

6.5 ARMA Model

ARMA (Autoregressive Moving Average , 自己回帰移動平均) Process

1. ARMA(p, q)

$$y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \epsilon_t + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2} + \cdots + \theta_q \epsilon_{t-q},$$

which is rewritten as:

$$\phi(L)y_t = \theta(L)\epsilon_t,$$

where $\phi(L) = 1 - \phi_1 L - \phi_2 L^2 - \cdots - \phi_p L^p$ and $\theta(L) = 1 + \theta_1 L + \theta_2 L^2 + \cdots + \theta_q L^q$.

2. Likelihood Function:

The variance-covariance matrix of Y , denoted by V , has to be computed.

Example: ARMA(1,1) Process: $y_t = \phi_1 y_{t-1} + \epsilon_t + \theta_1 \epsilon_{t-1}$

Obtain the autocorrelation coefficient.

The mean of y_t is to take the expectation on both sides.

$$E(y_t) = \phi_1 E(y_{t-1}) + E(\epsilon_t) + \theta_1 E(\epsilon_{t-1}),$$

where the second and third terms are zeros.

Therefore, we obtain:

$$E(y_t) = 0.$$

The autocovariance of y_t is to take the expectation, multiplying $y_{t-\tau}$ on both sides.

$$E(y_t y_{t-\tau}) = \phi_1 E(y_{t-1} y_{t-\tau}) + E(\epsilon_t y_{t-\tau}) + \theta_1 E(\epsilon_{t-1} y_{t-\tau}).$$

Each term is given by:

$$E(y_t y_{t-\tau}) = \gamma(\tau), \quad E(y_{t-1} y_{t-\tau}) = \gamma(\tau - 1),$$

$$E(\epsilon_t y_{t-\tau}) = \begin{cases} \sigma_\epsilon^2, & \tau = 0, \\ 0, & \tau = 1, 2, \dots, \end{cases} \quad E(\epsilon_{t-1} y_{t-\tau}) = \begin{cases} (\phi_1 + \theta_1)\sigma_\epsilon^2, & \tau = 0, \\ \sigma_\epsilon^2, & \tau = 1, \\ 0, & \tau = 2, 3, \dots \end{cases}$$

Therefore, we obtain;

$$\begin{aligned} \gamma(0) &= \phi_1 \gamma(1) + (1 + \phi_1 \theta_1 + \theta_1^2) \sigma_\epsilon^2, \\ \gamma(1) &= \phi_1 \gamma(0) + \theta_1 \sigma_\epsilon^2, \\ \gamma(\tau) &= \phi_1 \gamma(\tau - 1), \quad \tau = 2, 3, \dots \end{aligned}$$

From the first two equations, $\gamma(0)$ and $\gamma(1)$ are computed by:

$$\begin{pmatrix} 1 & -\phi_1 \\ -\phi_1 & 1 \end{pmatrix} \begin{pmatrix} \gamma(0) \\ \gamma(1) \end{pmatrix} = \sigma_\epsilon^2 \begin{pmatrix} 1 + \phi_1 \theta_1 + \theta_1^2 \\ \theta_1 \end{pmatrix}$$

$$\begin{pmatrix} \gamma(0) \\ \gamma(1) \end{pmatrix} = \sigma_\epsilon^2 \begin{pmatrix} 1 & -\phi_1 \\ -\phi_1 & 1 \end{pmatrix}^{-1} \begin{pmatrix} 1 + \phi_1 \theta_1 + \theta_1^2 \\ \theta_1 \end{pmatrix}$$

$$= \frac{\sigma_\epsilon^2}{1 - \phi_1^2} \begin{pmatrix} 1 & \phi_1 \\ \phi_1 & 1 \end{pmatrix} \begin{pmatrix} 1 + \phi_1\theta_1 + \theta_1^2 \\ \theta_1 \end{pmatrix} = \frac{\sigma_\epsilon^2}{1 - \phi_1^2} \begin{pmatrix} 1 + 2\phi_1\theta_1 + \theta_1^2 \\ (1 + \phi_1\theta_1)(\phi_1 + \theta_1) \end{pmatrix}.$$

Thus, the initial value of the autocorrelation coefficient is given by:

$$\rho(1) = \frac{(1 + \phi_1\theta_1)(\phi_1 + \theta_1)}{1 + 2\phi_1\theta_1 + \theta_1^2}.$$

We have:

$$\rho(\tau) = \phi_1\rho(\tau - 1).$$

ARMA(p, q) +drift:

$$y_t = \mu + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \epsilon_t + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2} + \cdots + \theta_q \epsilon_{t-q}.$$

