

5. **Another solution for  $\gamma(0)$ :**

From  $\gamma(0) = \phi_1\gamma(1) + \phi_2\gamma(2) + \sigma_\epsilon^2$ ,

$$\gamma(0) = \frac{\sigma_\epsilon^2}{1 - \phi_1\rho(1) - \phi_2\rho(2)}$$

where

$$\rho(1) = \frac{\phi_1}{1 - \phi_2}, \quad \rho(2) = \phi_1\rho(1) + \phi_2 = \frac{\phi_1^2 + (1 - \phi_2)\phi_2}{1 - \phi_2}.$$

6. **Autocorrelation Function of AR(2) Model:**

Given  $\rho(1)$  and  $\rho(2)$ ,

$$\rho(\tau) = \phi_1\rho(\tau - 1) + \phi_2\rho(\tau - 2), \quad \text{for } \tau = 3, 4, \dots,$$

7.  $\phi_{k,k}$  = Partial Autocorrelation Coefficient of AR(2) Process:

$$\begin{pmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(k-1) \\ \rho(1) & 1 & & \rho(k-3) & \rho(k-2) \\ \vdots & \vdots & & \vdots & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & 1 \end{pmatrix} \begin{pmatrix} \phi_{k,1} \\ \phi_{k,2} \\ \vdots \\ \phi_{k,k-1} \\ \phi_{k,k} \end{pmatrix} = \begin{pmatrix} \rho(1) \\ \rho(2) \\ \vdots \\ \rho(k) \end{pmatrix},$$

for  $k = 1, 2, \dots$ .

$$\phi_{k,k} = \frac{\begin{vmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(1) \\ \rho(1) & 1 & & \rho(k-3) & \rho(2) \\ \vdots & \vdots & & \vdots & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & \rho(k) \end{vmatrix}}{\begin{vmatrix} 1 & \rho(1) & \cdots & \rho(k-2) & \rho(k-1) \\ \rho(1) & 1 & & \rho(k-3) & \rho(k-2) \\ \vdots & \vdots & & \vdots & \vdots \\ \rho(k-1) & \rho(k-2) & \cdots & \rho(1) & 1 \end{vmatrix}}$$

Autocovariance Functions:

$$\gamma(1) = \phi_1\gamma(0) + \phi_2\gamma(1),$$

$$\gamma(2) = \phi_1\gamma(1) + \phi_2\gamma(0),$$

$$\gamma(\tau) = \phi_1\gamma(\tau - 1) + \phi_2\gamma(\tau - 2), \quad \text{for } \tau = 3, 4, \dots.$$

Autocorrelation Functions:

$$\rho(1) = \phi_1 + \phi_2\rho(1) = \frac{\phi_1}{1 - \phi_2},$$

$$\rho(2) = \phi_1\rho(1) + \phi_2 = \frac{\phi_1^2}{1 - \phi_2} + \phi_2,$$

$$\rho(\tau) = \phi_1\rho(\tau - 1) + \phi_2\rho(\tau - 2), \quad \text{for } \tau = 3, 4, \dots.$$

$$\phi_{1,1} = \rho(1) = \frac{\phi_1}{1 - \phi_2}$$

$$\phi_{2,2} = \frac{\begin{vmatrix} 1 & \rho(1) \\ \rho(1) & \rho(2) \end{vmatrix}}{\begin{vmatrix} 1 & \rho(1) \\ \rho(1) & 1 \end{vmatrix}} = \frac{\rho(2) - \rho(1)^2}{1 - \rho(1)^2} = \phi_2$$

$$\phi_{3,3} = \frac{\begin{vmatrix} 1 & \rho(1) & \rho(1) \\ \rho(1) & 1 & \rho(2) \\ \rho(2) & \rho(1) & \rho(3) \end{vmatrix}}{\begin{vmatrix} 1 & \rho(1) & \rho(2) \\ \rho(1) & 1 & \rho(1) \\ \rho(2) & \rho(1) & 1 \end{vmatrix}}$$

$$= \frac{(\rho(3) - \rho(1)\rho(2)) - \rho(1)^2(\rho(3) - \rho(1)) + \rho(2)\rho(1)(\rho(2) - 1)}{(1 - \rho(1)^2) - \rho(1)^2(1 - \rho(2)) + \rho(2)(\rho(1)^2 - \rho(2))} = 0.$$

## 8. Log-Likelihood Function — Innovation Form:

$$\log f(y_T, \dots, y_1) = \log f(y_2, y_1) + \sum_{t=3}^T \log f(y_t | y_{t-1}, \dots, y_1)$$

where

$$f(y_2, y_1) = \frac{1}{2\pi} \left| \begin{array}{cc} \gamma(0) & \gamma(1) \\ \gamma(1) & \gamma(0) \end{array} \right|^{-1/2} \exp \left( -\frac{1}{2} (y_1 \ y_2) \begin{pmatrix} \gamma(0) & \gamma(1) \\ \gamma(1) & \gamma(0) \end{pmatrix}^{-1} \begin{pmatrix} y_1 \\ y_2 \end{pmatrix} \right),$$

$$f(y_t | y_{t-1}, \dots, y_1) = \frac{1}{\sqrt{2\pi\sigma_\epsilon^2}} \exp \left( -\frac{1}{2\sigma_\epsilon^2} (y_t - \phi_1 y_{t-1} - \phi_2 y_{t-2})^2 \right).$$

