

## Econometrics II TA Session #3

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## Example 1: Binary Choice Model

- Consider the regression model:

$$y_i^* = X_i\beta + u_i, \quad u_i \sim (0, \sigma^2), \quad i = 1, 2, \dots, n,$$

where  $y_i^*$  is unobserved, but  $y_i$  is observed as 0 or 1, i.e.,

$$y_i = \begin{cases} 1, & \text{if } y_i^* > 0, \\ 0, & \text{if } y_i^* \leq 0. \end{cases}$$

- E.g.)  $y_i^*$ : productivity of market work (continuous variable)  
 $y_i$ : whether an individual is employed or not (discrete variable)
- Note that we do not specify the distribution of  $u_i$ .

- Consider the probability that  $y_i$  takes 1, i.e.,

$$\begin{aligned}\mathbb{P}(y_i = 1) &= \mathbb{P}(y_i^* > 0) \\ &= \mathbb{P}(u_i > -X_i\beta) \\ &= \mathbb{P}\left(\frac{u_i}{\sigma} > -X_i\frac{\beta}{\sigma}\right) \\ &= \mathbb{P}(u_i^* > -X_i\beta^*) \\ &= 1 - \mathbb{P}(u_i^* \leq -X_i\beta^*) \\ &= 1 - F(-X_i\beta^*) \\ &= F(X_i\beta^*),\end{aligned}$$

where the last equality holds if the distribution of  $u_i^*$  is symmetric.

- The distribution function of  $u_i^*$  is  $F(x) = \int_{-\infty}^x f(z)dz$ .
  - If  $u_i^*$  follows standard normal distribution, we call **Probit model**.

$$F(x) = \int_{-\infty}^x \frac{1}{2\pi} \exp\left(-\frac{1}{2}z^2\right) dz$$

- If  $u_i^*$  follows logistic distribution, we call **Logit model**.

$$F(x) = \frac{1}{1 + \exp(-x)}$$

- Since  $y_i$  is a binary variable,  $y_i$  follows Bernoulli distribution.
- Then, the density function of  $y_i$  is given by

$$\begin{aligned}f(y_i) &= [\mathbb{P}(y_i = 1)]^{y_i} [\mathbb{P}(y_i = 0)]^{1-y_i} \\&= [F(X_i\beta^*)]^{y_i} [1 - F(X_i\beta^*)]^{1-y_i}.\end{aligned}$$

- Using this density function, we define the likelihood function:

$$\begin{aligned}L(\beta^*) &= f(y_1, \dots, y_n) \\&= \prod_{i=1}^n f(y_i) \\&= \prod_{i=1}^n [F(X_i\beta^*)]^{y_i} [1 - F(X_i\beta^*)]^{1-y_i}.\end{aligned}$$

- The log-likelihood function is:

$$\log L(\beta^*) = \sum_{i=1}^n \left[ y_i \log F(X_i\beta^*) + (1 - y_i) \log[1 - F(X_i\beta^*)] \right].$$

- The F.O.C. is:

$$\begin{aligned}\frac{\partial \log L(\beta^*)}{\partial \beta^*} &= \sum_{i=1}^n \left( \frac{y_i X_i' f(X_i \beta^*)}{F(X_i \beta^*)} - \frac{(1 - y_i) X_i' f(X_i \beta^*)}{1 - F(X_i \beta^*)} \right) \\ &= \sum_{i=1}^n \frac{X_i' f_i(y_i - F_i)}{F_i(1 - F_i)} = 0,\end{aligned}$$

where  $f_i = f(X_i \beta^*)$  and  $F_i = F(X_i \beta^*)$ .

