

## 8.2 Unit Root (More Formally)

Consider  $y_t = y_{t-1} + \epsilon_t$  and  $y_0 = 0$ .

$$y_t = \epsilon_1 + \epsilon_2 + \cdots + \epsilon_t \sim N(0, t\sigma^2)$$

$$\frac{1}{\sqrt{T}}y_t = \frac{1}{\sqrt{T}}(\epsilon_1 + \epsilon_2 + \cdots + \epsilon_t) \sim N(0, \frac{t}{T}\sigma^2) \longrightarrow N(0, r\sigma^2)$$

where  $0 \leq r \leq 1$  and  $r = \frac{t}{T}$ .

Note that time interval  $(1, T)$  is transformed into  $(0, 1)$ , divided by  $T$ .

$$\frac{1}{\sqrt{T}\sigma}y_t = \frac{1}{\sqrt{T}\sigma}(\epsilon_1 + \epsilon_2 + \cdots + \epsilon_t) \sim N(0, \frac{t}{T}) \longrightarrow N(0, r) \equiv W(r)$$

As  $T$  ( $t$  at the same time) goes to infinity keeping  $r = \frac{t}{T}$ ,  $W(r)$  results in a continuous function of  $r$  where  $r$  takes any number between zero and one.

$W(r)$  is a normal random variable with mean zero and variance  $r$  and it is called the **Brownian motion**.

Moreover, we consider the following:

$$\frac{1}{\sqrt{T}\sigma}y_{t'} = \frac{1}{\sqrt{T}\sigma}(\epsilon_1 + \epsilon_2 + \cdots + \epsilon_{t'}) \longrightarrow N(0, r') \equiv W(r').$$

For  $t' > t$ , we have the following:

$$\begin{aligned}\frac{1}{\sqrt{T}\sigma}y_{t'} &= \frac{1}{\sqrt{T}\sigma}(\epsilon_1 + \cdots + \epsilon_t + \epsilon_{t+1} + \cdots + \epsilon_{t'}) \longrightarrow N(0, r') \equiv W(r') \\ &= \frac{1}{\sqrt{T}\sigma}y_t + \frac{1}{\sqrt{T}\sigma}(\epsilon_{t+1} + \cdots + \epsilon_{t'}).\end{aligned}$$

Therefore, we have

$$\frac{1}{\sqrt{T}\sigma}y_{t'} - \frac{1}{\sqrt{T}\sigma}y_t = \frac{1}{\sqrt{T}\sigma}(\epsilon_{t+1} + \cdots + \epsilon_{t'}) \longrightarrow N(0, r' - r) \equiv W(r') - W(r).$$

That is,  $W(r)$  is independent of  $W(r') - W(r)$  for  $r' > r$ .

Moreover, note as follows:

$$\frac{1}{T \sqrt{T} \sigma} \sum_{t=1}^T y_t = \frac{1}{T} \sum_{t=1}^T \left( \frac{y_t}{\sqrt{T} \sigma} \right) \longrightarrow \int_0^1 W(r) dr$$

where  $\frac{1}{T}$  and  $\sum_{t=1}^T$  are replaced by  $dr$  and  $\int_0^1$  as  $T$  goes to infinity.

We divide the time interval  $(0, 1)$  into  $T$  time intervals  $\left(\frac{t}{T}, \frac{t+1}{T}\right)$ .

That is, time interval  $(1, T)$  is transformed into  $(0, 1)$ .

(\*) We know that  $\frac{y_t}{\sqrt{T} \sigma} \longrightarrow W(r)$  as  $\frac{t}{T} \longrightarrow r$ .

**Summary: Properties of  $W(r)$  for  $0 < r < 1$ :**

1.  $W(r) \equiv N(0, r) \implies W(r)$  is a random variable.
2.  $W(1) \equiv N(0, 1)$
3.  $W(1)^2 \equiv \chi^2(1) \implies$  Remember that  $Z^2 \sim \chi^2(1)$  when  $Z \sim N(0, 1)$ .
4.  $W(r)$  is independent of  $W(r') - W(r)$  for  $r < r'$ .
5.  $W(r_4) - W(r_3)$  is independent of  $W(r_2) - W(r_1)$  for  $0 \leq r_1 < r_2 < r_3 < r_4 \leq 1$ .  
 $\implies$  The interval between  $r_4$  and  $r_3$  is not overlapped with the interval between  $r_2$  and  $r_1$ .