Note as follows:

$$\begin{pmatrix} \gamma(0) & \gamma(1) \\ \gamma(1) & \gamma(0) \end{pmatrix} = \gamma(0) \begin{pmatrix} 1 & \rho(1) \\ \rho(1) & 1 \end{pmatrix} = \gamma(0) \begin{pmatrix} 1 & \phi_1/(1 - \phi_2) \\ \phi_1/(1 - \phi_2) & 1 \end{pmatrix}.$$

9. **AR(2) +drift:**  $y_t = \mu + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \epsilon_t$

Mean:

Rewriting the AR(2)+drift model,

$$\phi(L)y_t = \mu + \epsilon_t$$

where  $\phi(L) = 1 - \phi_1 L - \phi_2 L^2$ .

Under the stationarity assumption, we can rewrite the AR(2)+drift model as follows:

$$y_t = \phi(L)^{-1}\mu + \phi(L)^{-1}\epsilon_t.$$

Therefore,

$$E(y_t) = \phi(L)^{-1}\mu + \phi(L)^{-1}E(\epsilon_t) = \phi(1)^{-1}\mu = \frac{\mu}{1 - \phi_1 - \phi_2}$$

**Example: AR( $p$ ) model:** Consider  $y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \epsilon_t$ .

### 1. Variance of AR( $p$ ) Process:

Under the stationarity condition (i.e., the  $p$  solutions of  $x$  from  $\phi(x) = 0$  are outside the unit circle),

$$\gamma(0) = \frac{\sigma_\epsilon^2}{1 - \phi_1 \rho(1) - \cdots - \phi_p \rho(p)}.$$

Note that  $\gamma(\tau) = \rho(\tau)\gamma(0)$ .

Solve the following simultaneous equations for  $\tau = 0, 1, \cdots, p$ :

$$\begin{aligned} \gamma(\tau) &= E((y_t - \mu)(y_{t-\tau} - \mu)) = E(y_t y_{t-\tau}) \\ &= \begin{cases} \phi_1 \gamma(\tau - 1) + \phi_2 \gamma(\tau - 2) + \cdots + \phi_p \gamma(\tau - p), & \text{for } \tau \neq 0, \\ \phi_1 \gamma(\tau - 1) + \phi_2 \gamma(\tau - 2) + \cdots + \phi_p \gamma(\tau - p) + \sigma_\epsilon^2, & \text{for } \tau = 0. \end{cases} \end{aligned}$$



## 2. Estimation of AR( $p$ ) Model:

### 1. OLS:

$$\min_{\phi_1, \dots, \phi_p} \sum_{t=p+1}^T (y_t - \phi_1 y_{t-1} - \phi_2 y_{t-2} - \dots - \phi_p y_{t-p})^2$$

### 2. MLE:

$$\max_{\phi_1, \dots, \phi_p} \log f(y_T, \dots, y_1)$$

where

$$\log f(y_T, \dots, y_1) = \log f(y_p, \dots, y_2, y_1) + \sum_{t=p+1}^T \log f(y_t | y_{t-1}, \dots, y_1),$$

$$f(y_p, \dots, y_2, y_1) = (2\pi)^{-p/2} |V|^{-1/2} \exp \left( -\frac{1}{2} (y_1 \ y_2 \ \dots \ y_p) V^{-1} \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_p \end{pmatrix} \right)$$

$$V = \gamma(0) \begin{pmatrix} 1 & \rho(1) & \cdots & \rho(p-2) & \rho(p-1) \\ \rho(1) & 1 & & \rho(p-3) & \rho(p-2) \\ \vdots & \vdots & & \vdots & \vdots \\ \rho(p-1) & \rho(p-2) & \cdots & \rho(1) & 1 \end{pmatrix}$$

$$f(y_t|y_{t-1}, \dots, y_1) = \frac{1}{\sqrt{2\pi\sigma_\epsilon^2}} \exp\left(-\frac{1}{2\sigma_\epsilon^2}(y_t - \phi_1 y_{t-1} - \phi_2 y_{t-2} - \cdots - \phi_p y_{t-p})^2\right)$$

### 3. Yule=Walker (ユール・ウォーカー) Equation:

Multiply  $y_{t-1}, y_{t-2}, \dots, y_{t-p}$  on both sides of  $y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \epsilon_t = y_t$ ,

take expectations for each case, and divide by the sample variance  $\hat{\gamma}(0)$ .

$$\begin{pmatrix} 1 & \hat{\rho}(1) & \cdots & \hat{\rho}(p-2) & \hat{\rho}(p-1) \\ \hat{\rho}(1) & 1 & & \hat{\rho}(p-3) & \hat{\rho}(p-2) \\ \vdots & \vdots & & \vdots & \vdots \\ \hat{\rho}(p-1) & \hat{\rho}(p-2) & \cdots & \hat{\rho}(1) & 1 \end{pmatrix} \begin{pmatrix} \phi_1 \\ \phi_2 \\ \vdots \\ \phi_{p-1} \\ \phi_p \end{pmatrix} = \begin{pmatrix} \hat{\rho}(1) \\ \hat{\rho}(2) \\ \vdots \\ \hat{\rho}(p) \end{pmatrix}$$

where

$$\hat{\gamma}(\tau) = \frac{1}{T} \sum_{t=\tau+1}^T (y_t - \hat{\mu})(y_{t-\tau} - \hat{\mu}), \quad \hat{\rho}(\tau) = \frac{\hat{\gamma}(\tau)}{\hat{\gamma}(0)}.$$

