

- **True Model** $y_t = y_{t-1} + \epsilon_t$ vs **Estimated Model** $y_t = \alpha + \phi y_{t-1} + \epsilon_t$: Under $\alpha = 0$ and $\phi = 1$, we estimate α and ϕ in the regression model:

$$y_t = \alpha + \phi y_{t-1} + \epsilon_t$$

OLSEs of α and ϕ are:

$$\begin{aligned} \begin{pmatrix} \hat{\alpha} \\ \hat{\phi} \end{pmatrix} &= \begin{pmatrix} T & \sum y_{t-1} \\ \sum y_{t-1} & \sum y_{t-1}^2 \end{pmatrix}^{-1} \begin{pmatrix} \sum y_t \\ \sum y_{t-1} y_t \end{pmatrix} = \begin{pmatrix} \alpha \\ \phi \end{pmatrix} + \begin{pmatrix} T & \sum y_{t-1} \\ \sum y_{t-1} & \sum y_{t-1}^2 \end{pmatrix}^{-1} \begin{pmatrix} \sum \epsilon_t \\ \sum y_{t-1} \epsilon_t \end{pmatrix} \\ &= \begin{pmatrix} \alpha \\ \phi \end{pmatrix} + \frac{1}{T \sum y_{t-1}^2 - (\sum y_{t-1})^2} \begin{pmatrix} \sum y_{t-1}^2 & -\sum y_{t-1} \\ -\sum y_{t-1} & T \end{pmatrix} \begin{pmatrix} \sum \epsilon_t \\ \sum y_{t-1} \epsilon_t \end{pmatrix} \\ &= \begin{pmatrix} \alpha \\ \phi \end{pmatrix} + \frac{1}{T \sum y_{t-1}^2 - (\sum y_{t-1})^2} \begin{pmatrix} (\sum y_{t-1}^2)(\sum \epsilon_t) - (\sum y_{t-1})(\sum y_{t-1} \epsilon_t) \\ -(\sum y_{t-1})(\sum \epsilon_t) + T(\sum y_{t-1} \epsilon_t) \end{pmatrix} \end{aligned}$$

$$= \begin{pmatrix} \alpha \\ \phi \end{pmatrix} + \begin{pmatrix} \frac{(\sum y_t^2)(\frac{1}{T} \sum \epsilon_t) - \bar{y} \sum y_{t-1} \epsilon_t}{\sum y_t^2 - T\bar{y}^2} \\ \frac{-\bar{y} \sum \epsilon_t + \sum y_{t-1} \epsilon_t}{\sum y_t^2 - T\bar{y}^2} \end{pmatrix}$$

Note that $\frac{1}{T} \sum y_{t-1} \approx \frac{1}{T} \sum y_t = \bar{y}$ and $\sum y_{t-1}^2 \approx \sum y_t^2$ for large T .

In the true model, $\alpha = 0$ and $\phi = 1$.

$$\begin{pmatrix} \hat{\alpha} \\ \hat{\phi} - 1 \end{pmatrix} = \begin{pmatrix} \frac{(\sum y_t^2)(\frac{1}{T} \sum \epsilon_t) - \bar{y} \sum y_{t-1} \epsilon_t}{\sum y_t^2 - T\bar{y}^2} \\ \frac{-\bar{y} \sum \epsilon_t + \sum y_{t-1} \epsilon_t}{\sum y_t^2 - T\bar{y}^2} \end{pmatrix}$$

For each element of the vector, we consider each term in the numerator and denominator.

- $\sum_{t=1}^T y_{t-1} \epsilon_t$:

Taking the square of $y_t = y_{t-1} + \epsilon_t$ on both sides, we obtain: $y_t^2 = y_{t-1}^2 + y_{t-1} \epsilon_t + \epsilon_t^2$.

Then, we can rewrite as: $y_{t-1} \epsilon_t = \frac{1}{2}(y_t^2 - y_{t-1}^2 - \epsilon_t^2)$ for $y_0 = 0$.

