

Econometrics I: Solutions of the homework #10

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1 Question 1

1.1 (1)

The likelihood function is defined as $L(\theta; X) = \prod_{i=1}^n f(X_i; \theta)$ in our case:

$$L(\theta; X) = (2\pi\sigma^2)^{-n/2} \exp\left(-\frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - u)^2\right)$$

The maximum likelihood estimate of θ is the θ such that:

$$\max_{\theta} L(\theta; X) \iff \max_{\theta} \log L(\theta; X)$$

*If you have any errors in handouts and materials, please contact me via lvang12@hotmail.com

Therefore we can take log for both sides and the log likelihood function can be written as:

$$\log L(\theta; X) = -\frac{n}{2} \log(2\pi) - \frac{n}{2} \log(\sigma^2) - \frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - u)^2$$

1.2 (2)

Obtain the first order condition:

$$\begin{aligned} \frac{\partial \log L(\theta; X)}{\partial u} &= \frac{1}{\sigma^2} \sum_{i=1}^n (X_i - u) = 0 \\ \frac{\partial \log L(\theta; X)}{\partial \sigma^2} &= -\frac{n}{2} \frac{1}{\sigma^2} + \frac{1}{2\sigma^4} \sum_{i=1}^n (X_i - u)^2 = 0 \end{aligned}$$

Solving the equation we can obtain the MLE estimators of u and σ as:

$$\begin{aligned} \hat{u} &= \frac{1}{n} \sum_{i=1}^n X_i \\ \hat{\sigma}^2 &= \frac{1}{n} \sum_{i=1}^n (X_i - \hat{u})^2 \end{aligned}$$

1.3 (3)

the variance of MLE estimators are calculated as:

$$V(\hat{\mu}) = \frac{1}{n^2} \sum_{i=1}^n V(X_i) = \frac{\sigma^2}{n}$$

$$\begin{aligned} V(\hat{\sigma}^2) &= V\left(\frac{1}{n} \sum_{i=1}^n (X_i - \hat{\mu})^2\right) \\ &= V\left(\frac{\sigma^2}{n\sigma^2} \sum_{i=1}^n (X_i - \hat{\mu})^2\right) \\ &= \frac{\sigma^4}{n^2} V\left(\sum_{i=1}^n \left(\frac{X_i - \hat{\mu}}{\sigma}\right)^2\right) \end{aligned}$$

Recall that $\sum_{i=1}^n \left(\frac{X_i - \hat{\mu}}{\sigma}\right)^2 \sim \chi^2(n-1)$ which has the variance $2(n-1)$ therefore:

$$V(\hat{\sigma}^2) = \frac{2\sigma^4(n-1)}{n^2}$$

1.4 (4)

Fisher's information matrix $I(\theta)$ is given as follows:

$$\begin{aligned}
 I(\theta) &= -E\left(\frac{\partial^2 \log L(\theta; X)}{\partial \theta \partial \theta'}\right) = -E\left(\begin{array}{cc} \frac{\partial^2 \log L(\theta; X)}{\partial u^2} & \frac{\partial^2 \log L(\theta; X)}{\partial u \partial \sigma^2} \\ \frac{\partial^2 \log L(\theta; X)}{\partial \sigma^2 \partial u} & \frac{\partial^2 \log L(\theta; X)}{\partial (\sigma^2)^2} \end{array}\right) \\
 &= -E\left(\begin{array}{cc} -\frac{n}{\sigma^2} & -\frac{n}{\sigma^4} \sum_{i=1}^n (X_i - u) \\ -\frac{n}{\sigma^4} \sum_{i=1}^n (X_i - u) & \frac{n}{2\sigma^4} - \frac{1}{\sigma^6} \sum_{i=1}^n (X_i - u)^2 \end{array}\right) \\
 &= \begin{pmatrix} \frac{n}{\sigma^2} & 0 \\ 0 & \frac{n}{2\sigma^4} \end{pmatrix}
 \end{aligned}$$

1.5 (5)

Compare (3) and (4) we can find when n is large the variance of MLE will approximate to $I(\theta)^{-1}$. i.e.

$$\begin{pmatrix} \hat{u} \\ \hat{\sigma}^2 \end{pmatrix} \sim N\left(\begin{pmatrix} u \\ \sigma^2 \end{pmatrix}, \begin{pmatrix} \frac{\sigma^2}{n} & 0 \\ 0 & \frac{2\sigma^4}{n} \end{pmatrix}\right)$$