3. **AR(p) + drift:**  $y_t = \mu + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \epsilon_t$

Mean:

$$\phi(L)y_t = \mu + \epsilon_t$$

where  $\phi(L) = 1 - \phi_1 L - \phi_2 L^2 - \dots - \phi_p L^p$ .

$$y_t = \phi(L)^{-1} \mu + \phi(L)^{-1} \epsilon_t$$

Taking the expectation on both sides,

$$\begin{aligned} E(y_t) &= \phi(L)^{-1} \mu + \phi(L)^{-1} E(\epsilon_t) = \phi(1)^{-1} \mu \\ &= \frac{\mu}{1 - \phi_1 - \phi_2 - \dots - \phi_p} \end{aligned}$$

#### 4. **Partial Autocorrelation of AR( $p$ ) Process:**

$\phi_{k,k} = 0$  for  $k = p + 1, p + 2, \dots$ .

## 6.4 MA Model

MA (Moving Average , 移動平均) Model:

1. MA(  $q$  )

$$y_t = \epsilon_t + \theta_1\epsilon_{t-1} + \theta_2\epsilon_{t-2} + \cdots + \theta_q\epsilon_{t-q},$$

which is rewritten as:

$$y_t = \theta(L)\epsilon_t,$$

where

$$\theta(L) = 1 + \theta_1L + \theta_2L^2 + \cdots + \theta_qL^q.$$

## 2. Invertibility (反転可能性):

The  $q$  solutions of  $x$  from  $\theta(x) = 1 + \theta_1x + \theta_2x^2 + \cdots + \theta_qx^q = 0$  are outside the unit circle.

$\implies$  MA( $q$ ) model is rewritten as AR( $\infty$ ) model.

**Example: MA(1) Model:**  $y_t = \epsilon_t + \theta_1\epsilon_{t-1}$

### 1. Mean of MA(1) Process:

$$E(y_t) = E(\epsilon_t + \theta_1\epsilon_{t-1}) = E(\epsilon_t) + \theta_1E(\epsilon_{t-1}) = 0$$

### 2. Autocovariance Function of MA(1) Process:

$$\begin{aligned}\gamma(0) &= E(y_t^2) = E(\epsilon_t + \theta_1\epsilon_{t-1})^2 = E(\epsilon_t^2 + 2\theta_1\epsilon_t\epsilon_{t-1} + \theta_1^2\epsilon_{t-1}^2) \\ &= E(\epsilon_t^2) + 2\theta_1E(\epsilon_t\epsilon_{t-1}) + \theta_1^2E(\epsilon_{t-1}^2) = (1 + \theta_1^2)\sigma_\epsilon^2\end{aligned}$$

$$\gamma(1) = E(y_t y_{t-1}) = E((\epsilon_t + \theta_1 \epsilon_{t-1})(\epsilon_{t-1} + \theta_1 \epsilon_{t-2})) = \theta_1 \sigma_\epsilon^2$$

$$\gamma(2) = E(y_t y_{t-2}) = E((\epsilon_t + \theta_1 \epsilon_{t-1})(\epsilon_{t-2} + \theta_1 \epsilon_{t-3})) = 0$$

### 3. Autocorrelation Function of MA(1) Process:

$$\rho(\tau) = \frac{\gamma(\tau)}{\gamma(0)} = \begin{cases} \frac{\theta_1}{1 + \theta_1^2}, & \text{for } \tau = 1, \\ 0, & \text{for } \tau = 2, 3, \dots \end{cases}$$

Let  $x$  be  $\rho(1)$ .

$$\frac{\theta_1}{1 + \theta_1^2} = x, \quad \text{i.e.,} \quad x\theta_1^2 - \theta_1 + x = 0.$$

$\theta_1$  should be a real number.

$$1 - 4x^2 > 0, \quad \text{i.e.,} \quad -\frac{1}{2} \leq \rho(1) \leq \frac{1}{2}.$$

#### 4. Invertibility Condition of MA(1) Process:

$$\begin{aligned}\epsilon_t &= -\theta_1\epsilon_{t-1} + y_t \\ &= (-\theta_1)^2\epsilon_{t-2} + y_t + (-\theta_1)y_{t-1} \\ &= (-\theta_1)^3\epsilon_{t-3} + y_t + (-\theta_1)y_{t-1} + (-\theta_1)^2y_{t-2} \\ &\quad \vdots \\ &= (-\theta_1)^s\epsilon_{t-s} + y_t + (-\theta_1)y_{t-1} + (-\theta_1)^2y_{t-2} + \cdots + (-\theta_1)^{t-s+1}y_{t-s+1}\end{aligned}$$

When  $(-\theta_1)^s\epsilon_{t-s} \rightarrow 0$ , the MA(1) model is written as the AR( $\infty$ ) model, i.e.,

$$y_t = -(-\theta_1)y_{t-1} - (-\theta_1)^2y_{t-2} - \cdots - (-\theta_1)^{t-s+1}y_{t-s+1} - \cdots + \epsilon_t$$

#### 5. Likelihood Function of MA(1) Process:

The autocovariance functions are:  $\gamma(0) = (1 + \theta_1^2)\sigma_\epsilon^2$ ,  $\gamma(1) = \theta_1\sigma_\epsilon^2$ , and  $\gamma(\tau) = 0$  for  $\tau = 2, 3, \dots$ .