Mean of ARMA(p, q) Process: $\phi(L)y_t = \mu + \theta(L)\epsilon_t$,

where $\phi(L) = 1 - \phi_1 L - \phi_2 L^2 - \cdots - \phi_p L^p$ and $\theta(L) = 1 + \theta_1 L + \theta_2 L^2 + \cdots + \theta_q L^q$.

$$y_t = \phi(L)^{-1} \mu + \phi(L)^{-1} \theta(L) \epsilon_t.$$

Therefore,

$$E(y_t) = \phi(L)^{-1} \mu + \phi(L)^{-1} \theta(L) E(\epsilon_t) = \phi(1)^{-1} \mu = \frac{\mu}{1 - \phi_1 - \phi_2 - \cdots - \phi_p}.$$

6.6 ARIMA Model

Autoregressive Integrated Moving Average (ARIMA , 自己回帰和分移動平均) Model

ARIMA(p, d, q) Process

$$\phi(L)\Delta^d y_t = \theta(L)\epsilon_t,$$

where $\Delta^d y_t = \Delta^{d-1}(1-L)y_t = \Delta^{d-1}y_t - \Delta^{d-1}y_{t-1} = (1-L)^d y_t$ for $d = 1, 2, \dots$, and $\Delta^0 y_t = y_t$.

例 : ARIMA(0,1,0) Model

Consider the model: $\Delta y_t = y_t - y_{t-1} + \epsilon_t$, $\epsilon_t \sim N(0, \sigma^2)$, $y_0 = 0$,

which is rewritten as: $y_t = \epsilon_t + \epsilon_{t-1} + \dots + \epsilon_1$.

$$E(y_t) = 0, \quad \gamma(0) = V(y_t) = \sigma^2 t, \quad \gamma(\tau) = \text{Cov}(y_t, y_{t-\tau}) = E(y_t y_{t-\tau}) = \sigma^2(t - \tau),$$

which implies that $\gamma(\tau)$ is time-dependent. $\implies y_t$ is not stationary.

$$\rho(\tau) = \frac{\text{Cov}(y_t, y_{t-\tau})}{\sqrt{V(y_t)} \sqrt{V(y_{t-\tau})}} = \frac{t - \tau}{\sqrt{t} \sqrt{t - \tau}} = \sqrt{\frac{t - \tau}{t}}.$$

That is, $\rho(\tau)$ gradually decreases with slow speed.

6.7 SARIMA Model

Seasonal ARIMA (SARIMA) Process:

1. SARIMA(p, d, q)

$$\phi(L)\Delta^d\Delta_s y_t = \theta(L)\epsilon_t,$$

where

$$\Delta_s y_t = (1 - L^s)y_t = y_t - y_{t-s}.$$

$s = 4$ when y_t denotes quarterly date and $s = 12$ when y_t represents monthly data.

6.8 Optimal Prediction

1. AR(p) Process: $y_t = \phi_1 y_{t-1} + \cdots + \phi_p y_{t-p} + \epsilon_t$

(a) Define:

$$E(y_{t+k}|Y_t) = y_{t+k|t},$$

where Y_t denotes all the information available at time t .

Taking the conditional expectation of $y_{t+k} = \phi_1 y_{t+k-1} + \cdots + \phi_p y_{t+k-p} + \epsilon_{t+k}$ on both sides,

$$y_{t+k|t} = \phi_1 y_{t+k-1|t} + \cdots + \phi_p y_{t+k-p|t},$$

where $y_{s|t} = y_s$ for $s \leq t$.

(b) Optimal prediction is given by solving the above differential equation.

2. MA(q) Process: $y_t = \epsilon_t + \theta_1 \epsilon_{t-1} + \cdots + \theta_q \epsilon_{t-q}$

(a) Let $\hat{\epsilon}_T, \hat{\epsilon}_{T-1}, \cdots, \hat{\epsilon}_1$ be the estimated errors.

(b) $y_{t+k} = \epsilon_{t+k} + \theta_1 \epsilon_{t+k-1} + \cdots + \theta_q \epsilon_{t+k-q}$

(c) Therefore,

$$y_{t+k|t} = \epsilon_{t+k|t} + \theta_1 \epsilon_{t+k-1|t} + \cdots + \theta_q \epsilon_{t+k-q|t},$$

where $\epsilon_{s|t} = 0$ for $s > t$ and $\epsilon_{s|t} = \hat{\epsilon}_s$ for $s \leq t$.