Taking a sum from $t = 1$ to T , we have:

$$\sum_{t=1}^T y_{t-1} \epsilon_t = \frac{1}{2} \sum_{t=1}^T (y_t^2 - y_{t-1}^2 - \epsilon_t^2) = \frac{1}{2} y_T^2 - \frac{1}{2} \sum_{t=1}^T \epsilon_t^2,$$

which is divided by $T\sigma^2$ on both sides, then we obtain:

$$\frac{1}{T\sigma^2} \sum_{t=1}^T y_{t-1} \epsilon_t = \frac{1}{2} \left(\frac{y_T}{\sqrt{T}\sigma} \right)^2 - \frac{1}{2\sigma^2} \frac{1}{T} \sum_{t=1}^T \epsilon_t^2 \longrightarrow \frac{1}{2} W(1)^2 - \frac{1}{2}.$$

Note that $\frac{y_T}{\sqrt{T}\sigma} = W(1)$ and $\frac{1}{T} \sum_{t=1}^T \epsilon_t^2 \longrightarrow E(\epsilon_t^2) = \sigma^2$.

• \bar{y} :

Note that $\frac{1}{T} \sum_{t=1}^T y_t = \bar{y}$ and $\frac{1}{T} \sum_{t=1}^T y_{t-1} = \bar{y}$. We can rewrite \bar{y} as follows:

$$\bar{y} = \frac{1}{T} \sum_{t=1}^T y_t = \sqrt{T}\sigma \frac{1}{T} \sum_{t=1}^T \frac{y_t}{\sqrt{T}\sigma}$$

which is rewritten as:

$$\frac{\bar{y}}{\sqrt{T}\sigma} = \frac{1}{T} \sum_{t=1}^T \frac{y_t}{\sqrt{T}\sigma} \longrightarrow \int_0^1 W(r) dr$$

Note that $\frac{y_t}{\sqrt{T}\sigma} \longrightarrow W(r)$ as $\frac{t}{T} \longrightarrow r$.

- $\sum_{t=1}^T \epsilon_t$:

From $y_T = \sum_{t=1}^T \epsilon_t$, we have: $\frac{1}{\sqrt{T}\sigma} \sum_{t=1}^T \epsilon_t = \frac{y_T}{\sqrt{T}\sigma} = W(1)$.

- $\sum_{t=1}^T y_{t-1}^2$:

From $\sum_{t=1}^T y_{t-1}^2 \approx \sum_{t=1}^T y_t^2$, we obtain: $\frac{1}{T} \sum_{t=1}^T \left(\frac{y_t}{\sqrt{T}\sigma}\right)^2 \rightarrow \int_0^1 W(r)^2 dr$.

Note that $\frac{y_t}{\sqrt{T}\sigma} \rightarrow W(r)$ as $\frac{t}{T} \rightarrow r$.

Thus, $\hat{\phi} - 1 = \frac{\sum y_{t-1}\epsilon_t - \bar{y} \sum \epsilon_t}{\sum y_t^2 - T\bar{y}^2}$ is rewritten as:

$$\begin{aligned} T(\hat{\phi} - 1) &= \frac{\frac{1}{T\sigma^2}(\sum y_{t-1}\epsilon_t - \bar{y} \sum \epsilon_t)}{\frac{1}{T^2\sigma^2}(\sum y_t^2 - T\bar{y}^2)} = \frac{\frac{1}{T\sigma^2} \sum y_{t-1}\epsilon_t - \left(\frac{\bar{y}}{\sqrt{T}\sigma}\right)\left(\frac{1}{\sqrt{T}\sigma} \sum \epsilon_t\right)}{\frac{1}{T} \sum \left(\frac{y_t}{\sqrt{T}\sigma}\right)^2 - \left(\frac{\bar{y}}{\sqrt{T}\sigma}\right)^2} \\ &\rightarrow \frac{\frac{1}{2}(W(1)^2 - 1) - W(1) \int_0^1 W(r)dr}{\int_0^1 W(r)^2 dr - \left(\int_0^1 W(r)dr\right)^2} \end{aligned}$$

Remember that OLSE of α is given by:

$$\hat{\alpha} = \alpha + \frac{(\sum y_t^2)\left(\frac{1}{T} \sum \epsilon_t\right) - \bar{y} \sum y_{t-1}\epsilon_t}{\sum y_t^2 - T\bar{y}^2}$$

Under $\alpha = 0$, $\hat{\alpha}$ is rewritten as follows:

$$\begin{aligned} \sqrt{T}\hat{\alpha} &= \frac{\sigma\left(\frac{1}{T}\sum\left(\frac{y_t}{\sqrt{T}\sigma}\right)^2\right)\left(\frac{1}{\sqrt{T}\sigma}\sum\epsilon_t\right) - \sigma\left(\frac{\bar{y}}{\sqrt{T}\sigma}\right)\left(\frac{1}{T\sigma^2}\sum y_{t-1}\epsilon_t\right)}{\frac{1}{T}\sum\left(\frac{y_t}{\sqrt{T}\sigma}\right)^2 - \left(\frac{\bar{y}}{\sqrt{T}\sigma}\right)^2} \\ &\rightarrow \frac{\sigma W(1)\int_0^1 W(r)^2 dr - \sigma^{\frac{1}{2}}(W(1)^2 - 1)\int_0^1 W(r)dr}{\int_0^1 W(r)^2 dr - \left(\int_0^1 W(r)dr\right)^2} \end{aligned}$$

Thus, convergence speed of $\hat{\phi}$ is different from that of $\hat{\alpha}$.