- The S.O.C. is:

$$\begin{aligned}\frac{\partial^2 \log L(\beta^*)}{\partial \beta^* \partial \beta^{*\prime}} &= \sum_{i=1}^n \frac{X_i' \frac{\partial f_i}{\partial \beta^*} (y_i - F_i)}{F_i(1 - F_i)} + \sum_{i=1}^n \frac{X_i' f_i \frac{\partial (f_i - F_i)}{\partial \beta^*}}{F_i(1 - F_i)} + \sum_{i=1}^n X_i' f_i (y_i - F_i) \frac{\partial [F_i(1 - F_i)]^{-1}}{\partial \beta^*} \\ &= \sum_{i=1}^n \frac{X_i' X_i f_i' (y_i - F_i)}{F_i(1 - F_i)} - \sum_{i=1}^n \frac{X_i' X_i f_i^2}{F_i(1 - F_i)} + \sum_{i=1}^n X_i' f_i (y_i - F_i) \frac{X_i f_i (1 - 2F_i)}{[F_i(1 - F_i)]^2}\end{aligned}$$

is a negative definite matrix.

- For maximization, the method of scoring is:

$$\begin{aligned}\beta^{*(j+1)} &= \beta^{*(j)} + \left[ -\mathbb{E}\left(\frac{\partial^2 \log L(\beta^{*(j)})}{\partial \beta^* \partial \beta^{*'}}\right) \right]^{-1} \frac{\partial \log L(\beta^{*(j)})}{\partial \beta^*} \\ &= \beta^{*(j)} + \left[ \sum_{i=1}^n \frac{X_i' X_i (f_i^{(j)})^2}{F_i^{(j)} (1 - F_i^{(j)})} \right]^{-1} \sum_{i=1}^n \frac{X_i' f_i^{(j)} (y_i - F_i^{(j)})}{F_i^{(j)} (1 - F_i^{(j)})},\end{aligned}$$

where  $F_i^{(j)} = F(X_i \beta^{*(j)})$  and  $f^{(j)} = f(X_i \beta^{*(j)})$ .

- Note that we use the following relationship:

$$\mathbb{E}[y_i] = \mathbb{P}(y_i = 1) = F_i(X_i \beta^*) = F_i.$$

- The Fisher information matrix is given by:

$$I(\beta^*) = -\mathbb{E}\left[\frac{\partial^2 \log L(\beta^*)}{\partial \beta^* \partial \beta^{*\prime}}\right] = \sum_{i=1}^n \frac{X_i' X_i f_i^2}{F_i(1-F_i)}.$$

- By the asymptotic normality,

$$\sqrt{n}(\hat{\beta}^* - \beta^*) \xrightarrow{d} \mathcal{N}\left(0, \lim_{n \rightarrow \infty} \left(-\frac{1}{n} \mathbb{E}\left[\frac{\partial^2 \log L(\beta^*)}{\partial \beta^* \partial \beta^{*\prime}}\right]^{-1}\right)\right),$$

where  $\hat{\beta}^* := \lim_{j \rightarrow \infty} \beta^{*(j)}$  denotes the MLE of  $\beta^*$ .

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## Example 2

- Consider the two utility functions:

$$U_{1i} = X_i\beta_1 + \epsilon_{1i}, \quad (1)$$

$$U_{2i} = X_i\beta_2 + \epsilon_{2i}. \quad (2)$$

- We purchase a good when  $U_{1i} > U_{2i}$  and do not purchase otherwise.
- $y_i$  takes 1 if we purchase the good and takes 0 otherwise.
- We can observe  $y_i$ , but can NOT observe  $U_{ji}$ ,  $j \in \{1, 2\}$ .

- Taking a difference between (1) and (2), we have

$$\begin{aligned} U_{1i} - U_{2i} &= X_i(\beta_1 - \beta_2) + (\epsilon_{1i} - \epsilon_{2i}) \\ \iff U_i^* &= X_i\beta^* + \epsilon^*, \end{aligned}$$

where  $U_i^* := U_{1i} - U_{2i}$ ,  $\beta^* := \beta_1 - \beta_2$  and  $\epsilon^* := \epsilon_{1i} - \epsilon_{2i}$ .

- Then, we have the following relationship:

$$y_i = \begin{cases} 1 & \text{if } U_i^* > 0, \\ 0 & \text{if } U_i^* \leq 0, \end{cases}$$

which is the same situation as Example 1 and the assumption that  $\epsilon_i^*$  follows a **symmetric** distribution is necessary.

