

$X_i$  is a random variable.

Consider the distribution of

$$\frac{1}{n} \frac{\partial \log L(\theta)}{\partial \theta} = \frac{1}{n} \sum_{i=1}^n \frac{\partial \log f(X_i; \theta)}{\partial \theta}.$$

We have to obtain mean and variance of  $\frac{\partial \log f(X_i; \theta)}{\partial \theta}$ .

Suppose that  $X_i$  is a continuous type of random variable.

$f(x_i; \theta)$  denotes the density function.

Therefore, we have:

$$\int f(x_i; \theta) dx_i = 1$$

Taking the derivative with respect to  $\theta$  on both sides, we obtain:

$$0 = \int \frac{\partial f(x_i; \theta)}{\partial \theta} dx_i = \int \frac{\partial \log f(x_i; \theta)}{\partial \theta} f(x_i; \theta) dx_i = E\left(\frac{\partial \log f(X_i; \theta)}{\partial \theta}\right)$$

Again, take the derivative with respect to  $\theta$  on both sides as follows:

$$\begin{aligned}
 0 &= \int \frac{\partial^2 \log f(x_i; \theta)}{\partial \theta \partial \theta'} f(x_i; \theta) + \frac{\partial \log f(x_i; \theta)}{\partial \theta} \frac{\partial f(x_i; \theta)}{\partial \theta'} dx_i \\
 &= \int \frac{\partial^2 \log f(x_i; \theta)}{\partial \theta \partial \theta'} f(x_i; \theta) dx_i + \int \frac{\partial \log f(x_i; \theta)}{\partial \theta} \frac{\partial \log f(x_i; \theta)}{\partial \theta'} f(x_i; \theta) dx_i \\
 &= E\left(\frac{\partial^2 \log f(X_i; \theta)}{\partial \theta \partial \theta'}\right) + E\left(\frac{\partial \log f(X_i; \theta)}{\partial \theta} \frac{\partial \log f(X_i; \theta)}{\partial \theta'}\right),
 \end{aligned}$$

i.e.,

$$-E\left(\frac{\partial^2 \log f(X_i; \theta)}{\partial \theta \partial \theta'}\right) = E\left(\frac{\partial \log f(X_i; \theta)}{\partial \theta} \frac{\partial \log f(X_i; \theta)}{\partial \theta'}\right) = V\left(\frac{\partial \log f(X_i; \theta)}{\partial \theta}\right) = \Sigma_i$$

Thus,  $\frac{\partial \log f(X_i; \theta)}{\partial \theta}$  is distributed with mean 0 and variance  $\Sigma_i$ .

Note as follows:

$$I(\theta) = -E\left(\frac{\partial^2 \log L(\theta)}{\partial \theta \partial \theta'}\right) = -\sum_{i=1}^n E\left(\frac{\partial^2 \log f(X_i; \theta)}{\partial \theta \partial \theta'}\right) = \sum_{i=1}^n \Sigma_i.$$

Using the central limit theorem (generalization) shown above, asymptotically we obtain the following distribution:

$$\frac{1}{\sqrt{n}} \frac{\partial \log L(\theta)}{\partial \theta} = \frac{1}{\sqrt{n}} \sum_{i=1}^n \frac{\partial \log f(X_i; \theta)}{\partial \theta} \longrightarrow N(0, \Sigma),$$

where  $\Sigma = \lim_{n \rightarrow \infty} \left( \frac{1}{n} I(\theta) \right)$ .

Let  $\tilde{\theta}$  be the maximum likelihood estimator.

Note that the MLE  $\tilde{\theta}$  satisfies:

$$\frac{\partial \log L(\tilde{\theta})}{\partial \theta} = \sum_{i=1}^n \frac{\partial \log f(X_i; \tilde{\theta})}{\partial \theta} = 0.$$

Linearizing  $\frac{\partial \log L(\tilde{\theta})}{\partial \theta}$  around  $\tilde{\theta} = \theta$ , we obtain:

$$0 = \frac{1}{\sqrt{n}} \frac{\partial \log L(\tilde{\theta})}{\partial \theta} \approx \frac{1}{\sqrt{n}} \frac{\partial \log L(\theta)}{\partial \theta} + \frac{1}{\sqrt{n}} \frac{\partial^2 \log L(\theta)}{\partial \theta \partial \theta'} (\tilde{\theta} - \theta),$$

where the rest of terms (i.e., the second-order term, the third-order term, ...) are ignored, which implies that the distribution of  $\frac{1}{\sqrt{n}} \frac{\partial \log L(\theta)}{\partial \theta}$  is asymptotically equivalent to that of  $\frac{1}{\sqrt{n}} \frac{\partial^2 \log L(\theta)}{\partial \theta \partial \theta'} (\tilde{\theta} - \theta)$ .