Neither $\sqrt{T}\hat{\alpha}$ nor $T(\hat{\phi} - 1)$ are normal.

8.3 Serially Correlated Errors

Consider the case where the error term is serially correlated.

8.3.1 Augmented Dickey-Fuller (ADF) Test

Consider the following AR(p) model:

$$y_t = \phi_1 y_{t-1} + \phi_2 y_{t-2} + \cdots + \phi_p y_{t-p} + \epsilon_t, \quad \epsilon_t \sim \text{iid}(0, \sigma^2),$$

which is rewritten as: $\phi(L)y_t = \epsilon_t$.

When the above model has a unit root, we have $\phi(1) = 0$, i.e., $\phi_1 + \phi_2 + \cdots + \phi_p = 1$.

The above AR(p) model is written as:

$$y_t = \rho y_{t-1} + \delta_1 \Delta y_{t-1} + \delta_2 \Delta y_{t-2} + \cdots + \delta_{p-1} \Delta y_{t-p+1} + \epsilon_t,$$

where $\rho = \phi_1 + \phi_2 + \cdots + \phi_p$ and $\delta_j = -(\phi_{j+1} + \phi_{j+2} + \cdots + \phi_p)$.

The null and alternative hypotheses are:

$$H_0 : \rho = 1 \text{ (Unit root),}$$

$$H_1 : \rho < 1 \text{ (Stationary).}$$

Use the t test, where we have the same asymptotic distributions.

We can utilize the same tables as before.

Choose p by AIC or SBIC.

Use $N(0, 1)$ to test $H_0 : \delta_j = 0$ against $H_1 : \delta_j \neq 0$ for $j = 1, 2, \dots, p - 1$.

Reference

Kurozumi (2008) “Economic Time Series Analysis and Unit Root Tests: Development and Perspective,” *Japan Statistical Society*, Vol.38, Series J, No.1, pp.39 – 57.

Download the above paper from:

http://ci.nii.ac.jp/vol_issue/nels/AA11989749/ISS00000426576_ja.html

Example of ADF Test

```
. gen time=_n
. tsset time
      time variable:  time, 1 to 516
      delta: 1 unit
. gen sexpend=expnd-112.expnd
(12 missing values generated)
. corrgram sexpend
```

LAG	AC	PAC	Q	Prob>Q	$^{-1}$ [Autocorrelation]	0 [Partial Autocor]	1
1	0.7177	0.7184	261.14	0.0000	-----	-----	-----
2	0.7036	0.3895	512.6	0.0000	-----	-----	---
3	0.7031	0.2817	764.23	0.0000	-----	-----	--
4	0.6366	0.0456	970.94	0.0000	-----	-----	
5	0.6413	0.1116	1181.1	0.0000	-----	-----	
6	0.6267	0.0815	1382.2	0.0000	-----	-----	
7	0.6208	0.0972	1580	0.0000	-----	-----	
8	0.6384	0.1286	1789.5	0.0000	-----	-----	-
9	0.5926	-0.0205	1970.5	0.0000	-----	-----	
10	0.5847	-0.0014	2146.9	0.0000	-----	-----	
11	0.5658	-0.0185	2312.6	0.0000	-----	-----	
12	0.4529	-0.2570	2418.9	0.0000	-----	-----	--
13	0.5601	0.2318	2581.8	0.0000	-----	-----	-
14	0.5393	0.1095	2733.2	0.0000	-----	-----	
15	0.5277	0.0850	2878.4	0.0000	-----	-----	

. varsoc d.sexpend, exo(l.sexpend) maxlag(25)