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## Example 3

- Consider the questionnaire:

$$y_i = \begin{cases} 1, & \text{if the } i\text{th person answers YES,} \\ 0, & \text{if the } i\text{th person answers NO.} \end{cases}$$

- Consider the following linear regression model:

$$y_i = X_i\beta + u_i.$$

- For instance, the question is "Do you have a car?" and  $X_i$  includes income, living place, and family size, etc.

- When  $\mathbb{E}[u_i] = 0$ ,

$$\mathbb{E}[y_i] = X_i\beta.$$

- Because of the linear function,  $X_i\beta$  takes the value from  $-\infty$  to  $\infty$ .
- For instance, if an individual  $i$  has a high income, lives in the countryside and has many children,  $X_i\beta$  takes a value that is greater than 1.
- However, since  $\mathbb{E}[y_i]$  means the probability that an individual  $i$  has a car,  $\mathbb{E}[y_i]$  must be in  $[0, 1]$ .

- Alternatively, consider the following model:

$$y_i = \mathbb{P}(y_i = 1) + u_i,$$

where  $u_i$  is a discrete type of random variable, i.e.,

$$u_i = \begin{cases} 1 - \mathbb{P}(y_i = 1) & \text{with prob. } \mathbb{P}(y_i = 1), \\ -\mathbb{P}(y_i = 1) & \text{with prob. } 1 - \mathbb{P}(y_i = 1) = \mathbb{P}(y_i = 0). \end{cases}$$

- Consider that  $\mathbb{P}(y_i = 1)$  is connected with the distribution function  $F(X_i\beta)$  as follows:

$$\mathbb{P}(y_i = 1) = F(X_i\beta).$$

- Assuming that  $F(\cdot)$  is normal distribution or logistic distribution results in probit model or logit model, respectively.

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## Example 4: Ordered probit or logit model

- Consider the regression model:

$$y_i^* = X_i\beta + u_i, \quad u_i \sim (0, 1), \quad i = 1, \dots, n,$$

where  $y_i^*$  is unobserved, but  $y_i$  is observed as  $1, 2, \dots, m$ , i.e.,

$$y_i = \begin{cases} 1, & \text{if } -\infty < y_i^* \leq a_1, \\ 2, & \text{if } a_1 < y_i^* \leq a_2, \\ \vdots \\ m, & \text{if } a_{m-1} < y_i^* \leq \infty, \end{cases}$$

where  $a_1, \dots, a_{m-1}$  are assumed to be known.

- For instance,  $y_i^*$  is hours worked per week and  $y_i$  is a discrete variable such that

$$y_i = \begin{cases} 1, & \text{if } y_i^* \leq 15, \\ 2, & \text{if } 15 < y_i^* \leq 20, \\ \vdots \\ 9, & \text{if } 50 < y_i^*. \end{cases}$$

- Consider the probability that  $y_i$  takes  $1, \dots, m$ .

$$\begin{aligned}\mathbb{P}(y_i = 1) &= \mathbb{P}(y_i^* \leq a_1) \\ &= \mathbb{P}(u_i < a_1 - X_i\beta) \\ &= F(a_1 - X_i\beta).\end{aligned}$$

$$\begin{aligned}\mathbb{P}(y_i = 2) &= \mathbb{P}(a_1 < y_i^* \leq a_2) \\ &= \mathbb{P}(a_1 - X_i\beta < u_i \leq a_2 - X_i\beta) \\ &= F(a_2 - X_i\beta) - F(a_1 - X_i\beta).\end{aligned}$$

$$\begin{aligned}\mathbb{P}(y_i = m) &= \mathbb{P}(a_{m-1} < y_i^*) \\ &= \mathbb{P}(a_{m-1} - X_i\beta < u_i) \\ &= 1 - F(a_{m-1} - X_i\beta).\end{aligned}$$

- We have

$$\mathbb{P}(y_i = j) = F(a_j - X_i\beta) - F(a_{j-1} - X_i\beta), \quad \forall j \in \{0, 1, \dots, m\}$$

where  $a_0 = -\infty$  and  $a_m = \infty$ .