- **True Model**  $y_t = y_{t-1} + \epsilon_t$  vs **Estimated Model**  $y_t = \phi y_{t-1} + \epsilon_t$ : Under  $\phi = 1$ , we estimate  $\phi$  in the regression model:

$$y_t = \phi y_{t-1} + \epsilon_t$$

OLS of  $\phi$  is:

$$\hat{\phi} = \frac{\sum_{t=1}^T y_{t-1} y_t}{\sum_{t=1}^T y_{t-1}^2} = \phi + \frac{\sum_{t=1}^T y_{t-1} \epsilon_t}{\sum_{t=1}^T y_{t-1}^2}$$

As mentioned above, the numerator is related to:

$$\frac{1}{\sigma^2 T} \sum_{t=1}^T y_{t-1} \epsilon_t = \frac{1}{2} \left( \frac{y_T}{\sigma \sqrt{T}} \right)^2 - \frac{1}{2\sigma^2} \frac{1}{T} \sum_{t=1}^T \epsilon_t^2 \longrightarrow \frac{1}{2} W(1)^2 - \frac{1}{2}$$

which is rewritten by using the Brownian motion  $W(1)$ .

The denominator is:

$$\frac{1}{\sigma^2 T^2} \sum_{t=1}^T y_{t-1}^2 \approx \frac{1}{\sigma^2 T^2} \sum_{t=1}^T y_t^2 = \frac{1}{T} \sum_{t=1}^T \left( \frac{y_t}{\sigma \sqrt{T}} \right)^2 \longrightarrow \int_0^1 W(r)^2 dr$$

where  $\frac{1}{T} \rightarrow dr$  and  $\frac{y_t}{\sigma\sqrt{T}} \rightarrow W(r)$  for  $\frac{t}{T} \rightarrow r$ .

Thus, under  $\phi = 1$ ,  $T(\hat{\phi} - \phi)$  is asymptotically distributed as follows:

$$T(\hat{\phi} - \phi) = T(\hat{\phi} - 1) = \frac{\frac{1}{\sigma^2 T} \sum_{t=1}^T y_{t-1} \epsilon_t}{\frac{1}{\sigma^2 T^2} \sum_{t=1}^T y_{t-1}^2} \rightarrow \frac{\frac{1}{2}(W(1)^2 - 1)}{\int_0^1 W(r)^2 dr}$$

•  $t$  value:

In the regression model:  $y_t - y_{t-1} \equiv \Delta y_t = \rho y_{t-1} + \epsilon_t$ , OLSE of  $\rho = \phi - 1$  is given by

$$\hat{\rho} = \hat{\phi} - 1.$$

$t$  value is  $\frac{\hat{\rho}}{s_\rho} = \frac{\hat{\phi} - 1}{s_\phi}$ , where  $s_\rho$  and  $s_\phi$  denote the standard errors of  $\hat{\rho}$  and  $\hat{\phi}$ .

Note that  $s_\rho = s_\phi$  because of  $V(\hat{\rho}) = V(\hat{\phi} - 1) = V(\hat{\phi})$ .

The standard error of  $\hat{\phi}$ , denoted by  $s_\phi$ , is given by:  $s_\phi^2 = \frac{s^2}{\sum_{t=1}^T y_{t-1}^2}$ , where  $s^2 = \frac{1}{T} \sum_{t=1}^T (y_t - \hat{\phi}y_{t-1})^2$ , called the standard error of regression.