The joint distribution of  $y_1, y_2, \dots, y_T$  is:

$$f(y_1, y_2, \dots, y_T) = \frac{1}{(2\pi)^{T/2}} |V|^{-1/2} \exp\left(-\frac{1}{2} Y' V^{-1} Y\right)$$

where

$$Y = \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_T \end{pmatrix}, \quad V = \sigma_\epsilon^2 \begin{pmatrix} 1 + \theta_1^2 & \theta_1 & 0 & \cdots & 0 \\ \theta_1 & 1 + \theta_1^2 & \theta_1 & \ddots & \vdots \\ 0 & \theta_1 & \ddots & \ddots & 0 \\ \vdots & \ddots & \ddots & 1 + \theta_1^2 & \theta_1 \\ 0 & \cdots & 0 & \theta_1 & 1 + \theta_1^2 \end{pmatrix}.$$

6. **MA(1) +drift:**  $y_t = \mu + \epsilon_t + \theta_1 \epsilon_{t-1}$

Mean of MA(1) Process:

$$y_t = \mu + \theta(L)\epsilon_t,$$

where  $\theta(L) = 1 + \theta_1 L$ .

Taking the expectation,

$$E(y_t) = \mu + \theta(L)E(\epsilon_t) = \mu.$$

**Example: MA(2) Model:**  $y_t = \epsilon_t + \theta_1\epsilon_{t-1} + \theta_2\epsilon_{t-2}$

1. Autocovariance Function of MA(2) Process:

$$\gamma(\tau) = \begin{cases} (1 + \theta_1^2 + \theta_2^2)\sigma_\epsilon^2, & \text{for } \tau = 0, \\ (\theta_1 + \theta_1\theta_2)\sigma_\epsilon^2, & \text{for } \tau = 1, \\ \theta_2\sigma_\epsilon^2, & \text{for } \tau = 2, \\ 0, & \text{otherwise.} \end{cases}$$

2. let  $-1/\beta_1$  and  $-1/\beta_2$  be two solutions of  $x$  from  $\theta(x) = 0$ .

For invertibility condition, both  $\beta_1$  and  $\beta_2$  should be less than one in absolute value.

Then, the MA(2) model is represented as:

$$\begin{aligned} y_t &= \epsilon_t + \theta_1\epsilon_{t-1} + \theta_2\epsilon_{t-2} \\ &= (1 + \theta_1L + \theta_2L^2)\epsilon_t \\ &= (1 + \beta_1L)(1 + \beta_2L)\epsilon_t \end{aligned}$$

AR(∞) representation of the MA(2) model is given by:

$$\begin{aligned}\epsilon_t &= \frac{1}{(1 + \beta_1 L)(1 + \beta_2 L)} y_t \\ &= \left( \frac{\beta_1 / (\beta_1 - \beta_2)}{1 + \beta_1 L} + \frac{-\beta_2 / (\beta_1 - \beta_2)}{1 + \beta_2 L} \right) y_t\end{aligned}$$

### 3. Likelihood Function:

$$f(y_1, y_2, \dots, y_T) = \frac{1}{(2\pi)^{T/2}} |V|^{-1/2} \exp\left(-\frac{1}{2} Y' V^{-1} Y\right)$$

where

$$Y = \begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_T \end{pmatrix}, \quad V = \sigma_\epsilon^2 \begin{pmatrix} 1 + \theta_1^2 + \theta_2^2 & \theta_1 + \theta_1\theta_2 & \theta_2 & & & 0 \\ \theta_1 + \theta_1\theta_2 & 1 + \theta_1^2 + \theta_2^2 & \theta_1 + \theta_1\theta_2 & \ddots & & \\ \theta_2 & \theta_1 + \theta_1\theta_2 & \ddots & \ddots & & \theta_2 \\ & \ddots & \ddots & 1 + \theta_1^2 + \theta_2^2 & \theta_1 + \theta_1\theta_2 & \\ 0 & & \theta_2 & \theta_1 + \theta_1\theta_2 & 1 + \theta_1^2 + \theta_2^2 & \end{pmatrix}$$

4. **MA(2) +drift:**  $y_t = \mu + \epsilon_t + \theta_1\epsilon_{t-1} + \theta_2\epsilon_{t-2}$

Mean:

$$y_t = \mu + \theta(L)\epsilon_t,$$

where  $\theta(L) = 1 + \theta_1L + \theta_2L^2$ .