We have already known the distribution of  $\frac{1}{\sqrt{n}} \frac{\partial \log L(\theta)}{\partial \theta}$  as follows:

$$\frac{1}{\sqrt{n}} \frac{\partial \log L(\theta)}{\partial \theta} \approx -\frac{1}{\sqrt{n}} \frac{\partial^2 \log L(\theta)}{\partial \theta \partial \theta'} (\tilde{\theta} - \theta) = \left( -\frac{1}{n} \frac{\partial^2 \log L(\theta)}{\partial \theta \partial \theta'} \right) \sqrt{n}(\tilde{\theta} - \theta) \longrightarrow N(0, \Sigma).$$

Note as follows:

$$-\frac{1}{n} \frac{\partial^2 \log L(\theta)}{\partial \theta \partial \theta'} \longrightarrow \lim_{n \rightarrow \infty} \left( \frac{1}{n} \mathbb{E} \left( -\frac{\partial^2 \log L(\theta)}{\partial \theta \partial \theta'} \right) \right) = \lim_{n \rightarrow \infty} \left( \frac{1}{n} I(\theta) \right) = \Sigma.$$

Thus,  $\left( -\frac{1}{n} \frac{\partial^2 \log L(\theta)}{\partial \theta \partial \theta'} \right) \sqrt{n}(\tilde{\theta} - \theta)$  asymptotically has the same distribution as  $\Sigma \sqrt{n}(\tilde{\theta} - \theta)$ .

Therefore,

$$\mathbb{V}(\Sigma \sqrt{n}(\tilde{\theta} - \theta)) = \Sigma \mathbb{V}(\sqrt{n}(\tilde{\theta} - \theta)) \Sigma' \longrightarrow \Sigma.$$

Note that  $\Sigma = \Sigma'$ . Thus, we have the asymptotic variance of  $\sqrt{n}(\tilde{\theta} - \theta)$  as follows:

$$V(\sqrt{n}(\tilde{\theta} - \theta)) \longrightarrow \Sigma^{-1}\Sigma\Sigma^{-1} = \Sigma^{-1}.$$

Finally, we obtain:

$$\sqrt{n}(\tilde{\theta} - \theta) \longrightarrow N(0, \Sigma^{-1}).$$

# 11 Consistency and Asymptotic Normality of OLSE

Regression model:  $y = X\beta + u$ ,  $u \sim (0, \sigma^2 I_n)$ .

## Consistency:

1. Let  $\hat{\beta}_n = (X'X)^{-1}X'y$  be the OLS with sample size  $n$ .

Consistency: As  $n$  is large,  $\hat{\beta}_n$  converges to  $\beta$ .

2. Assume the stationarity condition for  $X$ , i.e.,

$$\frac{1}{n}X'X \longrightarrow M_{xx}.$$

and no correlation between  $X$  and  $u$ , i.e.,

$$\frac{1}{n}X'u \longrightarrow 0.$$

3. Note that  $\frac{1}{n}X'X \rightarrow M_{xx}$  results in  $(\frac{1}{n}X'X)^{-1} \rightarrow M_{xx}^{-1}$ .

4. OLS is given by:

$$\hat{\beta}_n = \beta + (X'X)^{-1}X'u = \beta + (\frac{1}{n}X'X)^{-1}(\frac{1}{n}X'u).$$

Therefore,

$$\hat{\beta}_n \rightarrow \beta + M_{xx}^{-1} \times 0 = \beta$$

Thus, OLSE is a consistent estimator.

### **Asymptotic Normality:**

1. Asymptotic Normality of OLSE

$$\sqrt{n}(\hat{\beta}_n - \beta) \rightarrow N(0, \sigma^2 M_{xx}^{-1}), \quad \text{when } n \rightarrow \infty.$$

2. **Central Limit Theorem:** Greenberg and Webster (1983)

$Z_1, Z_2, \dots, Z_n$  are mutually independent.  $Z_i$  is distributed with mean  $\mu$  and variance  $\Sigma_i$  for  $i = 1, 2, \dots, n$ .