$$\begin{aligned}
 s^2 &= \frac{1}{T} \sum_{t=1}^T (y_t - \hat{\phi}y_{t-1})^2 \\
 &= \frac{1}{T} \sum_{t=1}^T (\epsilon_t - (\hat{\phi} - 1)y_{t-1})^2 \\
 &= \frac{1}{T} \sum_{t=1}^T \epsilon_t^2 - 2\frac{1}{T}\frac{1}{T}(\hat{\phi} - 1) \sum_{t=1}^T y_{t-1}\epsilon_t + \frac{1}{T}(\hat{\phi} - 1)^2 \sum_{t=1}^T y_{t-1}^2 \\
 &= \frac{1}{T} \sum_{t=1}^T \epsilon_t^2 - 2\frac{\sigma^2}{T}[T(\hat{\phi} - 1)]\left[\frac{1}{T\sigma^2} \sum_{t=1}^T y_{t-1}\epsilon_t\right] + \frac{\sigma^2}{T}[T(\hat{\phi} - 1)]^2\left[\frac{1}{T^2\sigma^2} \sum_{t=1}^T y_{t-1}^2\right] \\
 &\rightarrow \sigma^2.
 \end{aligned}$$

The random variables in  $[\cdot]$  converge in distribution..

Note that in the right hand side of the fourth line the second and third terms go to zero, because we know the followings:

$$\begin{aligned} \frac{1}{T} \sum_{t=1}^T \epsilon_t^2 &\rightarrow \sigma^2 \\ T(\hat{\phi} - 1) &\rightarrow \frac{\frac{1}{2}(W(1)^2 - 1)}{\int_0^1 W(r)^2 dr} \\ \frac{1}{T\sigma^2} \sum_{t=1}^T y_{t-1}\epsilon_t &\rightarrow \frac{1}{2}W(1)^2 - \frac{1}{2} \\ \frac{1}{T^2\sigma^2} \sum_{t=1}^T y_t^2 &\rightarrow \int_0^1 W(r)^2 dr \end{aligned}$$

Therefore, from  $s_\phi^2 = \frac{1}{T^2\sigma^2} \frac{s^2}{\frac{1}{T^2\sigma^2} \sum_{t=1}^T y_{t-1}^2}$ , we obtain  $T^2 s_\phi^2 \rightarrow \left(\int_0^1 W(r)^2 dr\right)^{-1}$ .



$t$  value is given by:

$$\frac{\hat{\phi} - 1}{s_{\phi}} = \frac{T(\hat{\phi} - 1)}{T s_{\phi}} \rightarrow \frac{\frac{1}{2}(W(1)^2 - 1) / \int_0^1 W(r)^2 dr}{\left(\int_0^1 W(r)^2 dr\right)^{-1/2}} = \frac{\frac{1}{2}(W(1)^2 - 1)}{\left(\int_0^1 W(r)^2 dr\right)^{1/2}},$$

which is not a normal distribution.

- **True Model**  $y_t = y_{t-1} + \epsilon_t$  vs **Estimated Model**  $y_t = \alpha + \phi y_{t-1} + \epsilon_t$ : Under  $\alpha = 0$  and  $\phi = 1$ , we estimate  $\alpha$  and  $\phi$  in the regression model:

$$y_t = \alpha + \phi y_{t-1} + \epsilon_t$$

OLSEs of  $\alpha$  and  $\phi$  are:

$$\begin{aligned} \begin{pmatrix} \hat{\alpha} \\ \hat{\phi} \end{pmatrix} &= \begin{pmatrix} T & \sum y_{t-1} \\ \sum y_{t-1} & \sum y_{t-1}^2 \end{pmatrix}^{-1} \begin{pmatrix} \sum y_t \\ \sum y_{t-1} y_t \end{pmatrix} = \begin{pmatrix} \alpha \\ \phi \end{pmatrix} + \begin{pmatrix} T & \sum y_{t-1} \\ \sum y_{t-1} & \sum y_{t-1}^2 \end{pmatrix}^{-1} \begin{pmatrix} \sum \epsilon_t \\ \sum y_{t-1} \epsilon_t \end{pmatrix} \\ &= \begin{pmatrix} \alpha \\ \phi \end{pmatrix} + \frac{1}{T \sum y_{t-1}^2 - (\sum y_{t-1})^2} \begin{pmatrix} \sum y_{t-1}^2 & -\sum y_{t-1} \\ -\sum y_{t-1} & T \end{pmatrix} \begin{pmatrix} \sum \epsilon_t \\ \sum y_{t-1} \epsilon_t \end{pmatrix} \\ &= \begin{pmatrix} \alpha \\ \phi \end{pmatrix} + \frac{1}{T \sum y_{t-1}^2 - (\sum y_{t-1})^2} \begin{pmatrix} (\sum y_{t-1}^2)(\sum \epsilon_t) - (\sum y_{t-1})(\sum y_{t-1} \epsilon_t) \\ -(\sum y_{t-1})(\sum \epsilon_t) + T(\sum y_{t-1} \epsilon_t) \end{pmatrix} \end{aligned}$$

