alpha_1 - \alpha_2 - \cdots - \alpha_p}$$

2. GARCH (p, q) Model

   $$\epsilon_t|\epsilon_{t-1}, \epsilon_{t-2}, \cdots, \epsilon_1 \sim N(0, h_t),$$

where

$$h_t = \alpha_0 + \alpha_1 \epsilon_{t-1}^2 + \cdots + \alpha_p \epsilon_{t-p}^2 + \beta_1 h_{t-1} + \cdots + \beta_q h_{t-q}.$$

3. Application to OLS (Case of ARCH(1) Model):

$$y_t = x_t \beta + \epsilon_t, \qquad \epsilon_t | \epsilon_{t-1}, \epsilon_{t-2}, \cdots, \epsilon_1 \sim N(0, \alpha_0 + \alpha_1 \epsilon_{t-1}^2).$$

The joint density of $\epsilon_1, \epsilon_2, \cdots, \epsilon_T$ is:

$$\begin{aligned} f(\epsilon_1, \cdots, \epsilon_T) &= f(\epsilon_1) \prod_{t=2}^{T} f(\epsilon_t | \epsilon_{t-1}, \cdots, \epsilon_1) \\ &= (2\pi)^{-1/2} \left(\frac{\alpha_0}{1-\alpha_1} \right)^{-1/2} \exp\left(-\frac{1}{2\alpha_0/(1-\alpha_1)} \epsilon_1^2 \right) \\ &\times (2\pi)^{-(T-1)/2} \prod_{t=2}^{T} (\alpha_0 + \alpha_1 \epsilon_{t-1}^2)^{-1/2} \exp\left(-\frac{1}{2} \sum_{t=2}^{T} \frac{\epsilon_t^2}{\alpha_0 + \alpha_1 \epsilon_{t-1}^2} \right). \end{aligned}$$

The log-likelihood function is:

$$
\begin{aligned}
&\log L(\beta, \alpha_0, \alpha_1; y_1, \cdots, y_T) \\
&\quad = -\frac{1}{2}\log(2\pi) - \frac{1}{2}\log\left(\frac{\alpha_0}{1-\alpha_1}\right) - \frac{1}{2\alpha_0/(1-\alpha_1)}(y_1 - x_1\beta)^2 \\
&\qquad -\frac{T-1}{2}\log(2\pi) - \frac{1}{2}\sum_{t=2}^{T}\log\left(\alpha_0 + \alpha_1(y_{t-1} - x_{t-1}\beta)^2\right) \\
&\qquad -\frac{1}{2}\sum_{t=2}^{T}\frac{(y_t - x_t\beta)^2}{\alpha_0 + \alpha_1(y_{t-1} - x_{t-1}\beta)^2}.
\end{aligned}
$$

Obtain α_0, α_1 and β such that the log-likelihood function is maximized.

$\alpha_0 > 0$ and $\alpha_1 > 0$ have to be satisfied.

These two conditions are explicitly included, when the model is modified to:

$\mathrm{E}(\epsilon_t^2 | \epsilon_{t-1}, \epsilon_{t-2}, \cdots, \epsilon_1) = \alpha_0^2 + \alpha_1^2 \epsilon_{t-1}^2$.

Testing the ARCH(1) Effect:

(a) Estimate $y_t = x_t\beta + u_t$ by OLS, and compute $\hat{\beta}$ and $\hat{u}_t = y_t - x_t\hat{\beta}$.

(b) Estimate $\hat{u}_t^2 = \alpha_0 + \alpha_1\hat{u}_{t-1}^2$ by OLS. If $\hat{\alpha}_1$ is significant, there is the ARCH(1) effect in the error term.

This test corresponds to LM test.