Therefore,

$$E(y_t) = \mu + \theta(L)E(\epsilon_t) = \mu$$

**Example: MA( $q$ ) Model:**  $y_t = \epsilon_t + \theta_1\epsilon_{t-1} + \theta_2\epsilon_{t-2} + \cdots + \theta_q\epsilon_{t-q}$

1. **Mean of MA( $q$ ) Process:**

$$E(y_t) = E(\epsilon_t + \theta_1\epsilon_{t-1} + \theta_2\epsilon_{t-2} + \cdots + \theta_q\epsilon_{t-q}) = 0$$

2. **Autocovariance Function of MA( $q$ ) Process:**

$$\gamma(\tau) = \begin{cases} \sigma_\epsilon^2(\theta_0\theta_\tau + \theta_1\theta_{\tau+1} + \cdots + \theta_{q-\tau}\theta_q) = \sigma_\epsilon^2 \sum_{i=0}^{q-\tau} \theta_i\theta_{\tau+i}, & \tau = 1, 2, \dots, q, \\ 0, & \tau = q + 1, q + 2, \dots, \end{cases}$$

where  $\theta_0 = 1$ .

3. MA( $q$ ) process is stationary.

4. **MA( $q$ ) + drift:**  $y_t = \mu + \epsilon_t + \theta_1\epsilon_{t-1} + \theta_2\epsilon_{t-2} + \cdots + \theta_q\epsilon_{t-q}$

Mean:

$$y_t = \mu + \theta(L)\epsilon_t,$$

where  $\theta(L) = 1 + \theta_1L + \theta_2L^2 + \dots + \theta_qL^q$ .

Therefore, we have:

$$E(y_t) = \mu + \theta(L)E(\epsilon_t) = \mu.$$



## 6.5 ARMA Model

ARMA (Autoregressive Moving Average , 自己回帰移動平均) Process

### 1. ARMA( $p, q$ )

$$y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \epsilon_t + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2} + \cdots + \theta_q \epsilon_{t-q},$$

which is rewritten as:

$$\phi(L)y_t = \theta(L)\epsilon_t,$$

where  $\phi(L) = 1 - \phi_1 L - \phi_2 L^2 - \cdots - \phi_p L^p$  and  $\theta(L) = 1 + \theta_1 L + \theta_2 L^2 + \cdots + \theta_q L^q$ .

### 2. Likelihood Function:

The variance-covariance matrix of  $Y$ , denoted by  $V$ , has to be computed.

**Example: ARMA(1,1) Process:**  $y_t = \phi_1 y_{t-1} + \epsilon_t + \theta_1 \epsilon_{t-1}$

Obtain the autocorrelation coefficient.

The mean of  $y_t$  is to take the expectation on both sides.

$$E(y_t) = \phi_1 E(y_{t-1}) + E(\epsilon_t) + \theta_1 E(\epsilon_{t-1}),$$

where the second and third terms are zeros.

Therefore, we obtain:

$$E(y_t) = 0.$$

The autocovariance of  $y_t$  is to take the expectation, multiplying  $y_{t-\tau}$  on both sides.

$$E(y_t y_{t-\tau}) = \phi_1 E(y_{t-1} y_{t-\tau}) + E(\epsilon_t y_{t-\tau}) + \theta_1 E(\epsilon_{t-1} y_{t-\tau}).$$

Each term is given by:

$$E(y_t y_{t-\tau}) = \gamma(\tau), \quad E(y_{t-1} y_{t-\tau}) = \gamma(\tau - 1),$$

$$E(\epsilon_t y_{t-\tau}) = \begin{cases} \sigma_\epsilon^2, & \tau = 0, \\ 0, & \tau = 1, 2, \dots, \end{cases} \quad E(\epsilon_{t-1} y_{t-\tau}) = \begin{cases} (\phi_1 + \theta_1)\sigma_\epsilon^2, & \tau = 0, \\ \sigma_\epsilon^2, & \tau = 1, \\ 0, & \tau = 2, 3, \dots \end{cases}$$

Therefore, we obtain;

$$\gamma(0) = \phi_1 \gamma(1) + (1 + \phi_1 \theta_1 + \theta_1^2) \sigma_\epsilon^2,$$

$$\gamma(1) = \phi_1 \gamma(0) + \theta_1 \sigma_\epsilon^2,$$

$$\gamma(\tau) = \phi_1 \gamma(\tau - 1), \quad \tau = 2, 3, \dots$$

From the first two equations,  $\gamma(0)$  and  $\gamma(1)$  are computed by:

$$\begin{pmatrix} 1 & -\phi_1 \\ -\phi_1 & 1 \end{pmatrix} \begin{pmatrix} \gamma(0) \\ \gamma(1) \end{pmatrix} = \sigma_\epsilon^2 \begin{pmatrix} 1 + \phi_1 \theta_1 + \theta_1^2 \\ \theta_1 \end{pmatrix}$$

$$\begin{pmatrix} \gamma(0) \\ \gamma(1) \end{pmatrix} = \sigma_\epsilon^2 \begin{pmatrix} 1 & -\phi_1 \\ -\phi_1 & 1 \end{pmatrix}^{-1} \begin{pmatrix} 1 + \phi_1 \theta_1 + \theta_1^2 \\ \theta_1 \end{pmatrix}$$

$$= \frac{\sigma_\epsilon^2}{1 - \phi_1^2} \begin{pmatrix} 1 & \phi_1 \\ \phi_1 & 1 \end{pmatrix} \begin{pmatrix} 1 + \phi_1\theta_1 + \theta_1^2 \\ \theta_1 \end{pmatrix} = \frac{\sigma_\epsilon^2}{1 - \phi_1^2} \begin{pmatrix} 1 + 2\phi_1\theta_1 + \theta_1^2 \\ (1 + \phi_1\theta_1)(\phi_1 + \theta_1) \end{pmatrix}.$$