$$= \begin{pmatrix} \alpha \\ \phi \end{pmatrix} + \begin{pmatrix} \frac{(\sum y_t^2)(\frac{1}{T} \sum \epsilon_t) - \bar{y} \sum y_{t-1} \epsilon_t}{\sum y_t^2 - T\bar{y}^2} \\ \frac{-\bar{y} \sum \epsilon_t + \sum y_{t-1} \epsilon_t}{\sum y_t^2 - T\bar{y}^2} \end{pmatrix}$$

Note that  $\frac{1}{T} \sum y_{t=1} \approx \frac{1}{T} \sum y_t = \bar{y}$  and  $\sum y_{t-1}^2 \approx \sum y_t^2$  for large  $T$ .

In the true model,  $\alpha = 0$  and  $\phi = 1$ .

$$\begin{pmatrix} \hat{\alpha} \\ \hat{\phi} - 1 \end{pmatrix} = \begin{pmatrix} \frac{(\sum y_t^2)(\frac{1}{T} \sum \epsilon_t) - \bar{y} \sum y_{t-1} \epsilon_t}{\sum y_t^2 - T\bar{y}^2} \\ \frac{-\bar{y} \sum \epsilon_t + \sum y_{t-1} \epsilon_t}{\sum y_t^2 - T\bar{y}^2} \end{pmatrix}$$

For each element of the vector, we consider each term in the numerator and denominator.

- $\sum_{t=1}^T y_{t-1} \epsilon_t$ :

Taking the square of  $y_t = y_{t-1} + \epsilon_t$  on both sides, we obtain:  $y_t^2 = y_{t-1}^2 + y_{t-1} \epsilon_t + \epsilon_t^2$ .

Then, we can rewrite as:  $y_{t-1} \epsilon_t = \frac{1}{2}(y_t^2 - y_{t-1}^2 - \epsilon_t^2)$  for  $y_0 = 0$ .

Taking a sum from  $t = 1$  to  $T$ , we have:

$$\sum_{t=1}^T y_{t-1} \epsilon_t = \frac{1}{2} \sum_{t=1}^T (y_t^2 - y_{t-1}^2 - \epsilon_t^2) = \frac{1}{2} y_T^2 - \frac{1}{2} \sum_{t=1}^T \epsilon_t^2,$$

which is divided by  $T\sigma^2$  on both sides, then we obtain:

$$\frac{1}{T\sigma^2} \sum_{t=1}^T y_{t-1} \epsilon_t = \frac{1}{2} \left( \frac{y_T}{\sqrt{T}\sigma} \right)^2 - \frac{1}{2\sigma^2} \frac{1}{T} \sum_{t=1}^T \epsilon_t^2 \rightarrow \frac{1}{2} W(1)^2 - \frac{1}{2}.$$

Note that  $\frac{y_T}{\sqrt{T}\sigma} = W(1)$  and  $\frac{1}{T} \sum_{t=1}^T \epsilon_t^2 \rightarrow E(\epsilon_t^2) = \sigma^2$ .

•  $\bar{y}$ :

Note that  $\frac{1}{T} \sum_{t=1}^T y_t = \bar{y}$  and  $\frac{1}{T} \sum_{t=1}^T y_{t-1} = \bar{y}$ . We can rewrite  $\bar{y}$  as follows:

$$\bar{y} = \frac{1}{T} \sum_{t=1}^T y_t = \sqrt{T}\sigma \frac{1}{T} \sum_{t=1}^T \frac{y_t}{\sqrt{T}\sigma}$$

which is rewritten as:

$$\frac{\bar{y}}{\sqrt{T}\sigma} = \frac{1}{T} \sum_{t=1}^T \frac{y_t}{\sqrt{T}\sigma} \rightarrow \int_0^1 W(r)dr$$

Note that  $\frac{y_t}{\sqrt{T}\sigma} \rightarrow W(r)$  as  $\frac{t}{T} \rightarrow r$ .