Example: GARCH(1,1) Model

```
. arch sdex l.sdex l2.sdex, arch(1) garch(1)

(setting optimization to BHHH)
Iteration 0:   log likelihood = -5089.3558
Iteration 1:   log likelihood = -5086.7468
.........
.........

Iteration 22:  log likelihood = -5064.9328  (backed up)
Iteration 23:  log likelihood = -5064.9328

ARCH family regression
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Sample: 16 - 516
Distribution: Gaussian
Log likelihood = -5064.933

Number of obs = 501
Wald chi2(2) = 225.19
Prob > chi2 = 0.0000

sdex	Coef.	OPG Std. Err.	z	P>|z|	[95% Conf.	Interval]
sdex						
sdex						
L1.	-.6357273	.0426939	-14.89	0.000	-.7194059	-.5520488
L2.	-.370862	.0466222	-7.95	0.000	-.4622398	-.2794842
_cons	-55.28043	261.2057	-0.21	0.832	-567.2341	456.6733
ARCH						
arch						
L1.	.041632	.0123474	3.37	0.001	.0174317	.0658324
garch						
L1.	.9526041	.0148639	64.09	0.000	.9234715	.9817367
_cons	312143.8	227564.3	1.37	0.170	-133873.9	758161.6

8 Vector Autoregressive (VAR) Model – Causality, Impulse Response Function and etc

Vector Autoregressive Process:

$$y_t = \mu + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \epsilon_t,$$

where

$$y_t : k \times 1, \qquad \mu : k \times 1, \qquad \epsilon_t : k \times 1, \qquad \phi_i : k \times k.$$

Rewriting the above equation,

$$\phi(L) y_t = \mu + \epsilon_t,$$

where $\phi(L) = I_k - \phi_1 L - \phi_2 L^2 - \cdots - \phi_p L^p$.

VAR(1) Model:

$$y_t = \phi_1 y_{t-1} + \epsilon_t, \qquad \text{i.e.,} \qquad (I_k - \phi_1 L) y_t = \epsilon_t.$$

When y_t is stationary, we obtain:

$$\begin{aligned} y_t &= (I_k - \phi_1 L)^{-1} \epsilon_t \\ &= (I_k + \phi_1 L + \phi_1^2 L^2 + \phi_1^3 L^3 + \cdots) \epsilon_t \\ &= \epsilon_t + \phi_1 \epsilon_{t-1} + \phi_1^2 \epsilon_{t-2} + \phi_1^3 \epsilon_{t-3} + \cdots \end{aligned}$$

VAR(1)=VMA(∞)

VAR(2) Model:

$$y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \epsilon_t, \qquad \text{i.e.,} \qquad (I_k - \phi_1 L - \phi_2 L^2) y_{t-1} = \epsilon_t.$$

When y_t is stationary, we obtain:

$$\begin{aligned} y_{t-1} &= (I_k - \phi_1 L - \phi_2 L^2)^{-1}\epsilon_t \\ &= \epsilon_t + \theta_1\epsilon_{t-1} + \theta_2\epsilon_{t-2} + \cdots \end{aligned}$$

VAR(2)=VMA(∞)

VAR(p) Model:

$$y_t = \mu + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \epsilon_t,$$

i.e.,

$$(I_k - \phi_1 L - \phi_2 L^2 - \cdots - \phi_p L^p) y_{t-1} = \epsilon_t.$$

When y_t is stationary, we obtain:

$$\begin{aligned} y_t &= (I_k - \phi_1 L - \phi_2 L^2 - \cdots - \phi_p L^p)^{-1}\epsilon_t \\ &= \epsilon_t + \theta_1\epsilon_{t-1} + \theta_2\epsilon_{t-2} + \cdots \end{aligned}$$

VAR(p)=VMA(∞)

8.1 Autocovariance Matrix and Autocorrelation Matrix

Let y_t be a $k \times 1$ vector.

Autocovariance Function Matrix:

$$\Gamma(\tau) = \mathrm{E}((y_t - \mu)(y_{t-\tau} - \mu)'), \qquad \tau = 0, 1, 2, \cdots,$$

where $\mathrm{E}(y_t) = \mu$. $\Gamma(\tau)$ is a $k \times k$ matrix.

$$\Gamma(\tau) = \Gamma(-\tau)'$$

Autocorrelation Function Matrix:

$$\rho(\tau) = D^{-1/2}\Gamma(\tau)D^{-1/2},$$

where the (i, j)th element of D is given by $\gamma_{ii}(\tau) = \mathrm{V}(y_{it})$ for $i = j$ and zero otherwise.

$$\rho(\tau) = \rho(-\tau)'$$

8.2 Granger Cuasality Test (グレンジャー因果性テスト)

Consider the bivariate case:

$$\begin{pmatrix} y_{1,t} \\ y_{2,t} \end{pmatrix} = \begin{pmatrix} \mu_1 \\ \mu_2 \end{pmatrix} + \begin{pmatrix} \phi_{11,1} & \phi_{12,1} \\ \phi_{21,1} & \phi_{22,1} \end{pmatrix} \begin{pmatrix} y_{1,t-1} \\ y_{2,t-1} \end{pmatrix} + \cdots + \begin{pmatrix} \phi_{11,p} & \phi_{12,p} \\ \phi_{21,p} & \phi_{22,p} \end{pmatrix} \begin{pmatrix} y_{1,t-p} \\ y_{2,t-p} \end{pmatrix} + \begin{pmatrix} \epsilon_1 \\ \epsilon_2 \end{pmatrix}$$

Unrestricted Model (Sum of Squared Residuals, denoted by SSR_1):

$$y_{1,t} = \mu_1 + \begin{pmatrix} \phi_{11,1} & \phi_{12,1} \end{pmatrix} \begin{pmatrix} y_{1,t-1} \\ y_{2,t-1} \end{pmatrix} + \cdots + \begin{pmatrix} \phi_{11,p} & \phi_{12,p} \end{pmatrix} \begin{pmatrix} y_{1,t-p} \\ y_{2,t-p} \end{pmatrix} + \epsilon_1$$

$H_0: \ \phi_{12,1} = \phi_{12,2} = \cdots = \phi_{12,p} = 0$

When H_0 is correct, we say there is no causality from y_2 to y_1.

$\Longrightarrow$ Granger Causality Test.

Restricted Model (Sum of Squared Residuals, denoted by SSR_0):

$$y_{1,t} = \mu_1 + \begin{pmatrix} \phi_{11,1} & 0 \end{pmatrix} \begin{pmatrix} y_{1,t-1} \\ y_{2,t-1} \end{pmatrix} + \cdots + \begin{pmatrix} \phi_{11,p} & 0 \end{pmatrix} \begin{pmatrix} y_{1,t-p} \\ y_{2,t-p} \end{pmatrix} + \epsilon_1$$

Asymptotically, we have the following distribution:

$$F = \frac{(\mathrm{SSR}_0 - \mathrm{SSR}_1)/p}{\mathrm{SSR}_1/(T-2p-1)} \sim F(p, T-2p-1),$$

or

$$pF \sim \chi^2(p).$$