Selection-order criteria

Sample: 39 - 516

Number of obs

=

478

lag	LL	LR	df	p	FPE	AIC	HQIC	SBIC
0	-4917.7				5.1e+07	20.5845	20.5914	20.6019
1	-4878.69	78.013	1	0.000	4.3e+07	20.4255	20.4358	20.4516
2	-4858.95	39.481	1	0.000	4.0e+07	20.3471	20.3608	20.382
3	-4858.46	.97673	1	0.323	4.0e+07	20.3492	20.3664	20.3928
4	-4855.44	6.0461	1	0.014	4.0e+07	20.3407	20.3613	20.3931
5	-4853.84	3.1904	1	0.074	4.0e+07	20.3383	20.3623	20.3993
6	-4851.58	4.5304	1	0.033	4.0e+07	20.333	20.3604	20.4027
7	-4847.61	7.942	1	0.005	3.9e+07	20.3205	20.3514	20.399
8	-4847.51	.20154	1	0.653	3.9e+07	20.3243	20.3586	20.4115
9	-4847.51	.00096	1	0.975	3.9e+07	20.3285	20.3662	20.4244
10	-4847.43	.16024	1	0.689	4.0e+07	20.3323	20.3735	20.437
11	-4831.38	32.094	1	0.000	3.7e+07	20.2694	20.3139	20.3828
12	-4818.46	25.834	1	0.000	3.5e+07	20.2195	20.2675	20.3416*
13	-4815.64	5.6341	1	0.018	3.5e+07	20.2119	20.2633	20.3427
14	-4813.98	3.321	1	0.068	3.5e+07	20.2091	20.264	20.3487
15	-4813.38	1.2007	1	0.273	3.5e+07	20.2108	20.2691	20.3591
16	-4810.57	5.6184	1	0.018	3.5e+07	20.2032	20.265	20.3603
17	-4808.7	3.7539	1	0.053	3.5e+07	20.1996	20.2647	20.3653
18	-4806.12	5.1557	1	0.023	3.4e+07	20.195	20.2616	20.3674
19	-4804.6	3.0319	1	0.082	3.4e+07	20.1908	20.2628	20.374
20	-4804.6	2.7e-05	1	0.996	3.5e+07	20.195	20.2704	20.3869
21	-4797.33	14.542	1	0.000	3.4e+07	20.1688	20.2476	20.3694
22	-4794.2	6.2571*	1	0.012	3.3e+07*	20.1598*	20.2422*	20.3692
23	-4793.42	1.5626	1	0.211	3.3e+07	20.1608	20.2465	20.3788
24	-4792.85	1.1533	1	0.283	3.3e+07	20.1625	20.2517	20.3893

```
| 25 | -4792.78 .13518 1 0.713 3.4e+07 20.1664 20.259 20.402 |
+-----+
Endogenous: D.sexpend
Exogenous: L.sexpend _cons
```

```
. dfuller sexpend, lags(22)
```

```
Augmented Dickey-Fuller test for unit root          Number of obs   =          481
```

	Test Statistic	Interpolated Dickey-Fuller		
		1% Critical Value	5% Critical Value	10% Critical Value
Z(t)	-1.627	-3.442	-2.871	-2.570

MacKinnon approximate p-value for Z(t) = 0.4689

```
. dfuller sexpend, lags(12)
```

```
Augmented Dickey-Fuller test for unit root          Number of obs   =          491
```

	Test Statistic	Interpolated Dickey-Fuller		
		1% Critical Value	5% Critical Value	10% Critical Value
Z(t)	-2.399	-3.441	-2.870	-2.570

MacKinnon approximate p-value for Z(t) = 0.1420

⇒ Unit root is detected.

8.4 Cointegration (共和分)

1. For a scalar y_t , when $\Delta y_t = y_t - y_{t-1}$ is a white noise (i.e., iid), we write $\Delta y_t \sim I(1)$.

2. Definition of Cointegration:

Suppose that each series in a $g \times 1$ vector y_t is $I(1)$, i.e., each series has unit root, and that a linear combination of each series (i.e, $a'y_t$ for a nonzero vector a) is $I(0)$, i.e., stationary.

Then, we say that y_t has a cointegration.

a is called the cointegrating vector.

3. Example:

Suppose that $y_t = (y_{1,t}, y_{2,t})'$ is the following vector autoregressive process:

$$y_{1,t} = \phi_1 y_{2,t} + \epsilon_{1,t},$$

$$y_{2,t} = y_{2,t-1} + \epsilon_{2,t}.$$

Then,

$$\Delta y_{1,t} = \phi_1 \epsilon_{2,t} + \epsilon_{1,t} - \epsilon_{1,t-1}, \quad (\text{MA}(1) \text{ process}),$$