

# Econometrics I's Homework #2

**Deadline: July 22, 2025, 10:20AM**

- The answer should be written in English or Japanese.
- Your name and student ID number should be included in your answer sheet.

1 Consider the following regression model:

$$y = X\beta + u = i\beta_1 + X_2\beta_2 + u$$

where  $y$ ,  $X = (i, X_2)$ ,  $\beta = \begin{pmatrix} \beta_1 \\ \beta_2 \end{pmatrix}$  and  $u$  denote  $T \times 1$ ,  $T \times k$ ,  $k \times 1$  and  $T \times 1$  matrices.  $k$  and  $T$  are the number of explanatory variables and the sample size. Note that  $i = (1, 1, \dots, 1)'$ .  $\beta_1$  is a constant term.  $\beta_2$  is a  $(k-1) \times 1$  vector, and  $X_2$  is a  $T \times (k-1)$  matrix.  $u_1, u_2, \dots, u_T$  are mutually independently and **normally** distributed with mean zero and variance  $\sigma^2$ , i.e.,  $u \sim N(0, \sigma^2 I_T)$ .  $\beta$  are a vector of unknown parameters to be estimated. Let  $\hat{\beta}$  be the ordinary least squares estimator of  $\beta$ . We define a vector of residuals as  $e = y - X\hat{\beta}$ .

- (1) Show that  $Mi = 0$  and  $Me = e$ , where  $M = I_T - \frac{1}{T}ii'$ .
- (2) Show that  $e'e = y'My - \hat{\beta}'X'MX\hat{\beta} = y'My - \hat{\beta}_2'X_2'MX_2\hat{\beta}_2$ , where  $\hat{\beta} = \begin{pmatrix} \hat{\beta}_1 \\ \hat{\beta}_2 \end{pmatrix}$ .
- (3) The coefficient of determination is defined as:  $R^2 = 1 - \frac{e'e}{y'My}$ , where  $M = I_T - \frac{1}{T}ii'$ .

Show that  $R^2 = \frac{\hat{\beta}_2'X_2'MX_2\hat{\beta}_2}{y'My}$

- (4) Let  $R$  be an  $G \times k$  vector with  $G \leq k$ . Derive the distribution of  $R\hat{\beta}$ .
- (5) The null hypothesis is given by  $H_0 : R\beta = r$ .  
Under the null hypothesis, derive the distribution of  $R\hat{\beta}$ .
- (6) We want to test  $\beta_2 = 0$ . What are  $G$ ,  $R$  and  $r$ ?

(7) We know that  $\frac{(R\hat{\beta} - r)'(R(X'X)^{-1}R')^{-1}(R\hat{\beta} - r)/G}{e'e/(T - k)} \sim F(G, T - k).$

Show that  $\frac{\hat{\beta}_2'X_2'MX_2\hat{\beta}_2/G}{e'e/(T - k)} \sim F(G, T - k).$

(8) When we test  $H_0 : \beta_2 = 0$ , the test statistic is given by  $\frac{R^2/(k - 1)}{(1 - R^2)/(T - k)}.$

Show that  $\frac{R^2/(k - 1)}{(1 - R^2)/(T - k)} \sim F(k - 1, T - k).$

(9) Suppose that we have the following estimation result:

$$y_i = \begin{matrix} 0.3 \\ (3.163) \end{matrix} + \begin{matrix} 0.65 \\ (0.240) \end{matrix} X_i, \quad s = 1.07, \quad R^2 = 0.786$$

where the sample size is  $T = 4$ , ( ) indicates the standard error of the corresponding coefficient estimate,  $s$  denote the standard error of regression, and  $R^2$  represents the coefficient of determination.

Let  $\beta$  be the coefficient of  $X_i$ .

We want to test  $\beta = 0$  for the regression model:  $y_i = \alpha + \beta X_i + u_i$ .

First, using the estimate 0.65 and its standard error of coefficient (i.e., 0.240), test  $\beta = 0$ .

Second, using  $R^2$ , test  $\beta = 0$ .

Compare both test statistics and make sure  $t^2(T - k) = F(1, T - k)$

You can find various distribution tables in standard introductory statistics textbooks.

2 Consider the regression model:

$$y = X\beta + u, \quad u \sim N(0, \sigma^2 I_T),$$

where  $y$ ,  $X$ ,  $\beta$  and  $u$  are  $T \times 1$ ,  $T \times k$ ,  $k \times 1$  and  $T \times 1$ .

Let  $\hat{\beta}$  be the ordinary least squares estimator, and  $\tilde{\beta}$  be the ordinary least squares estimator restricted to  $R\beta = r$ , where  $R$  and  $r$  are  $G \times k$ ,  $G \times 1$  and  $G \leq k$ .  $\hat{u}$  and  $\tilde{u}$  are defined as the OLS residual and the restricted OLS residual, respectively.

$$y = X\hat{\beta} + \hat{u} \tag{A}$$

$$y = X\tilde{\beta} + \tilde{u}, \quad R\tilde{\beta} = r \tag{B}$$

(1) Derive the restricted OLS  $\tilde{\beta}$ .