- Define the following indicator functions:

$$I_{ij} = \begin{cases} 1, & \text{if } y_i = j, \\ 0, & \text{otherwise.} \end{cases}$$

- The likelihood function is:

$$\begin{aligned} L(\beta) &= \prod_{i=1}^n \left[ F(a_1 - X_i\beta) \right]^{I_{i1}} \left[ F(a_2 - X_i\beta) - F(a_1 - X_i\beta) \right]^{I_{i2}} \cdots \left[ 1 - F(a_{m-1} - X_i\beta) \right]^{I_{im}} \\ &= \prod_{i=1}^n \prod_{j=1}^m \left[ F(a_j - X_i\beta) - F(a_{j-1} - X_i\beta) \right]^{I_{ij}}. \end{aligned}$$

- The log-likelihood function is:

$$\log L(\beta) = \sum_{i=1}^n \sum_{j=1}^m I_{ij} \log \left[ F(a_j - X_i \beta) - F(a_{j-1} - X_i \beta) \right].$$

- The first derivative of  $\log L(\beta)$  w.r.t.  $\beta$  is:

$$\frac{\partial \log L(\beta)}{\partial \beta} = \sum_{i=1}^n \sum_{j=1}^m \frac{-I_{ij} X_i' \left[ f(a_j - X_i \beta) - f(a_{j-1} - X_i \beta) \right]}{F(a_j - X_i \beta) - F(a_{j-1} - X_i \beta)} = 0.$$

- Usually, normal distribution or logistic distribution is chosen for  $F(\cdot)$ .

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## Example 5: Multinomial logit model

- The  $i$ th individual has  $m + 1$  choices, i.e.,  $j = 0, 1, \dots, m$ .

$$\begin{aligned}\mathbb{P}(y_i = j) &= \frac{\exp(X_i\beta_j)}{\sum_{j=0}^m \exp(X_i\beta_j)} \\ &= \frac{\exp(X_i\beta_j)}{\exp(X_i\beta_0) + \exp(X_i\beta_1) + \dots + \exp(X_i\beta_m)} \\ &= \frac{\exp(X_i\beta_j)}{1 + \exp(X_i\beta_1) + \dots + \exp(X_i\beta_m)} =: P_{ij},\end{aligned}$$

where  $\beta_0 = 0$ .

- Different from the ordered probit or logit model, the order does not matter.
- For instance,  $X_i$  is IQ and  $y_i$  indicates occupations:  $y_i$  takes 1 if  $i$  is a cook, takes 2 if  $i$  is a professor and takes 3 if  $i$  is an artist.
- In using the multinomial logit model, a choice is set to be a comparison.
- Therefore, if we set being a cook as a comparison, we interpret the estimation results as follows:
  - one unit of increase in IQ increases the probability of being a professor **relative to being a cook** by A%;
  - one unit of increase in IQ increases the probability of being an artist **relative to being a cook** by B%.

- Note that

$$P_{i0} = \frac{1}{1 + \exp(X_i\beta_1) + \cdots + \exp(X_i\beta_m)}$$

- Then, we have

$$\frac{P_{ij}}{P_{i0}} = \exp(X_i\beta_j) \iff \log \frac{P_{ij}}{P_{i0}} = X_i\beta_j.$$

- The log-likelihood function is:

$$\log L(\beta_1, \dots, \beta_m) = \sum_{i=1}^n \sum_{j=1}^m d_{ij} \log P_{ij},$$

where  $d_{ij} = 1$  when the  $i$ th individual chooses  $j$ th choice, and  $d_{ij} = 0$  otherwise.

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## Example 6: Nested logit model

- Consider the following 2 steps:
  - Choose YES or NO with probability  $P_Y$  and  $P_N = 1 - P_Y$ , respectively. Go to the next step only if YES is chosen.
  - Choose A or B with probability  $P_{A|Y}$  and  $P_{B|Y}$ , respectively.
- For instance, the individual decides whether or not to buy a car in the first step and chooses Audi or BMW.
- Assume the logistic distribution.