Thus, the initial value of the autocorrelation coefficient is given by:

$$\rho(1) = \frac{(1 + \phi_1\theta_1)(\phi_1 + \theta_1)}{1 + 2\phi_1\theta_1 + \theta_1^2}.$$

We have:

$$\rho(\tau) = \phi_1\rho(\tau - 1).$$

### ARMA( $p, q$ ) +drift:

$$y_t = \mu + \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \epsilon_t + \theta_1 \epsilon_{t-1} + \theta_2 \epsilon_{t-2} + \cdots + \theta_q \epsilon_{t-q}.$$

Mean of ARMA( $p, q$ ) Process:  $\phi(L)y_t = \mu + \theta(L)\epsilon_t$ ,

where  $\phi(L) = 1 - \phi_1 L - \phi_2 L^2 - \cdots - \phi_p L^p$  and  $\theta(L) = 1 + \theta_1 L + \theta_2 L^2 + \cdots + \theta_q L^q$ .

$$y_t = \phi(L)^{-1} \mu + \phi(L)^{-1} \theta(L) \epsilon_t.$$

Therefore,

$$E(y_t) = \phi(L)^{-1} \mu + \phi(L)^{-1} \theta(L) E(\epsilon_t) = \phi(1)^{-1} \mu = \frac{\mu}{1 - \phi_1 - \phi_2 - \cdots - \phi_p}.$$

## 6.6 ARIMA Model

Autoregressive Integrated Moving Average (ARIMA , 自己回帰和分移動平均) Model

**ARIMA( $p, d, q$ ) Process**

$$\phi(L)\Delta^d y_t = \theta(L)\epsilon_t,$$

where  $\Delta^d y_t = \Delta^{d-1}(1-L)y_t = \Delta^{d-1}y_t - \Delta^{d-1}y_{t-1} = (1-L)^d y_t$  for  $d = 1, 2, \dots$ , and  $\Delta^0 y_t = y_t$ .

**例 : ARIMA(0,1,0) Model**

Consider the model:  $\Delta y_t = y_t - y_{t-1} + \epsilon_t$ ,  $\epsilon_t \sim N(0, \sigma^2)$ ,  $y_0 = 0$ ,

which is rewritten as:  $y_t = \epsilon_t + \epsilon_{t-1} + \dots + \epsilon_1$ .

$$E(y_t) = 0, \quad \gamma(0) = V(y_t) = \sigma^2 t, \quad \gamma(\tau) = \text{Cov}(y_t, y_{t-\tau}) = E(y_t y_{t-\tau}) = \sigma^2(t - \tau),$$

which implies that  $\gamma(\tau)$  is time-dependent.  $\implies y_t$  is not stationary.

$$\rho(\tau) = \frac{\text{Cov}(y_t, y_{t-\tau})}{\sqrt{V(y_t)} \sqrt{V(y_{t-\tau})}} = \frac{t - \tau}{\sqrt{t} \sqrt{t - \tau}} = \sqrt{\frac{t - \tau}{t}}.$$

That is,  $\rho(\tau)$  gradually decreases with slow speed.

## 6.7 SARIMA Model

Seasonal ARIMA (SARIMA) Process:

1. SARIMA( $p, d, q$ )

$$\phi(L)\Delta^d\Delta_s y_t = \theta(L)\epsilon_t,$$

where

$$\Delta_s y_t = (1 - L^s)y_t = y_t - y_{t-s}.$$

$s = 4$  when  $y_t$  denotes quarterly date and  $s = 12$  when  $y_t$  represents monthly data.

## 6.8 Optimal Prediction

1. AR( $p$ ) Process:  $y_t = \phi_1 y_{t-1} + \dots + \phi_p y_{t-p} + \epsilon_t$

(a) Define:

$$E(y_{t+k}|Y_t) = y_{t+k|t},$$

where  $Y_t$  denotes all the information available at time  $t$ .

Taking the conditional expectation of  $y_{t+k} = \phi_1 y_{t+k-1} + \cdots + \phi_p y_{t+k-p} + \epsilon_{t+k}$  on both sides,

$$y_{t+k|t} = \phi_1 y_{t+k-1|t} + \cdots + \phi_p y_{t+k-p|t},$$

where  $y_{s|t} = y_s$  for  $s \leq t$ .

(b) Optimal prediction is given by solving the above differential equation.

2. MA( $q$ ) Process:  $y_t = \epsilon_t + \theta_1 \epsilon_{t-1} + \cdots + \theta_q \epsilon_{t-q}$

(a) Let  $\hat{\epsilon}_T, \hat{\epsilon}_{T-1}, \cdots, \hat{\epsilon}_1$  be the estimated errors.

(b)  $y_{t+k} = \epsilon_{t+k} + \theta_1 \epsilon_{t+k-1} + \cdots + \theta_q \epsilon_{t+k-q}$



(c) Therefore,

$$y_{t+k|t} = \epsilon_{t+k|t} + \theta_1 \epsilon_{t+k-1|t} + \cdots + \theta_q \epsilon_{t+k-q|t},$$

where  $\epsilon_{s|t} = 0$  for  $s > t$  and  $\epsilon_{s|t} = \hat{\epsilon}_s$  for  $s \leq t$ .