(2) Show the following:

$$\frac{(\tilde{u}'\tilde{u} - \hat{u}'\hat{u})/G}{\hat{u}'\hat{u}/(T-k)} \sim F(G, T-k).$$

(3) Suppose that the coefficients of determination in (A) and (B) are given  $\hat{R}^2$  and  $\tilde{R}^2$ , respectively. Show the following:

$$\frac{(\hat{R}^2 - \tilde{R}^2)/G}{(1 - \hat{R}^2)/(T-k)} \sim F(G, T-k).$$

3 We consider estimating the following three production functions.

$$\log(Y_t) = \alpha_0 + \alpha_1 \log(K_t) + \alpha_2 \log(L_t) + u_t$$

$$\log(y_t) = \beta_0 + \beta_1 \log(k_t) + u_t$$

$$\log(Y_t) = \gamma_0 + \gamma_1 \log(K_t) + \gamma_2 \log(L_t) + \gamma_3 D_t + \gamma_4 D_t \log(K_t) + \gamma_5 D_t \log(L_t) + u_t$$

The estimation period is 1969 – 1997 (it's too old!). Let  $Y_t$  be GDP (10 billion yen, 1992 price),  $K_t$  be the net worth (10 billion yen, deflated by the GDP deflator),  $L_t$  be the number of employees,  $D_t$  be the dummy variable, which is one after 1991 and zero before 1991,  $y_t$  be the per capita GDP (10 billion yen, 1992 price,  $y_t = Y_t/L_t$ ), and  $k_t$  be the per capita net worth (10 billion yen, deflated by the GDP deflator,  $k_t = K_t/L_t$ ). The error terms  $u_1, u_2, \dots, u_T$  are mutually independently, identically and normally distributed.

The following estimation results are obtained.

$$\begin{aligned} \log(Y_t) = & - \frac{30.6242}{(7.283)} + \frac{.230042}{(5.054)} \log(K_t) + \frac{2.23565}{(8.266)} \log(L_t) \\ & R^2 = .986684, \quad \bar{R}^2 = .985659, \quad \hat{\sigma}^2 = .00141869 \end{aligned}$$

$$\begin{aligned} \log(y_t) = & - \frac{3.53058}{(41.08)} + \frac{.504043}{(19.62)} \log(k_t) \\ & R^2 = .934448, \quad \bar{R}^2 = .932020, \quad \hat{\sigma}^2 = .00354801 \end{aligned}$$

$$\begin{aligned} \log(Y_t) = & - \frac{34.6168}{(3.630)} + \frac{.204302}{(2.588)} \log(K_t) + \frac{2.48045}{(4.155)} \log(L_t) \\ & - \frac{54.8287}{(1.090)} D_t + \frac{.243766}{(.4665)} D_t \log(K_t) + \frac{2.84275}{(1.134)} D_t \log(L_t) \\ & R^2 = .987960, \quad \bar{R}^2 = .985342, \quad \hat{\sigma}^2 = .00145010 \end{aligned}$$

Note that the values in the parentheses denote the  $t$  values,  $R^2$  is the coefficient of determination,  $\bar{R}^2$  is the adjusted  $R^2$ , and  $\hat{\sigma}^2$  is the variance estimate of regression.

Answer the following questions.

- (1) Test  $H_0 : \alpha_1 = \alpha_2 = 0$ .
- (2) Test whether the production function is homogeneous.
- (3) Test whether the structural change occurred after 1991.

For each question, show the testing procedure in detail.

4

- (1) Show that there exists  $P$  such that  $\Omega = PP'$  when  $\Omega$  is a symmetric and positive definite matrix.
- (2) Consider the following regression model:

$$y_t = x_t\beta + u_t, \quad u_t \sim N(0, \sigma^2 z_t^2), \quad t = 1, 2, \dots, T,$$

where  $u_1, u_2, \dots, u_T$  are mutually independent.

What is the variance-covariance matrix, denoted by  $\sigma^2\Omega$ , of  $u = (u_1, u_2, \dots, u_T)'$ .

- (3) Consider the following regression model:

$$y_t = x_t\beta + u_t, \quad u_t = \rho u_{t-1} + \epsilon_t, \quad \epsilon_t \sim N(0, \sigma^2), \quad t = 1, 2, \dots, T,$$

where  $\epsilon_1, \epsilon_2, \dots, \epsilon_T$  are mutually independent.

What is the variance-covariance matrix, denoted by  $\sigma^2\Omega$ , of  $u = (u_1, u_2, \dots, u_T)'$ .

- (4) Let  $b$  be the best linear unbiased estimator of  $\beta$  under the following regression model:

$$y = X\beta + u, \quad u \sim N(0, \sigma^2\Omega),$$

where  $y, X, \beta$  and  $u$  are  $T \times 1, T \times k, k \times 1$  and  $T \times 1$ .

Derive  $b$ .

- (5) Consider the regression model in (4). We have two estimators,  $\hat{\beta}$  and  $b$ , to estimate  $\beta$ , where  $\hat{\beta}$  denotes the OLS estimator and  $b$  indicates the GLS estimator.