Then, we have the following result:

$$\frac{1}{\sqrt{n}} \sum_{i=1}^n (Z_i - \mu) \longrightarrow N(0, \Sigma),$$

where

$$\Sigma = \lim_{n \rightarrow \infty} \left( \frac{1}{n} \sum_{i=1}^n \Sigma_i \right).$$

Note that the distribution of  $Z_i$  is not assumed.

3. Define  $Z_i = x'_i u_i$ . Then,  $\Sigma_i = V(Z_i) = \sigma^2 x'_i x_i$ .



4.  $\Sigma$  is defined as:

$$\Sigma = \lim_{n \rightarrow \infty} \left( \frac{1}{n} \sum_{i=1}^n \sigma^2 x'_i x_i \right) = \sigma^2 \lim_{n \rightarrow \infty} \left( \frac{1}{n} X'X \right) = \sigma^2 M_{xx},$$

where

$$X = \begin{pmatrix} x_1 \\ x_2 \\ \vdots \\ x_n \end{pmatrix}$$

5. Applying Central Limit Theorem (Greenberg and Webster (1983)), we obtain the following:

$$\frac{1}{\sqrt{n}} \sum_{i=1}^n x'_i u_i = \frac{1}{\sqrt{n}} X' u \longrightarrow N(0, \sigma^2 M_{xx}).$$

On the other hand, from  $\hat{\beta}_n = \beta + (X'X)^{-1} X' u$ , we can rewrite as:

$$\sqrt{n}(\hat{\beta} - \beta) = \left( \frac{1}{n} X'X \right)^{-1} \frac{1}{\sqrt{n}} X' u.$$

$$\begin{aligned}
V\left(\left(\frac{1}{n}X'X\right)^{-1}\frac{1}{\sqrt{n}}X'u\right) &= E\left(\left(\frac{1}{n}X'X\right)^{-1}\frac{1}{\sqrt{n}}X'u\left(\left(\frac{1}{n}X'X\right)^{-1}\frac{1}{\sqrt{n}}X'u\right)'\right) \\
&= \left(\frac{1}{n}X'X\right)^{-1}\left(\frac{1}{n}X'E(uu')X\right)\left(\frac{1}{n}X'X\right)^{-1} \\
&= \sigma^2\left(\frac{1}{n}X'X\right)^{-1}\left(\frac{1}{n}X'X\right)\left(\frac{1}{n}X'X\right)^{-1} \\
&\longrightarrow \sigma^2M_{xx}^{-1}M_{xx}M_{xx}^{-1} = \sigma^2M_{xx}^{-1}.
\end{aligned}$$

Therefore,

$$\sqrt{n}(\hat{\beta} - \beta) \longrightarrow N(0, \sigma^2M_{xx}^{-1})$$

$\implies$  Asymptotic normality (漸近的正規性) of OLSE

The distribution of  $u_i$  is not assumed.

## 12 Instrumental Variable (操作変数法)

### 12.1 Measurement Error (測定誤差)

Errors in Variables

1. True regression model:

$$y = \tilde{X}\beta + u$$

2. Observed variable:

$$X = \tilde{X} + V$$

$V$ : is called the **measurement error** (測定誤差 or 観測誤差).

3. For the elements which do not include measurement errors in  $X$ , the corresponding elements in  $V$  are zeros.

4. Regression using observed variable:

$$y = X\beta + (u - V\beta)$$

OLS of  $\beta$  is:

$$\hat{\beta} = (X'X)^{-1}X'y = \beta + (X'X)^{-1}X'(u - V\beta)$$

5. Assumptions:

(a) The measurement error in  $X$  is uncorrelated with  $\tilde{X}$  in the limit. i.e.,

$$\text{plim}\left(\frac{1}{n}\tilde{X}'V\right) = 0.$$

Therefore, we obtain the following:

$$\text{plim}\left(\frac{1}{n}X'X\right) = \text{plim}\left(\frac{1}{n}\tilde{X}'\tilde{X}\right) + \text{plim}\left(\frac{1}{n}V'V\right) = \Sigma + \Omega$$