3. ARMA( $p, q$ ) Process:  $y_t = \phi_1 y_{t-1} + \cdots + \phi_p y_{t-p} + \epsilon_t + \theta_1 \epsilon_{t-1} + \cdots + \theta_q \epsilon_{t-q}$

(a)  $y_{t+k} = \phi_1 y_{t+k-1} + \cdots + \phi_p y_{t+k-p} + \epsilon_{t+k} + \theta_1 \epsilon_{t+k-1} + \cdots + \theta_q \epsilon_{t+k-q}$

(b) Optimal prediction is:

$$y_{t+k|t} = \phi_1 y_{t+k-1|t} + \cdots + \phi_p y_{t+k-p|t} + \epsilon_{t+k|t} + \theta_1 \epsilon_{t+k-1|t} + \cdots + \theta_q \epsilon_{t+k-q|t},$$

where  $y_{s|t} = y_s$  and  $\epsilon_{s|t} = \hat{\epsilon}_s$  for  $s \leq t$ , and  $\epsilon_{s|t} = 0$  for  $s > t$ .

## 6.9 Identification (識別 , または , 同定)

We have the following two approaches for model specification.

1. Based on AIC or SBIC given  $d, s$ , we obtain  $p, q$ .

(a) AIC (Akaike's Information Criterion , 赤池の情報量基準)

$$\text{AIC} = -2 \log(\text{likelihood}) + 2k,$$

where  $k = p + q$ , which is the number of parameters estimated.

(b) SBIC (Shwarz's Bayesian Information Criterion)

$$\text{SBIC} = -2 \log(\text{likelihood}) + k \log T,$$

where  $T$  denotes the number of observations.

2. From the sample autocorrelation coefficient function  $\hat{\rho}(k)$  and the sample partial autocorrelation coefficient function  $\hat{\phi}_{k,k}$  for  $k = 1, 2, \dots$ , we obtain  $p, d, q, s$ .

	AR( $p$ ) Process	MA( $q$ ) Process
Autocorrelation Function	Gradually decreasing	$\rho(k) = 0,$ $k = q + 1, q + 2, \dots$
Partial Autocorrelation Function	$\phi(k, k) = 0,$ $k = p + 1, p + 2, \dots$	Gradually decreasing

- (a) Compute  $\Delta_s y_t$  to remove seasonality.

Compute the autocovariance functions of  $\Delta_s y_t$ .

If the autocovariance functions have period  $s$ , we take  $(1 - L^s)$ , again.

- (b) Determine the order of difference.

Compute the partial autocovariance functions every time.

If the autocovariance functions decrease as  $\tau$  is large, go to the next step.

- (c) Determine the order of AR terms (i.e.,  $p$ ).

Compute the partial autocovariance functions every time.

The partial autocovariance functions are close to zero after some  $\tau$ , go to the next step.

- (d) Determine the order of MA terms (i.e.,  $q$ ).

Compute the autocovariance functions every time.

If the autocovariance functions are randomly around zero, end of the procedure.

## 6.10 Example of SARIMA using Consumption Data

Construct SARIMA model using monthly and seasonally unadjusted consumption expenditure data and STATA12.

Estimation Period: Jan., 1970 — Dec., 2012 ( $T = 516$ )

```
. gen time=_n                                Generate time.
. tsset time                                  Defined as time series data.
    time variable:  time, 1 to 516
      delta:       1 unit
. corrgram expend                             Variable name: expend
                                           corrgram: Compute autocorrelation
                                           and partial autocorrelation.
```

LAG	AC	PAC	Q	Prob>Q	-1 [Autocorrelation]	0	1 [Partial Autocor]
1	0.8488	0.8499	373.88	0.0000	-----		-----
2	0.8231	0.3858	726.18	0.0000	-----		----
3	0.8716	0.5266	1122	0.0000	-----		-----
4	0.8706	0.4025	1517.6	0.0000	-----		----
5	0.8498	0.3447	1895.3	0.0000	-----		---
6	0.8085	0.0074	2237.9	0.0000	-----		---
7	0.8378	0.1528	2606.5	0.0000	-----		-
8	0.8460	0.1467	2983	0.0000	-----		-
9	0.8342	0.3006	3349.9	0.0000	-----		--

```

10      0.7735  -0.1518      3666  0.0000
11      0.7852  -0.1185     3992.3  0.0000
12      0.9234   0.9442     4444.5  0.0000
13      0.7754  -0.5486     4764.1  0.0000
14      0.7482  -0.3248     5062.1  0.0000
15      0.7963  -0.2392     5400.5  0.0000

```

```

|-----|
|-----|
|-----|
|-----|
|-----|
|-----|

```

Autocorrelation does not approach zero for large lag.  
Time series has unit root.

```
. gen dexp=expend-l.expend
(1 missing value generated)
```

```
Generate dexp=expend-expend(-1),
excluding unit root.
```

```
. corrgram dexp
```