Obtain  $E(\hat{\beta})$  and  $V(\hat{\beta})$ .

Show that  $V(\hat{\beta}) - V(b)$  is a positive definite matrix.

5 Suppose that  $u_1, u_2, \dots, u_T$  are mutually independently and normally distributed with  $E(u_t) = 0$  and  $V(u_t) = \sigma^2$  for all  $t = 1, 2, \dots, T$ .

Consider the following regression model:

$$y_t = \alpha + \beta x_t + u_t$$

Answer the following questions.

- (1) Construct the likelihood function of  $(\alpha, \beta, \sigma^2)$ .
- (2) Obtain the maximum likelihood estimator of  $\theta = (\alpha, \beta, \sigma^2)'$ , denoted by  $\tilde{\theta} = (\tilde{\alpha}, \tilde{\beta}, \tilde{\sigma}^2)'$ .
- (3) Obtain the information matrix  $I(\theta)$ .
- (4) Compare OLSE (i.e.,  $\hat{\beta}$ ) and MLE (i.e.,  $\tilde{\beta}$ ) of  $\beta$  with respect to mean and variance. How about OLSE (i.e.,  $s^2$ ) and MLE (i.e.,  $\hat{\sigma}^2$ ) of  $\sigma^2$ ?

6 Suppose that  $\epsilon_1, \epsilon_2, \dots, \epsilon_T$  are mutually independently and normally distributed with  $E(\epsilon_t) = 0$  and  $V(\epsilon_t) = \sigma^2$  for all  $t = 1, 2, \dots, T$ .

Consider the following regression model:

$$y_t = \rho y_{t-1} + \epsilon_t, \quad \epsilon_t \sim N(0, \sigma^2)$$

Answer the following questions.

- (1) Obtain mean and variance of  $y$  for  $y = (y_1, y_2, \dots, y_T)'$ .
- (2) Using (1), construct the likelihood function of  $(\rho, \sigma^2)$ .
- (3) Obtain unconditional mean and variance of  $y_t$ . Obtain conditional mean and variance of  $y_t$  given  $y_{t-1}, y_{t-2}, \dots$ .
- (4) Using (3), construct the likelihood function in the innovation form.
- (5) Comparing the two likelihood function, obtain the matrix  $P$  such that  $\Omega = PP'$ , where  $\Omega$  denotes the variance obtained in (2).

7 Suppose that  $u_1, u_2, \dots, u_T$  are mutually independently distributed with  $E(u_t) = 0$  and  $V(u_t) = \sigma^2$  for all  $t = 1, 2, \dots, T$ .

Consider the following regression model:

$$y = X\beta + u,$$

where  $y, X, \beta$  and  $u$  are  $T \times 1, T \times k, k \times 1$  and  $T \times 1$  matrices or vectors. Answer the following questions.

- (1) Let  $\hat{\beta}$  be the ordinary least squares estimator of  $\beta$ . Show that  $\hat{\beta}$  is a consistent estimator of  $\beta$ . You have to make clear the underlying assumptions.
- (2) As  $T$  goes to infinity, what is the asymptotic distribution of  $\frac{1}{\sqrt{T}}X'u$ ?
- (3) Obtain the asymptotic distribution of  $\sqrt{T}(\hat{\beta} - \beta)$ .

8 Suppose that  $X_1, X_2, \dots, X_T$  are mutually independently distributed with the density functions  $f(x_i; \theta)$ ,  $i = 1, 2, \dots, T$ .

- (1) Let  $\hat{\theta}$  be the maximum likelihood estimator of  $\theta$ . Show that  $\hat{\theta}$  is a consistent estimator of  $\theta$ . You have to make clear the underlying assumptions.
- (2) As  $T$  goes to infinity, what is the asymptotic distribution of  $\frac{1}{\sqrt{T}} \sum_{i=1}^T \frac{\partial \log f(X_i; \theta)}{\partial \theta}$ ?
- (3) Obtain the asymptotic distribution of  $\sqrt{T}(\hat{\theta} - \theta)$ .

9 Suppose that  $u_1, u_2, \dots, u_T$  are mutually independently distributed with  $E(u_t) = 0$  and  $V(u_t) = \sigma^2$  for all  $t = 1, 2, \dots, T$ .

Consider the following regression model:

$$y = X\beta + u,$$

where  $y$ ,  $X$ ,  $\beta$  and  $u$  are  $T \times 1$ ,  $T \times k$ ,  $k \times 1$  and  $T \times 1$  matrices or vectors. Answer the following questions.

- (1) When  $X$  is correlated with  $u$ , show that OLSE of  $\beta$ , i.e.,  $\hat{\beta}$ , is inconsistent.
- (2)  $X$  is correlated with  $u$ . Suppose that we have the  $T \times k$  matrix  $Z$  which is uncorrelated with  $u$  and correlated with  $X$ . Obtain a consistent estimator of  $\beta$ , using  $Z$ .
- (3) Obtain an asymptotic distribution of the consistent estimator in (2).