3. ARMA(p, q) Process: $y_t = \phi_1 y_{t-1} + \cdots + \phi_p y_{t-p} + \epsilon_t + \theta_1 \epsilon_{t-1} + \cdots + \theta_q \epsilon_{t-q}$

(a) $y_{t+k} = \phi_1 y_{t+k-1} + \cdots + \phi_p y_{t+k-p} + \epsilon_{t+k} + \theta_1 \epsilon_{t+k-1} + \cdots + \theta_q \epsilon_{t+k-q}$

(b) Optimal prediction is:

$$y_{t+k|t} = \phi_1 y_{t+k-1|t} + \cdots + \phi_p y_{t+k-p|t} + \epsilon_{t+k|t} + \theta_1 \epsilon_{t+k-1|t} + \cdots + \theta_q \epsilon_{t+k-q|t},$$

where $y_{s|t} = y_s$ and $\epsilon_{s|t} = \hat{\epsilon}_s$ for $s \leq t$, and $\epsilon_{s|t} = 0$ for $s > t$.

6.9 Identification (識別 , または , 同定)

We have the following two approaches for model specification.

1. Based on AIC or SBIC given d, s , we obtain p, q .

(a) AIC (Akaike's Information Criterion , 赤池の情報量基準)

$$\text{AIC} = -2 \log(\text{likelihood}) + 2k,$$

where $k = p + q$, which is the number of parameters estimated.

(b) SBIC (Shwarz's Bayesian Information Criterion)

$$\text{SBIC} = -2 \log(\text{likelihood}) + k \log T,$$

where T denotes the number of observations.

2. From the sample autocorrelation coefficient function $\hat{\rho}(k)$ and the sample partial autocorrelation coefficient function $\hat{\phi}_{k,k}$ for $k = 1, 2, \dots$, we obtain p, d, q, s .

	AR(p) Process	MA(q) Process
Autocorrelation Function	Gradually decreasing	$\rho(k) = 0,$ $k = q + 1, q + 2, \dots$
Partial Autocorrelation Function	$\phi(k, k) = 0,$ $k = p + 1, p + 2, \dots$	Gradually decreasing

- (a) Compute $\Delta_s y_t$ to remove seasonality.

Compute the autocovariance functions of $\Delta_s y_t$.

If the autocovariance functions have period s , we take $(1 - L^s)$, again.

- (b) Determine the order of difference.

Compute the partial autocovariance functions every time.

If the autocovariance functions decrease as τ is large, go to the next step.

- (c) Determine the order of AR terms (i.e., p).

Compute the partial autocovariance functions every time.

The partial autocovariance functions are close to zero after some τ , go to the next step.

- (d) Determine the order of MA terms (i.e., q).

Compute the autocovariance functions every time.

If the autocovariance functions are randomly around zero, end of the procedure.

6.10 Example of SARIMA using Consumption Data

Construct SARIMA model using monthly and seasonally unadjusted consumption expenditure data and STATA12.

Estimation Period: Jan., 1970 — Dec., 2012 ($T = 516$)

```
. gen time=_n                                Generate time.
. tsset time                                  Defined as time series data.
    time variable:  time, 1 to 516
      delta:       1 unit
. corrgram expend                             Variable name: expend
                                              corrgram: Compute autocorrelation
                                              and partial autocorrelation.
```

LAG	AC	PAC	Q	Prob>Q	-1	0	1	-1	0	1
					[Autocorrelation]			[Partial Autocor]		
1	0.8488	0.8499	373.88	0.0000						
2	0.8231	0.3858	726.18	0.0000						
3	0.8716	0.5266	1122	0.0000						
4	0.8706	0.4025	1517.6	0.0000						
5	0.8498	0.3447	1895.3	0.0000						
6	0.8085	0.0074	2237.9	0.0000						
7	0.8378	0.1528	2606.5	0.0000						
8	0.8460	0.1467	2983	0.0000						
9	0.8342	0.3006	3349.9	0.0000						

```

10      0.7735  -0.1518      3666  0.0000
11      0.7852  -0.1185     3992.3  0.0000
12      0.9234   0.9442     4444.5  0.0000
13      0.7754  -0.5486     4764.1  0.0000
14      0.7482  -0.3248     5062.1  0.0000
15      0.7963  -0.2392     5400.5  0.0000

```

```

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```

- Autocorrelation does not approach zero for large lag.
- Time series has unit root.

```

. gen dexp=expend-l.expd
(1 missing value generated)

```

Generate dexp=expend-expend(-1), excluding unit root.