LAG	AC	PAC	Q	Prob>Q	<sup>-1</sup> [Autocorrelation]	<sup>0</sup> [Partial Autocor]	<sup>1</sup>
1	-0.4316	-0.4329	96.485	0.0000	---	---	
2	-0.2546	-0.5441	130.13	0.0000	--	----	
3	0.1721	-0.4091	145.53	0.0000	-	----	
4	0.0667	-0.3459	147.85	0.0000		--	
5	0.0715	-0.0036	150.52	0.0000			
6	-0.2428	-0.1489	181.36	0.0000	-	-	
7	0.0711	-0.1400	184.01	0.0000		-	
8	0.0668	-0.2900	186.36	0.0000		--	
9	0.1704	0.1681	201.64	0.0000	-		-
10	-0.2485	0.1306	234.21	0.0000	-		-
11	-0.4293	-0.9305	331.56	0.0000	---	-----	
12	0.9773	0.6768	837.12	0.0000	-----	-----	-----
13	-0.4152	0.3778	928.56	0.0000	---		----
14	-0.2583	0.2688	964.03	0.0000	--		--

15            0.1712    0.0406    979.63   0.0000

Big autocorrelation at lag 12.

. gen sdex=dexp-l12.dexp  
(13 missing values generated)

Generate sdex=dexp-dexp(-12),  
excluding seasonality.

. corrgram sdex

LAG	AC	PAC	Q	Prob>Q	-1 [Autocorrelation]	0	1	-1 [Partial Autocor]	0	1
1	-0.4752	-0.4753	114.28	0.0000	---			---		
2	-0.0244	-0.3235	114.58	0.0000				--		
3	0.1163	-0.0759	121.46	0.0000						
4	-0.1246	-0.1365	129.37	0.0000						
5	0.0341	-0.1016	129.96	0.0000						
6	-0.0151	-0.1136	130.08	0.0000						
7	-0.0395	-0.1413	130.88	0.0000						
8	0.1123	0.0092	137.35	0.0000						
9	-0.0664	-0.0100	139.62	0.0000						
10	0.0168	0.0069	139.76	0.0000						
11	0.1642	0.2422	153.68	0.0000		-				-
12	-0.3888	-0.2469	231.9	0.0000	---					
13	0.2242	-0.1205	257.96	0.0000		-				
14	-0.0147	-0.0941	258.07	0.0000						
15	-0.0708	-0.0591	260.68	0.0000						

Big autocorrelation at lag 1 (ignore lag 12).

MA(1)

Big partial autocorrelation at lags 1 and 2.

AR(2)

```
. arima sdex, ar(1,2) ma(1)
```

```
Model specification is:
```

```
sdex ~ ARMA(2,1), i.e., expend ~ SARIMA(2,1,1).
```

```
(setting optimization to BHHH)
```

```
Iteration 0: log likelihood = -5107.4608
```

```
Iteration 1: log likelihood = -5102.391
```

```
Iteration 2: log likelihood = -5099.9071
```

```
Iteration 3: log likelihood = -5099.4216
```

```
Iteration 4: log likelihood = -5099.2463
```

```
(switching optimization to BFGS)
```

```
Iteration 5: log likelihood = -5099.2361
```

```
Iteration 6: log likelihood = -5099.2346
```

```
Iteration 7: log likelihood = -5099.2346
```

```
Iteration 8: log likelihood = -5099.2346
```

```
ARIMA regression
```

```
Sample: 14 - 516
```

```
Log likelihood = -5099.235
```

```
Number of obs = 503
```

```
Wald chi2(3) = 973.93
```

```
Prob > chi2 = 0.0000
```

-----						
sdex	Coef.	OPG Std. Err.	z	P> z	[95% Conf. Interval]	
-----						
sdex						
_cons	-15.64573	59.17574	-0.26	0.791	-131.628	100.3366
-----						



```
ARMA
```

ar							
L1.	.1271774	.0581883	2.19	0.029	.0131304	.2412244	
L2.	.1009983	.053626	1.88	0.060	-.0041068	.2061034	
ma							
L1.	-.8343264	.0419364	-19.90	0.000	-.9165202	-.7521326	
/sigma		6111.128	139.0105	43.96	0.000	5838.673	6383.584

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

```
. estat ic
```

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
.	503	.	-5099.235	5	10208.47	10229.57

Note: N=Obs used in calculating BIC; see [R] BIC note

```
. predict resid, r
(13 missing values generated)
```

. corrgram resid

Make sure sdex ~ ARMA(2,1), using the residuals.

LAG	AC	PAC	Q	Prob>Q	$^{-1}$ [Autocorrelation]	$^0$	$^1$ [Partial Autocor]
1	-0.0132	-0.0132	.08814	0.7666			
2	-0.0095	-0.0097	.1341	0.9351			
3	0.1248	0.1246	8.0433	0.0451			
4	-0.0644	-0.0624	10.154	0.0379			
5	-0.0001	0.0011	10.154	0.0710			
6	-0.0138	-0.0309	10.252	0.1144			
7	-0.0032	0.0126	10.257	0.1745			
8	0.0958	0.0938	14.97	0.0597			
9	-0.0317	-0.0255	15.487	0.0784			
10	0.0126	0.0112	15.569	0.1127			
11	-0.0053	-0.0305	15.583	0.1573			
12	-0.3773	-0.3837	89.235	0.0000	---		---
13	0.0408	0.0258	90.098	0.0000			
14	-0.0233	-0.0307	90.381	0.0000			
15	-0.0911	-0.0059	94.703	0.0000			

Big autocorrelation and  
big partial correlation at lag 12.  
We need to re-specify the model.