- $\sum_{t=1}^T \epsilon_t$ :

From  $y_T = \sum_{t=1}^T \epsilon_t$ , we have:  $\frac{1}{\sqrt{T}\sigma} \sum_{t=1}^T \epsilon_t = \frac{y_T}{\sqrt{T}\sigma} = W(1)$ .

- $\sum_{t=1}^T y_{t-1}^2$ :

From  $\sum_{t=1}^T y_{t-1}^2 \approx \sum_{t=1}^T y_t^2$ , we obtain:  $\frac{1}{T} \sum_{t=1}^T \left(\frac{y_t}{\sqrt{T}\sigma}\right)^2 \rightarrow \int_0^1 W(r)^2 dr$ .

Note that  $\frac{y_t}{\sqrt{T}\sigma} \rightarrow W(r)$  as  $\frac{t}{T} \rightarrow r$ .

Thus,  $\hat{\phi} - 1 = \frac{\sum y_{t-1}\epsilon_t - \bar{y} \sum \epsilon_t}{\sum y_t^2 - T\bar{y}^2}$  is rewritten as:

$$\begin{aligned} T(\hat{\phi} - 1) &= \frac{\frac{1}{T\sigma^2}(\sum y_{t-1}\epsilon_t - \bar{y} \sum \epsilon_t)}{\frac{1}{T^2\sigma^2}(\sum y_t^2 - T\bar{y}^2)} = \frac{\frac{1}{T\sigma^2} \sum y_{t-1}\epsilon_t - \left(\frac{\bar{y}}{\sqrt{T}\sigma}\right)\left(\frac{1}{\sqrt{T}\sigma} \sum \epsilon_t\right)}{\frac{1}{T} \sum \left(\frac{y_t}{\sqrt{T}\sigma}\right)^2 - \left(\frac{\bar{y}}{\sqrt{T}\sigma}\right)^2} \\ &\rightarrow \frac{\frac{1}{2}(W(1)^2 - 1) - W(1) \int_0^1 W(r)dr}{\int_0^1 W(r)^2 dr - \left(\int_0^1 W(r)dr\right)^2} \end{aligned}$$

Remember that OLSE of  $\alpha$  is given by:

$$\hat{\alpha} = \alpha + \frac{(\sum y_t^2)\left(\frac{1}{T} \sum \epsilon_t\right) - \bar{y} \sum y_{t-1}\epsilon_t}{\sum y_t^2 - T\bar{y}^2}$$

Under  $\alpha = 0$ ,  $\hat{\alpha}$  is rewritten as follows:

$$\begin{aligned} \sqrt{T}\hat{\alpha} &= \frac{\sigma\left(\frac{1}{T} \sum \left(\frac{y_t}{\sqrt{T}\sigma}\right)^2\right)\left(\frac{1}{\sqrt{T}\sigma} \sum \epsilon_t\right) - \sigma\left(\frac{\bar{y}}{\sqrt{T}\sigma}\right)\left(\frac{1}{T\sigma^2} \sum y_{t-1}\epsilon_t\right)}{\frac{1}{T} \sum \left(\frac{y_t}{\sqrt{T}\sigma}\right)^2 - \left(\frac{\bar{y}}{\sqrt{T}\sigma}\right)^2} \\ &\rightarrow \frac{\sigma W(1) \int_0^1 W(r)^2 dr - \sigma \frac{1}{2}(W(1)^2 - 1) \int_0^1 W(r) dr}{\int_0^1 W(r)^2 dr - \left(\int_0^1 W(r) dr\right)^2} \end{aligned}$$

Thus, convergence speed of  $\hat{\phi}$  is different from that of  $\hat{\alpha}$ .

Neither  $\sqrt{T}\hat{\alpha}$  nor  $T(\hat{\phi} - 1)$  are normal.