- The probability that the  $i$ th individual chooses NO is:

$$P_{N,i} = \frac{1}{1 + \exp(X_i\beta)}.$$

- The probability that the  $i$ th individual chooses YES and A is:

$$P_{A|Y,i}P_{Y,i} = P_{A|Y,i}(1 - P_{N,i}) = \frac{\exp(Z_i\alpha)}{1 + \exp(Z_i\alpha)} \frac{\exp(X_i\beta)}{1 + \exp(X_i\beta)}.$$

- The probability that the  $i$ th individual chooses YES and B is:

$$P_{B|Y,i}P_{Y,i} = (1 - P_{A|Y,i})(1 - P_{N,i}) = \frac{1}{1 + \exp(Z_i\alpha)} \frac{\exp(X_i\beta)}{1 + \exp(X_i\beta)}.$$

- $X_i$  means variables which affect the decision making on whether to buy a car:
  - annual income
  - distance from the nearest station
- $Z_i$  means variables which characterize a car:
  - speed
  - fuel-efficiency
  - car company

- The likelihood function is:

$$\begin{aligned} L(\alpha, \beta) &= \prod_{i=1}^n P_{N,i}^{\mathbf{I}_{1i}} \left\{ \left[ (1 - P_{N,i}) P_{A|Y,i} \right]^{\mathbf{I}_{2i}} \left[ (1 - P_{N,i})(1 - P_{A|Y,i}) \right]^{1-\mathbf{I}_{2i}} \right\}^{1-\mathbf{I}_{1i}} \\ &= \prod_{i=1}^n P_{N,i}^{\mathbf{I}_{1i}} (1 - P_{N,i})^{1-\mathbf{I}_{1i}} \left[ P_{A|Y,i}^{\mathbf{I}_{2i}} (1 - P_{A|Y,i})^{1-\mathbf{I}_{2i}} \right]^{1-\mathbf{I}_{1i}}, \end{aligned}$$

where

- $I_{1i}$  takes 1 if  $i$ th individual decides not to buy a car in the first step and takes 0 otherwise;
- $I_{2i}$  takes 1 if  $i$ th individual chooses A in the second step and takes 0 otherwise.

- Let individual  $i \in \mathcal{N}$  decide not to buy a car,  $i \in \mathcal{A}$  choose A and  $i \in \mathcal{B}$  choose B.
- Then, the likelihood function becomes:

$$\begin{aligned} L(\alpha, \beta) &= \prod_{i=1}^n P_{N,i}^{I_{1i}} (1 - P_{N,i})^{1-I_{1i}} \left[ P_{A|Y,i}^{I_{2i}} (1 - P_{A|Y,i})^{1-I_{2i}} \right]^{1-I_{1i}} \\ &= \prod_{i \in \mathcal{N}} P_{N,i} \times \prod_{i \in \mathcal{A}} (1 - P_{N,i}) P_{A|Y,i} \times \prod_{i \in \mathcal{B}} (1 - P_{N,i}) (1 - P_{A|Y,i}). \end{aligned}$$

- The log-likelihood function is:

$$\log L(\alpha, \beta) = \sum_{i \in \mathcal{N}} P_{N,i} + \sum_{i \in \mathcal{A}} (1 - P_{N,i}) P_{A|Y,i} + \sum_{i \in \mathcal{B}} (1 - P_{N,i}) (1 - P_{A|Y,i}).$$

- Substituting the expressions above into the log-likelihood yields:

$$\begin{aligned}\log L(\alpha, \beta) &= \sum_{i \in \mathcal{N}} \textcolor{red}{P_{N,i}} + \sum_{i \in \mathcal{A}} (1 - P_{N,i}) P_{A|Y,i} + \sum_{i \in \mathcal{B}} (1 - P_{N,i})(1 - P_{A|Y,i}) \\ &= \sum_{i \in \mathcal{N}} \frac{1}{1 + \exp(X_i \beta)} + \sum_{i \in \mathcal{A}} \frac{\exp(Z_i \alpha)}{1 + \exp(Z_i \alpha)} \frac{\exp(X_i \beta)}{1 + \exp(X_i \beta)} \\ &\quad + \sum_{i \in \mathcal{B}} \frac{1}{1 + \exp(Z_i \alpha)} \frac{\exp(X_i \beta)}{1 + \exp(X_i \beta)}.\end{aligned}$$

- Using this, we can consider the F.O.C. and S.O.C. w.r.t.  $\alpha$  and  $\beta$ .