```

. corrgram dexp

```

LAG	AC	PAC	Q	Prob>Q	⁻¹ [Autocorrelation]	⁰ [Partial Autocor]	¹
1	-0.4316	-0.4329	96.485	0.0000	---	---	
2	-0.2546	-0.5441	130.13	0.0000	--	----	
3	0.1721	-0.4091	145.53	0.0000	-	----	
4	0.0667	-0.3459	147.85	0.0000		--	
5	0.0715	-0.0036	150.52	0.0000			
6	-0.2428	-0.1489	181.36	0.0000	-	-	
7	0.0711	-0.1400	184.01	0.0000		-	
8	0.0668	-0.2900	186.36	0.0000		--	
9	0.1704	0.1681	201.64	0.0000	-		-
10	-0.2485	0.1306	234.21	0.0000	-		-
11	-0.4293	-0.9305	331.56	0.0000	---	-----	
12	0.9773	0.6768	837.12	0.0000	-----		-----
13	-0.4152	0.3778	928.56	0.0000	---		---
14	-0.2583	0.2688	964.03	0.0000	--		--

15 0.1712 0.0406 979.63 0.0000 | - |

Big autocorrelation at lag 12.

. gen sdex=dexp-l12.dexp
(13 missing values generated)

Generate sdex=dexp-dexp(-12),
excluding seasonality.

. corrgram sdex

LAG	AC	PAC	Q	Prob>Q	⁻¹ [Autocorrelation]	⁰ [Partial Autocor]	¹
1	-0.4752	-0.4753	114.28	0.0000	---	---	
2	-0.0244	-0.3235	114.58	0.0000		--	
3	0.1163	-0.0759	121.46	0.0000			
4	-0.1246	-0.1365	129.37	0.0000		-	
5	0.0341	-0.1016	129.96	0.0000			
6	-0.0151	-0.1136	130.08	0.0000			
7	-0.0395	-0.1413	130.88	0.0000		-	
8	0.1123	0.0092	137.35	0.0000			
9	-0.0664	-0.0100	139.62	0.0000			
10	0.0168	0.0069	139.76	0.0000			
11	0.1642	0.2422	153.68	0.0000	-		-
12	-0.3888	-0.2469	231.9	0.0000	---	-	
13	0.2242	-0.1205	257.96	0.0000	-		
14	-0.0147	-0.0941	258.07	0.0000			
15	-0.0708	-0.0591	260.68	0.0000			

Big autocorrelation at lag 1 (ignore lag 12).

MA(1)

Big partial autocorrelation at lags 1 and 2.

AR(2)


```
. arima sdex, ar(1,2) ma(1)
```

```
Model specification is:  
sdex ~ ARMA(2,1), i.e., expend ~ SARIMA(2,1,1).
```

```
(setting optimization to BHHH)
```

```
Iteration 0: log likelihood = -5107.4608  
Iteration 1: log likelihood = -5102.391  
Iteration 2: log likelihood = -5099.9071  
Iteration 3: log likelihood = -5099.4216  
Iteration 4: log likelihood = -5099.2463  
(switching optimization to BFGS)  
Iteration 5: log likelihood = -5099.2361  
Iteration 6: log likelihood = -5099.2346  
Iteration 7: log likelihood = -5099.2346  
Iteration 8: log likelihood = -5099.2346
```

```
ARIMA regression
```

```
Sample: 14 - 516
```

```
Log likelihood = -5099.235
```

```
Number of obs      =      503  
Wald chi2(3)       =      973.93  
Prob > chi2        =      0.0000
```

sdex	Coef.	OPG Std. Err.	z	P> z	[95% Conf. Interval]
_cons	-15.64573	59.17574	-0.26	0.791	-131.628 100.3366

```
ARMA
```

ar							
L1.	.1271774	.0581883	2.19	0.029	.0131304	.2412244	
L2.	.1009983	.053626	1.88	0.060	-.0041068	.2061034	
ma							
L1.	-.8343264	.0419364	-19.90	0.000	-.9165202	-.7521326	
/sigma		6111.128	139.0105	43.96	0.000	5838.673	6383.584

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

```
. estat ic
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	503	.	-5099.235	5	10208.47	10229.57

Note: N=Obs used in calculating BIC; see [R] BIC note

```
. predict resid, r
(13 missing values generated)
```

. corrgram resid

Make sure $\text{sdex} \sim \text{ARMA}(2,1)$, using the residuals.

LAG	AC	PAC	Q	Prob>Q	$^{-1}$ [Autocorrelation]	0 [Partial Autocor]	1
1	-0.0132	-0.0132	.08814	0.7666			
2	-0.0095	-0.0097	.1341	0.9351			
3	0.1248	0.1246	8.0433	0.0451			
4	-0.0644	-0.0624	10.154	0.0379			
5	-0.0001	0.0011	10.154	0.0710			
6	-0.0138	-0.0309	10.252	0.1144			
7	-0.0032	0.0126	10.257	0.1745			
8	0.0958	0.0938	14.97	0.0597			
9	-0.0317	-0.0255	15.487	0.0784			
10	0.0126	0.0112	15.569	0.1127			
11	-0.0053	-0.0305	15.583	0.1573			
12	-0.3773	-0.3837	89.235	0.0000	---	---	
13	0.0408	0.0258	90.098	0.0000			
14	-0.0233	-0.0307	90.381	0.0000			
15	-0.0911	-0.0059	94.703	0.0000			

Big autocorrelation and
big partial correlation at lag 12.
We need to re-specify the model.